



Massachusetts Fisheries Recovery Commission

C/O Center for Marine Science and Technology

706 South Rodney French Boulevard

New Bedford, MA. 02744-1221

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ADDENDUM: ECONOMIC IMPACT OF REGULATIONS ON THE MARINE FISHERIES INDUSTRY IN THE STATE OF MASSACHUSETTS

Brett M. Baden, Ph.D.*

Marcelo Bianconi, Ph.D.**

**August 27, 2006: Marine Fisheries Industry Economic Impact Study for the State of
Massachusetts**

Ref: Contract on Bid No. 06-41: Marine Fisheries Industry Economic Impact Study

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Any remaining errors and/or shortcomings are our sole responsibility.



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Executive Summary

The revised econometric analysis of sales tax data shows negative impacts of Amendments on coastal towns where fishing is active. For Amendment 5, in 1994-1995 US dollars, in the worst case scenario, coastal towns had average losses of \$2,168 per month and a total loss of \$2,691,000 per month. In the best case scenario, coastal towns had average losses of \$354 per month and a total loss of \$176,468. For Amendment 13, the effect on the Port of Gloucester is given by, in 2005 US dollars, in the worst case scenario, an average loss of \$17,339 per month per business unit and a total loss of \$11,400,000 across the town per month. In the best case scenario, the town had average of no losses, of \$0 per month per business unit, and a total gain of \$1,000,000 across the town per month. Hence, the predictive power of future fisheries regulations may have further impacts on gross sales tax receipts that lie within the worst and best case scenarios obtained here. The employment results continue to show that Massachusetts' fishing communities have experienced higher rates of unemployment, lost employment in fishing after the passage of Amendment 13, and lost employment in fishing during the 1990s. In theory, we show that large restrictions can impose negative expected profits which may take a productive unit out of the business activity.

AD.I. Addendum to Data and Econometric Models of the Massachusetts Gross Sales Tax Data

In Baden and Bianconi (2006), chapter II, the econometric analysis of gross sales tax receipts data was done exploring the cross-sectional and time series variation of the data. However, the analysis did not include the possibility of autocorrelation across time of sales tax receipts within each town in the State. It is important to check the robustness of the results previously obtained when autocorrelations are taken into account. Here, we include evidence of the existence of autocorrelation across time, and show how it affects the results previously obtained.

AD.I.1. Amendment 5: May 1993 to April 1995 Period

First, we provide time series analysis and intervention analysis for Amendment 5, focusing on Boston, Gloucester, New Bedford, Fairhaven, Plymouth, and Marshfield.

Boston

Figure AD.1.1 presents the (logarithm of) average real gross sales tax receipts (upper line), the line below represents the introduction of the amendment on May 1, 1994. For the city of Boston, we observe a slight additional variation after the amendment.

Figure AD1.1

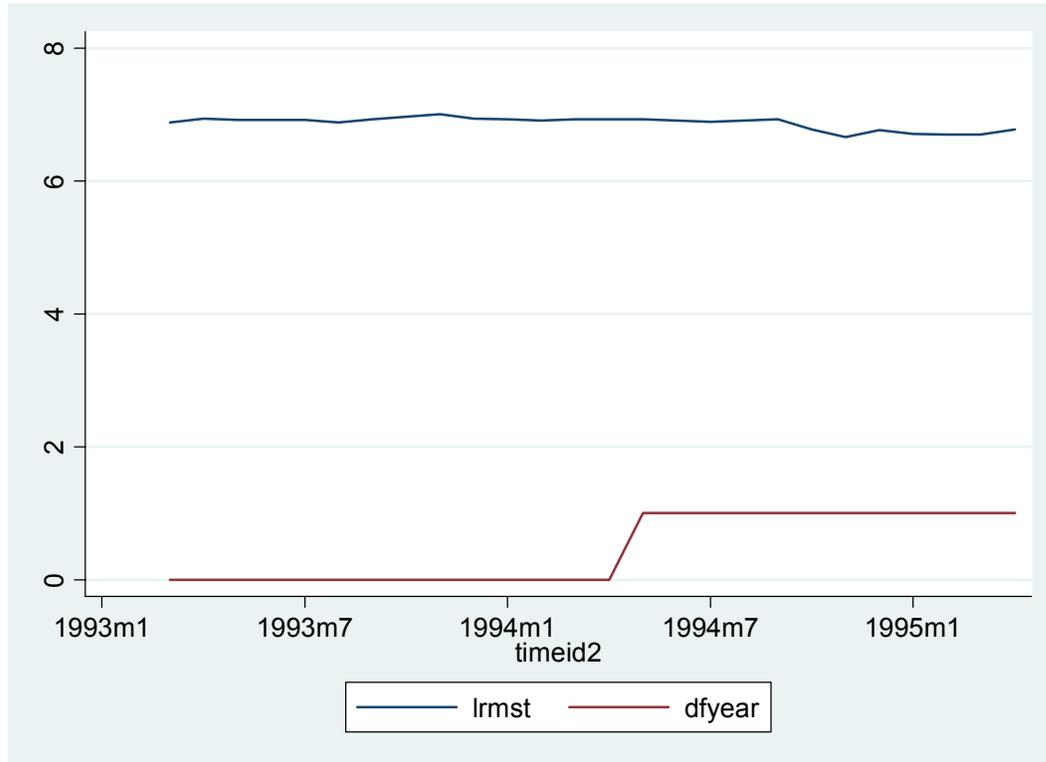


Table AD.1.1 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of

the amendment. We find some evidence of a statistically negative effect, about 4.8% decline, with a 5% significance level.¹

Table: AD.1.1

Linear regression		Number of obs =		25		
		F(2, 22) =		28.55		
		Prob > F =		0.0000		
		R-squared =		0.7467		
		Root MSE =		.05106		

	lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

	lrmst					
	L1.	.7020638	.1637072	4.29	0.000	.3625559 1.041572
	dfyear	-.0488181	.0283635	-1.72	0.099	-.1076404 .0100043
	_cons	2.067471	1.133922	1.82	0.082	-.2841395 4.419081

Figure AD.1.2 presents the (logarithm of) total real gross sales tax receipts (upper line), the line below represents the introduction of the amendment in May 1994. We also observe a very slight additional variation after the amendment.

Figure AD.1.2

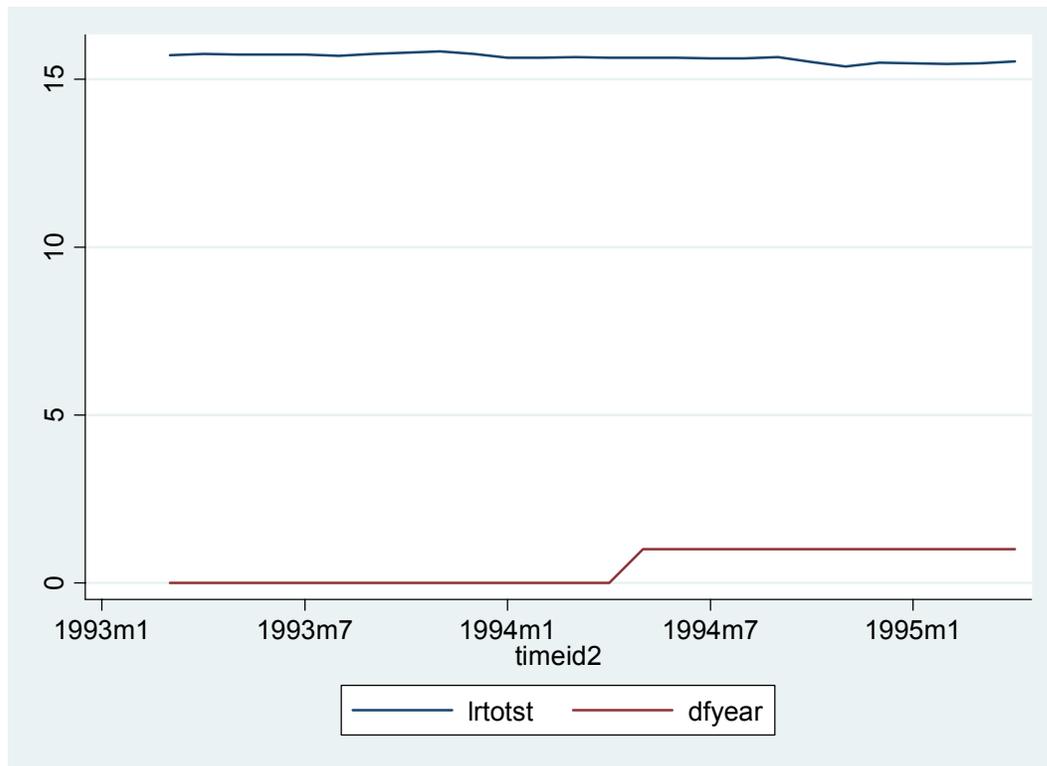


Table AD.1.2 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of

¹ Significance levels are for one-tailed tests. Additional lags were tested in all intervention and other regressions, when statistically significant they were included, otherwise they were excluded.

the amendment. We find some evidence of a negative effect, about 5.8% decline, with a 5.3% significance level.

Table AD.1.2

Linear regression						Number of obs =	25
						F(2, 22) =	41.59
						Prob > F =	0.0000
						R-squared =	0.8052
						Root MSE =	.05538

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	lrtotst						
	L1.	.7020817	.154243	4.55	0.000	.3822012	1.021962
	dfyear	-.0580054	.0345588	-1.68	0.107	-.1296759	.0136652
	_cons	4.684886	2.42598	1.93	0.066	-.346288	9.716059

Gloucester

Figure AD.1.3 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. For Gloucester, we observe additional variation after the amendment.

Figure AD.1.3

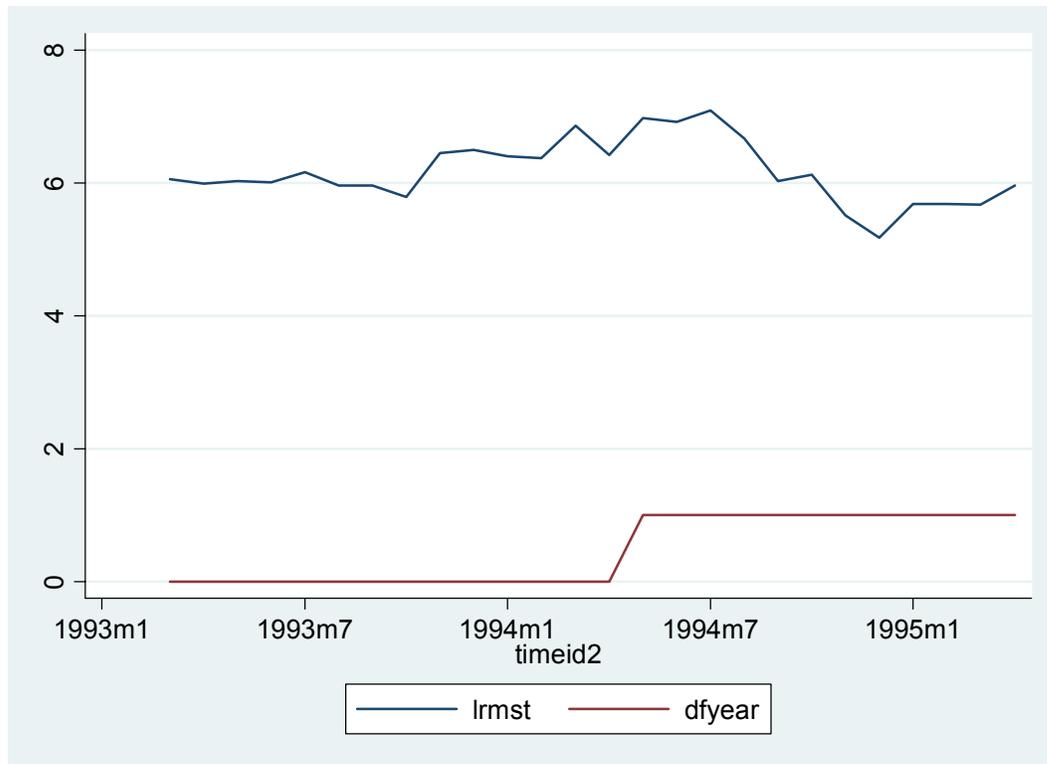


Table AD.1.3 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, but not statistically significant.

Table AD.1.3

Linear regression		Number of obs =		25		
		F(2, 22) =		16.71		
		Prob > F =		0.0000		
		R-squared =		0.5780		
		Root MSE =		.32836		
lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrmst						
L1.	.7557838	.1358592	5.56	0.000	.4740291	1.037539
dfyear	-.0737762	.1340773	-0.55	0.588	-.3518354	.2042831
_cons	1.541461	.848474	1.82	0.083	-.2181662	3.301088

Figure AD.1.4 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. We also observe additional variation after the amendment.

Figure AD.1.4

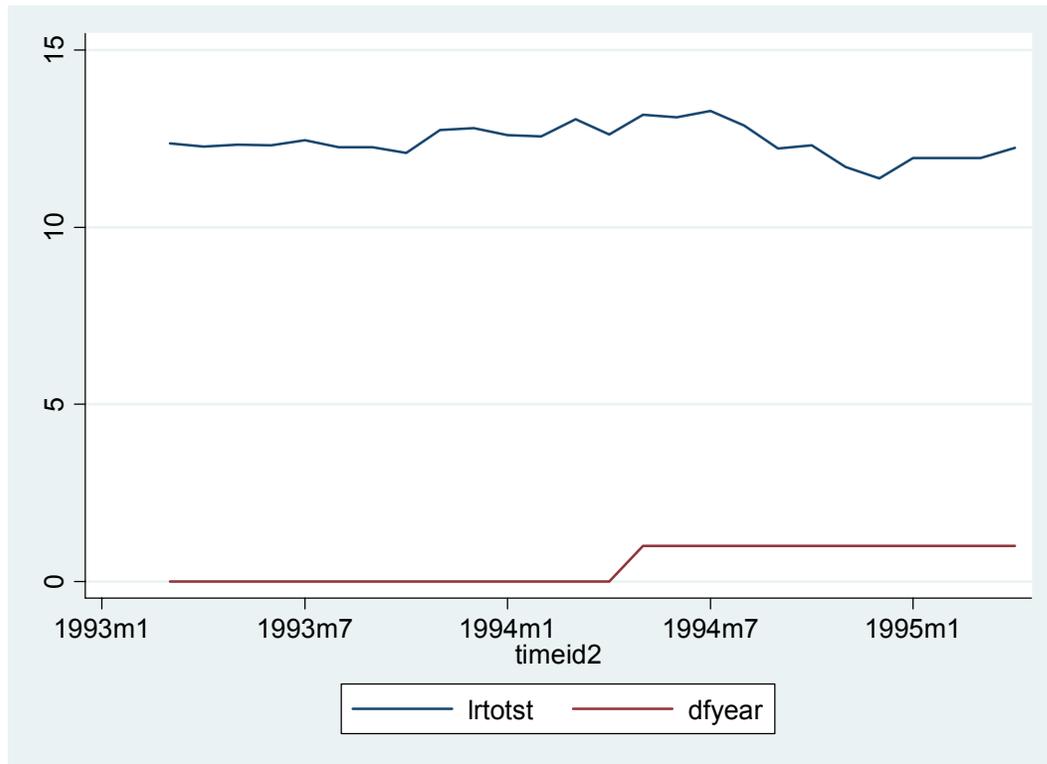


Table AD.1.4 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of

the amendment. We also find a negative effect, but not statistically significant for Gloucester.

Table AD.1.4

Linear regression						Number of obs =	25
						F(2, 22) =	12.95
						Prob > F =	0.0002
						R-squared =	0.5360
						Root MSE =	.33262

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	l1	.7207727	.1484292	4.86	0.000	.4129493	1.028596
	dfyear	-.0769807	.1372089	-0.56	0.580	-.3615346	.2075732
	_cons	3.501684	1.857282	1.89	0.073	-.3500843	7.353452

New Bedford

Figure AD.1.5 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. We observe a slight hump-shaped pattern after the amendment for New Bedford.

Figure AD.1.5

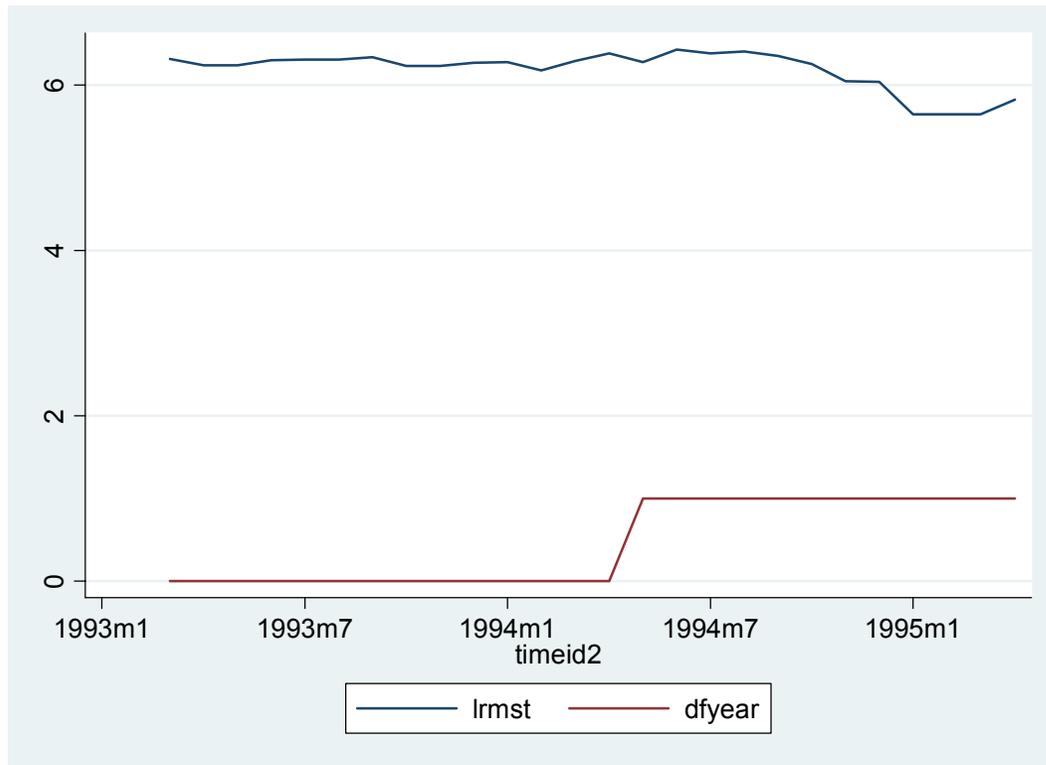


Table AD.1.5 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of

the amendment. We find weak evidence of a negative effect, about 7.2% decline, with a 7% significance level.

Table AD.1.5

Linear regression		Number of obs = 25				
		F(2, 22) = 46.96				
		Prob > F = 0.0000				
		R-squared = 0.7904				
		Root MSE = .11421				
lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrmst						
L1.	.8667046	.0923897	9.38	0.000	.6751	1.058309
dfyear	-.0718274	.0463115	-1.55	0.135	-.1678716	.0242168
_cons	.841227	.5792763	1.45	0.161	-.3601185	2.042573

Figure AD.1.6 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. We again observe a hump-shaped pattern after the amendment.

Figure AD.1.6

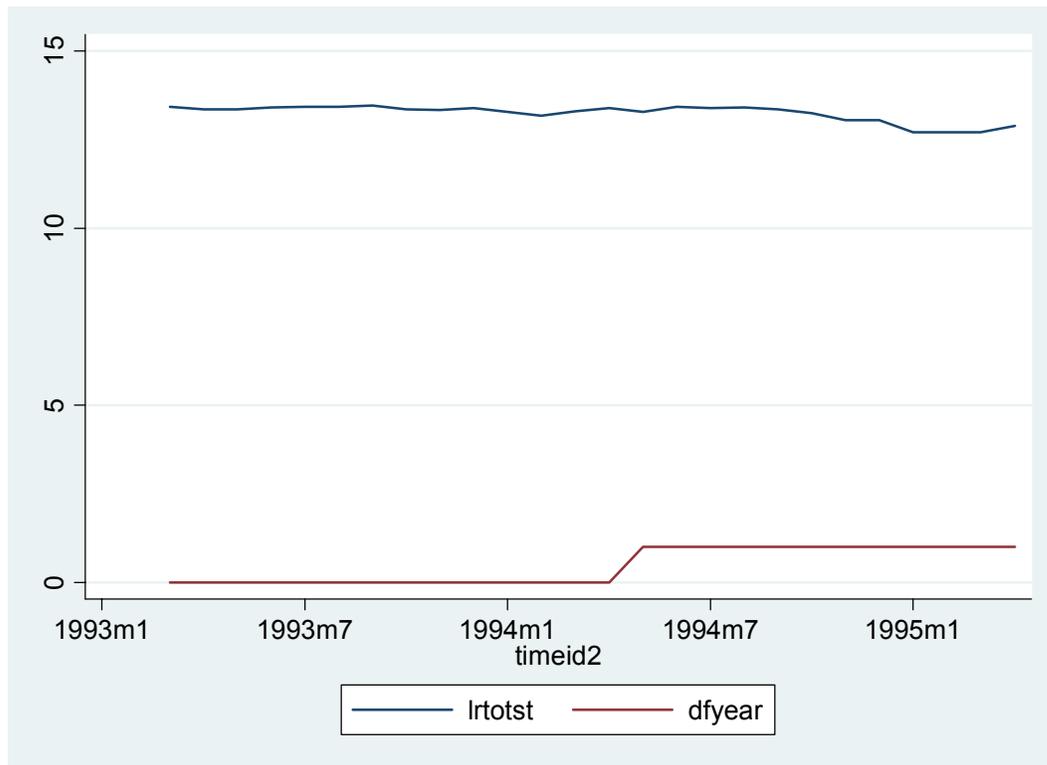


Table AD.1.6 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of

the amendment. We find weak evidence of a negative effect, about 7.1% decline, with a 7% statistical significance level.

Table AD.1.6

Linear regression						Number of obs =	25
						F(2, 22) =	50.64
						Prob > F =	0.0000
						R-squared =	0.8147
						Root MSE =	.10888

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	lrtotst						
	L1.	.8511604	.0947545	8.98	0.000	.6546516	1.047669
	dfyear	-.0717439	.0464427	-1.54	0.137	-.1680602	.0245725
	_cons	1.985678	1.2678	1.57	0.132	-.6435774	4.614933

Fairhaven

Figure AD.1.7 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. We observe some additional variation after the amendment for Fairhaven.

Figure AD.1.7

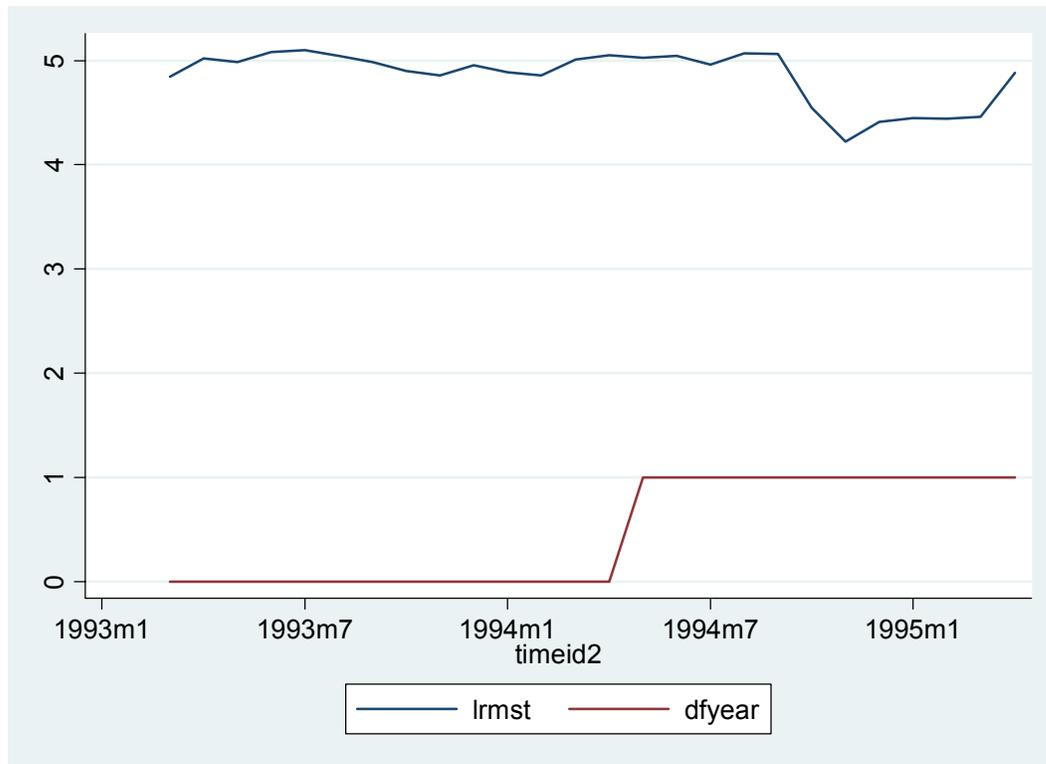


Table AD.1.7 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, but not statistically significant.

Table AD.1.7

Linear regression		Number of obs =		25		
		F(2, 22) =		18.33		
		Prob > F =		0.0000		
		R-squared =		0.6457		
		Root MSE =		.16234		
lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrmst						
L1.	.6926132	.1631972	4.24	0.000	.3541629	1.031064
dfyear	-.1023845	.079878	-1.28	0.213	-.2680414	.0632724
_cons	1.541932	.8115539	1.90	0.071	-.1411278	3.224992

Figure AD.1.8 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. We observe a slight additional variation after the amendment.

Figure AD.1.8

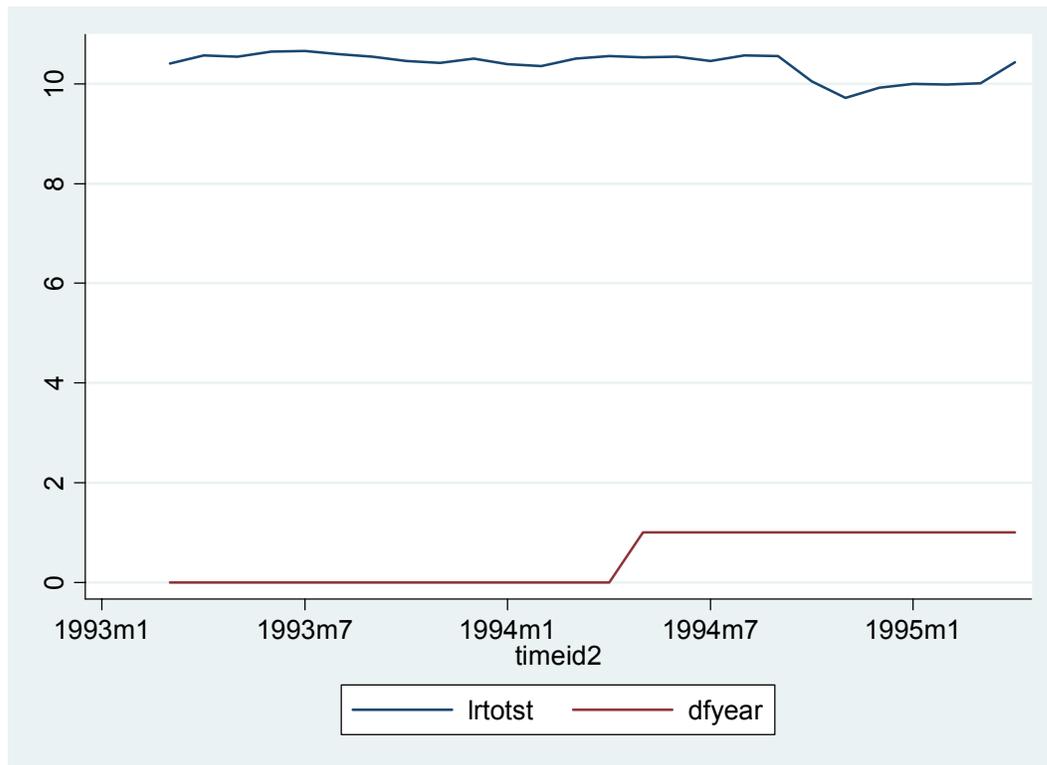


Table AD.1.8 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, but not statistically significant.

Table AD.1.8

Linear regression		Number of obs =		25		
		F(2, 22) =		19.58		
		Prob > F =		0.0000		
		R-squared =		0.6489		
		Root MSE =		.16422		

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

	l1.	.6810002	.1618996	4.21	0.000	.345241 1.016759
	dfyear	-.1072909	.0832944	-1.29	0.211	-.2800329 .0654511
	_cons	3.363831	1.702691	1.98	0.061	-.1673332 6.894995

Plymouth

Figure AD.1.9 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. We observe some additional variation after the amendment.

Figure AD.1.9

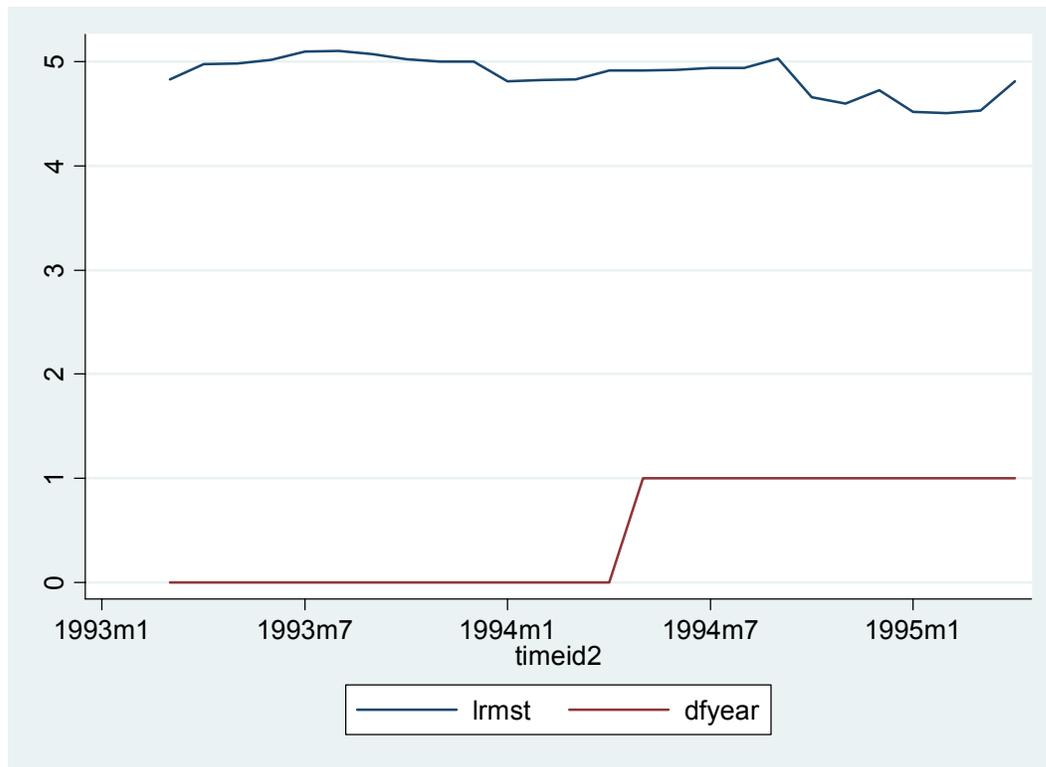


Table AD.1.9 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, but not statistically significant for Plymouth.

Table AD.1.9

Linear regression		Number of obs = 25				
		F(2, 22) = 20.18				
		Prob > F = 0.0000				
		R-squared = 0.6438				
		Root MSE = .11507				

	lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

	lrmst					
	L1.	.6453961	.1912497	3.37	0.003	.2487685 1.042024
	dfyear	-.086538	.0666969	-1.30	0.208	-.2248588 .0517829
	_cons	1.768497	.9507453	1.86	0.076	-.2032278 3.740222

Figure AD.1.10 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. We observe a slight additional variation after the amendment.

Figure AD.1.10

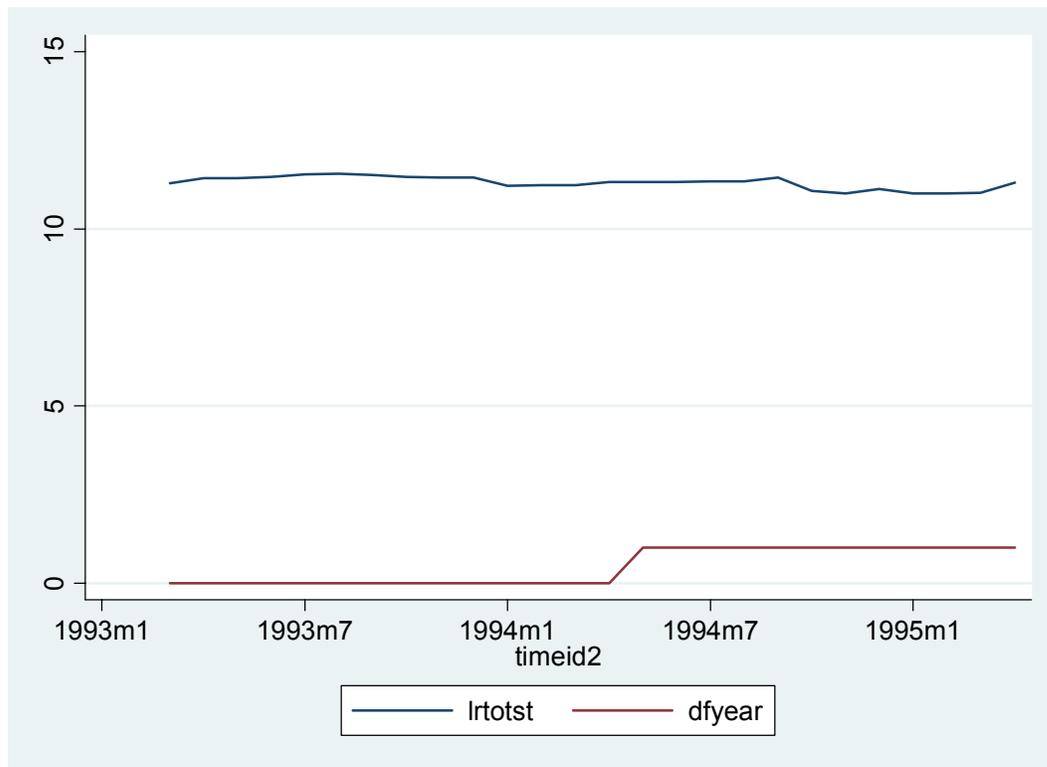


Table AD.1.0 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, but not statistically significant.

Table AD.1.10

Linear regression		Number of obs =		25		
		F(2, 22) =		20.83		
		Prob > F =		0.0000		
		R-squared =		0.6324		
		Root MSE =		.11381		

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

	lrtotst					
	L1.	.6291977	.1948568	3.23	0.004	.2250895 1.033306
	dfyear	-.0839341	.068947	-1.22	0.236	-.2269214 .0590533
	_cons	4.23336	2.222615	1.90	0.070	-.3760609 8.842781

Marshfield

Figure AD.1.11 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. We observe more variation across all period for Marshfield.

Figure AD.1.11

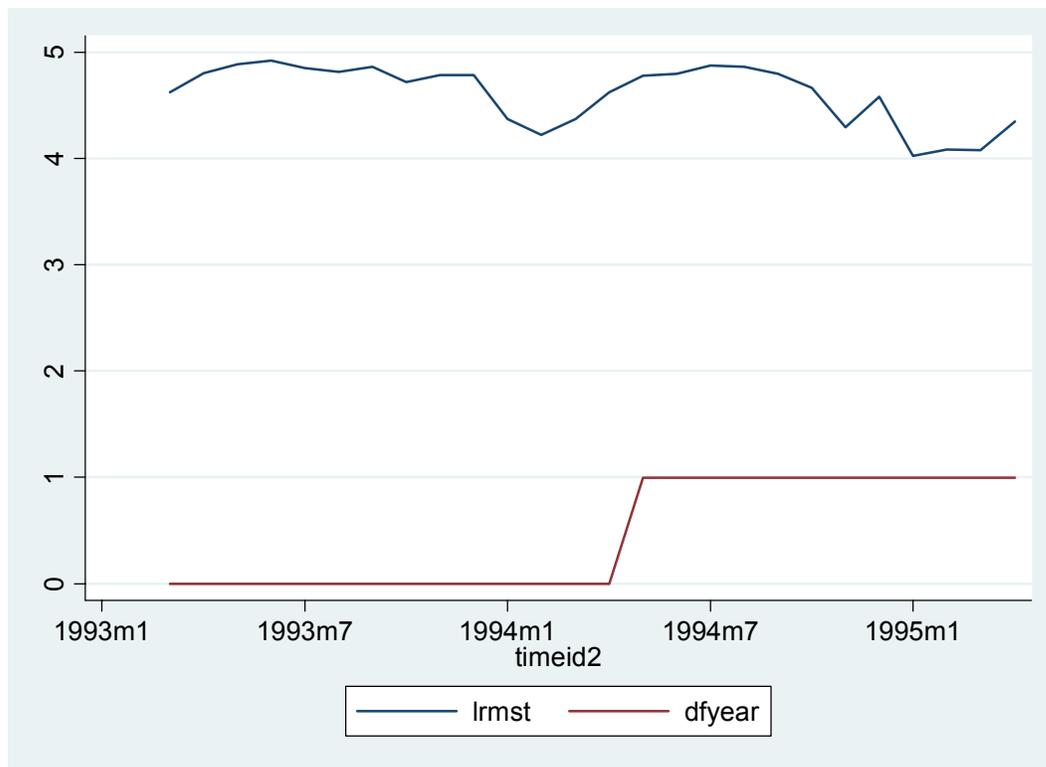


Table AD.1.11 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, but not statistically significant.

Table AD.1.11

Linear regression						Number of obs =	25
						F(2, 22) =	29.99
						Prob > F =	0.0000
						R-squared =	0.5714
						Root MSE =	.19826

	lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	lrmst						
	L1.	.7290684	.1022929	7.13	0.000	.5169259	.9412108
	dfyear	-.0647785	.0837421	-0.77	0.447	-.2384489	.1088919
	_cons	1.272012	.4872598	2.61	0.016	.2614975	2.282527

Figure AD.1.12 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 1994. We observe more variation across all period..

Figure AD.1.12

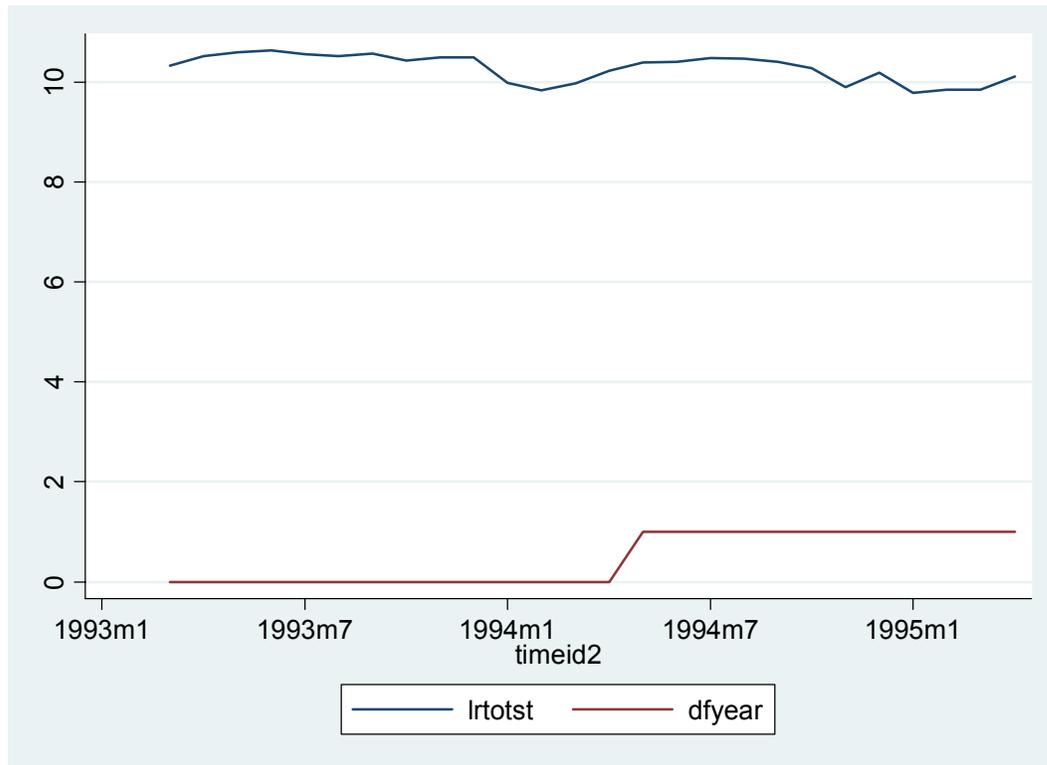


Table AD.1.12 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, but not statistically significant.

Table AD.1.12

Linear regression		Number of obs =		25		
		F(2, 22) =		23.95		
		Prob > F =		0.0000		
		R-squared =		0.5748		
		Root MSE =		.19297		

		Robust				
lrtotst		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]

lrtotst						
L1.		.7213359	.1109822	6.50	0.000	.491173 .9514989
dfyear		-.0565143	.0799462	-0.71	0.487	-.2223125 .1092839
_cons		2.885355	1.154992	2.50	0.020	.4900479 5.280661

In summary, the time series intervention evidence for Amendment 5 is as follows:

Table AD.1.12a

	Policy Intervention Variable	Statistically Significant
Boston	Negative	Yes
Gloucester	Negative	No
New Bedford	Negative	No
Fairhaven	Negative	No
Plymouth	Negative	No
Marshfield	Negative	No

AD.I.1.1. Amendment 5: May 1993 to April 1995 Period - Autocorrelation

We proceed now with estimates of the regressions in Baden and Bianconi (2006), but including lagged dependent variables, and with other assumptions as well.

First, in reference to Table 2.2 in Baden and Bianconi (2006), Table AD.1.13 shows the simple differences-indifferences estimator without further factors, fixed effects or controls, but with a lagged dependent variable. Note that original results are not changed in any dramatic fashion, but the autocorrelation of taxes is statistically significant, with a two month history in this case.² Note also that taking into account the lagged dependent variables increases the explanatory power of the regressions considerably, relative to the original regression in Table 2.2.

² We have performed several other econometric models and tests of fixed and random effects, available upon request. Results are unchanged.

Table AD.1.13

Linear regression							Number of obs = 7978
							F(5, 7972) =13240.51
							Prob > F = 0.0000
							R-squared = 0.9442
							Root MSE = .29254

lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]		

lrmst							
L1.	.7442927	.0300629	24.76	0.000	.6853615	.803224	
L2.	.2406551	.0295319	8.15	0.000	.1827648	.2985454	
dfyear	-.062464	.0067811	-9.21	0.000	-.0757568	-.0491713	
dtreat	-.0033531	.0125926	-0.27	0.790	-.0280379	.0213317	
dyeartreat	.0023568	.0233305	0.10	0.920	-.043377	.0480906	
_cons	.0868473	.0212589	4.09	0.000	.0451743	.1285202	

Linear regression							Number of obs = 7978
							F(5, 7972) =47992.09
							Prob > F = 0.0000
							R-squared = 0.9819
							Root MSE = .29142

lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]		

lrtotst							
L1.	.7525383	.0302483	24.88	0.000	.6932437	.811833	
L2.	.2436991	.0298928	8.15	0.000	.1851014	.3022968	
dfyear	-.0433426	.0066894	-6.48	0.000	-.0564556	-.0302297	
dtreat	.0032653	.0130321	0.25	0.802	-.022281	.0288117	
dyeartreat	.0000211	.023393	0.00	0.999	-.0458353	.0458775	
_cons	.0384676	.0229291	1.68	0.093	-.0064793	.0834146	

In reference to Tables 2.7-2.10 in Baden and Bianconi (2006), we present in Table AD.1.14 models with separate difference-in-difference estimators for each town of the fishing industry group, reflecting changes in each town relative to all other towns in the state. However, here we exclude other fishing towns, and include lagged dependent variables. In Table AD.1.14, we note that the results previously obtained are not robust. The time series autocorrelation of the tax receipts has more explanatory power and the cross sectional evidence is weakened.

Table AD.1.14

Linear regression						Number of obs = 7967	
						F(71, 7895) = 1921.91	
						Prob > F = 0.0000	
						R-squared = 0.9551	
						Root MSE = .26128	

lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]		

lrmst							
L1.	.7139483	.0272806	26.17	0.000	.6604711	.7674255	
L2.	.2756499	.0272069	10.13	0.000	.2223171	.3289826	
dfyear	-.0643485	.0062367	-10.32	0.000	-.0765741	-.0521229	
dtreat	-.0106406	.0190186	-0.56	0.576	-.0479221	.0266408	
dyeartreat	.0522775	.0464371	1.13	0.260	-.0387514	.1433064	
dfeb	.1399109	.0219726	6.37	0.000	.0968388	.182983	
dmar	.2647972	.0199974	13.24	0.000	.225597	.3039974	
dapr	.3841265	.0220181	17.45	0.000	.3409652	.4272878	
dmay	.3062554	.0202533	15.12	0.000	.2665535	.3459572	
djun	.2851283	.0200077	14.25	0.000	.2459079	.3243487	
djul	.2236462	.0209932	10.65	0.000	.1824939	.2647985	
daug	.2542912	.0200383	12.69	0.000	.2150109	.2935716	
dsep	.2202719	.0201644	10.92	0.000	.1807444	.2597993	
doct	.0993815	.0229275	4.33	0.000	.0544376	.1443254	
dnov	.1002066	.0223267	4.49	0.000	.0564403	.1439729	
ddec	.2908418	.0224127	12.98	0.000	.246907	.3347767	
dboston	.0195993	.0386232	0.51	0.612	-.0561124	.0953111	
dchatam	.0028996	.0694195	0.04	0.967	-.133181	.1389801	
dfairhaven	.0056989	.0289683	0.20	0.844	-.0510867	.0624845	
dgloucester	.0665489	.0861426	0.77	0.440	-.1023134	.2354112	
dmarshfield	-.0250851	.0352297	-0.71	0.476	-.0941447	.0439744	
dnewbedford	.02438	.0302798	0.81	0.421	-.0349764	.0837363	
dplymouth	-.0060727	.0248507	-0.24	0.807	-.0547867	.0426412	
dprovincetown	-.0135327	.082284	-0.16	0.869	-.1748311	.1477657	
dscituate	.007596	.0385873	0.20	0.844	-.0680454	.0832373	
dbeverly	-.0102718	.0320365	-0.32	0.748	-.0730718	.0525282	
deastham	.003526	.0655217	0.05	0.957	-.124914	.1319659	
dedgartown	-.0178232	.0943736	-0.19	0.850	-.2028204	.167174	
dharwich	.0107813	.0442502	0.24	0.808	-.0759609	.0975234	
dwestport	.0150642	.0249394	0.60	0.546	-.0338237	.0639521	
dmarblehead	-.0092377	.0404991	-0.23	0.820	-.0886266	.0701511	
dnantucket	.0167553	.0745199	0.22	0.822	-.1293235	.162834	
dnewburyport	.0002102	.056452	0.00	0.997	-.1104506	.110871	
drockport	.0000613	.0443717	0.00	0.999	-.0869189	.0870415	
dsandwich	.027157	.0609075	0.45	0.656	-.0922377	.1465518	
dswamp	-.0099235	.0394964	-0.25	0.802	-.0873469	.0674999	
dwellfleet	-.0013824	.1133443	-0.01	0.990	-.2235672	.2208023	
dyarmouth	.0119995	.0396239	0.30	0.762	-.0656738	.0896728	
dsalem	.0128536	.0319433	0.40	0.687	-.0497636	.0754709	
dmanchester	.0246584	.0511217	0.48	0.630	-.0755537	.1248705	
dorleans	.0111371	.0330535	0.34	0.736	-.0536564	.0759307	
doakb	.016459	.1148089	0.14	0.886	-.2085967	.2415147	
dtisbury	.0036947	.0528354	0.07	0.944	-.0998767	.107266	
ddartmouth	.0186071	.0306497	0.61	0.544	-.0414745	.0786887	
dtboston	-.0072655	.0608854	-0.12	0.905	-.126617	.112086	
dtchatam	-.0954083	.1108112	-0.86	0.389	-.3126276	.1218111	
dtfairhaven	-.0241559	.07653	-0.32	0.752	-.174175	.1258632	
dtgloucester	-.108352	.1484017	-0.73	0.465	-.3992586	.1825546	
dtmarshfield	.0022225	.0661967	0.03	0.973	-.1275406	.1319857	
dtnewbedford	-.0630977	.0558201	-1.13	0.258	-.1725199	.0463244	

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dtplymouth	-.0006434	.0547319	-0.01	0.991	-.1079325	.1066456
dtprovince~n	-.108413	.2010048	-0.54	0.590	-.5024355	.2856095
dtscituate	-.0013097	.1187624	-0.01	0.991	-.2341154	.2314959
dteastham	-.1285599	.1169872	-1.10	0.272	-.3578858	.100766
dtedgardtown	-.0206485	.1593545	-0.13	0.897	-.3330254	.2917284
dtharwich	-.0492016	.0751428	-0.65	0.513	-.1965013	.0980982
dtwestport	-.033191	.0595145	-0.56	0.577	-.1498552	.0834732
dtmarblehead	-.0497002	.0791436	-0.63	0.530	-.2048426	.1054422
dtnantucket	-.0277093	.1106829	-0.25	0.802	-.2446771	.1892585
dtnewburyp~t	.0021287	.0829154	0.03	0.980	-.1604075	.1646649
dtrockport	-.0721718	.0747707	-0.97	0.334	-.2187421	.0743985
dtswamp	-.1230408	.179101	-0.69	0.492	-.4741261	.2280445
dtswamp	-.0448389	.0647544	-0.69	0.489	-.1717747	.082097
dtwellfleet	-.0467202	.1830025	-0.26	0.798	-.4054535	.3120131
dtymouth	-.0228779	.0828293	-0.28	0.782	-.1852452	.1394895
dtsealem	-.0162487	.0577407	-0.28	0.778	-.1294357	.0969383
dtmanchester	-.104601	.0816229	-1.28	0.200	-.2646034	.0554014
dtorleans	-.0260683	.0693566	-0.38	0.707	-.1620255	.1098889
dtoakb	-.0996095	.1940096	-0.51	0.608	-.4799198	.2807007
dttisbury	-.0221011	.1222417	-0.18	0.857	-.2617271	.2175249
dtdartmouth	-.08306	.0803825	-1.03	0.301	-.240631	.0745111
_cons	-.1506941	.0271255	-5.56	0.000	-.2038672	-.0975209

Linear regression

Number of obs = 7967
 F(71, 7895) = 7034.25
 Prob > F = 0.0000
 R-squared = 0.9856
 Root MSE = .25901

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrtotst						
L1.	.7214126	.0277769	25.97	0.000	.6669625	.7758627
L2.	.2765529	.0276941	9.99	0.000	.2222652	.3308407
dfyear	-.0449614	.0060961	-7.38	0.000	-.0569114	-.0330115
dtreat	-.0072685	.015766	-0.46	0.645	-.038174	.023637
dyeartreat	.0510754	.0494919	1.03	0.302	-.0459418	.1480927
dfeb	.1462289	.0218432	6.69	0.000	.1034104	.1890474
dmar	.273487	.0194708	14.05	0.000	.2353192	.3116549
dapr	.3925167	.0215326	18.23	0.000	.3503071	.4347264
dmay	.312041	.0197343	15.81	0.000	.2733567	.3507254
djun	.2905136	.0195232	14.88	0.000	.2522429	.3287843
djul	.2287195	.0205768	11.12	0.000	.1883835	.2690555
daug	.259501	.0195992	13.24	0.000	.2210813	.2979206
dsep	.2253334	.0196906	11.44	0.000	.1867347	.2639322
doct	.1045806	.0225838	4.63	0.000	.0603104	.1488509
dnov	.1064037	.022046	4.83	0.000	.0631876	.1496198
ddec	.2978994	.0220333	13.52	0.000	.2547083	.3410906
dboston	.0080696	.0343592	0.23	0.814	-.0592834	.0754226
dchatam	.0125005	.0699778	0.18	0.858	-.1246746	.1496755
dfairhaven	.0085377	.0249843	0.34	0.733	-.0404381	.0575136
dgloucester	.0551272	.0854265	0.65	0.519	-.1123314	.2225858
dmarshfield	-.0245359	.0385227	-0.64	0.524	-.1000506	.0509788
dnewbedford	.013353	.0233389	0.57	0.567	-.0323974	.0591035
dplymouth	.0003012	.0220417	0.01	0.989	-.0429065	.0435088
dprovincet~n	.0009327	.0817928	0.01	0.991	-.1594029	.1612683
dscituate	.0155711	.0364002	0.43	0.669	-.0557829	.0869252
dbeverly	-.0127243	.0343911	-0.37	0.711	-.08014	.0546913
deastham	.0024975	.068126	0.04	0.971	-.1310474	.1360424
dedgardtown	-.0169229	.0966941	-0.18	0.861	-.206469	.1726231

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dharwich	.0217285	.0431767	0.50	0.615	-.0629092	.1063663
dwestport	.0192533	.0197167	0.98	0.329	-.0193967	.0579032
dmarblehead	-.0059603	.0416923	-0.14	0.886	-.0876883	.0757677
dnantucket	.0228842	.0742901	0.31	0.758	-.1227441	.1685125
dnewburyport	-.000106	.0558494	-0.00	0.998	-.1095856	.1093736
drockport	.0085147	.0458053	0.19	0.853	-.0812758	.0983052
dsandwich	.0154651	.0501976	0.31	0.758	-.0829354	.1138656
dswamp	-.0106642	.0369561	-0.29	0.773	-.0831079	.0617794
dwellfleet	.0093262	.1127602	0.08	0.934	-.2117135	.230366
dyarmouth	.0072285	.0437291	0.17	0.869	-.0784921	.0929491
dsalem	.0065124	.0318398	0.20	0.838	-.0559021	.0689268
dmanchester	.0200396	.0436937	0.46	0.647	-.0656116	.1056908
dorleans	.0149847	.0316955	0.47	0.636	-.0471468	.0771162
doakb	.030361	.1147272	0.26	0.791	-.1945346	.2552567
dtisbury	.0056361	.0512518	0.11	0.912	-.0948309	.1061031
ddartmouth	.0192455	.0309278	0.62	0.534	-.0413811	.0798722
dtboston	-.0101575	.0629153	-0.16	0.872	-.1334882	.1131732
dtchatam	-.1005223	.111434	-0.90	0.367	-.3189624	.1179178
dtfairhaven	-.0298328	.0796338	-0.37	0.708	-.1859361	.1262704
dtgloucester	-.106299	.1530262	-0.69	0.487	-.4062709	.1936729
dtmarshfield	.0122749	.0685584	0.18	0.858	-.1221177	.1466675
dtnewbedford	-.0623755	.0556675	-1.12	0.263	-.1714985	.0467476
dtplymouth	-.0045442	.0584039	-0.08	0.938	-.1190313	.1099429
dtprovince~n	-.1318013	.2049282	-0.64	0.520	-.5335148	.2699121
dtscituate	-.008136	.1223407	-0.07	0.947	-.2479562	.2316841
dteastham	-.1078282	.1141751	-0.94	0.345	-.3316416	.1159851
dtedgardtown	-.0123594	.1597038	-0.08	0.938	-.3254212	.3007023
dtharwich	-.0573084	.0768814	-0.75	0.456	-.2080163	.0933995
dtwestport	-.0288465	.0626459	-0.46	0.645	-.1516489	.093956
dtmarblehead	-.0482197	.0791236	-0.61	0.542	-.2033228	.1068834
dtnantucket	-.0341738	.1119877	-0.31	0.760	-.2536994	.1853518
dtnewburyp~t	-.000206	.087964	-0.00	0.998	-.1726388	.1722268
dtrockport	-.0750243	.0769183	-0.98	0.329	-.2258045	.075756
dtsandwich	-.1035172	.1839523	-0.56	0.574	-.4641124	.257078
dswamp	-.0443804	.0668368	-0.66	0.507	-.1753982	.0866374
dwellfleet	-.0599417	.1841434	-0.33	0.745	-.4209114	.3010279
dtarmouth	-.0177201	.0815462	-0.22	0.828	-.1775722	.1421319
dsalem	-.0192385	.060546	-0.32	0.751	-.1379246	.0994476
dmanchester	-.0889544	.072834	-1.22	0.222	-.2317283	.0538196
dorleans	-.0363421	.072347	-0.50	0.615	-.1781613	.1054771
dtoakb	-.1172873	.1949371	-0.60	0.547	-.4994157	.2648411
dtisbury	-.0319251	.1237086	-0.26	0.796	-.2744266	.2105764
dtdartmouth	-.0891465	.0822059	-1.08	0.278	-.2502918	.0719988
_cons	-.1988406	.0287099	-6.93	0.000	-.2551196	-.1425616

In reference to the next set of regressions, Tables 2.11-2.13, in Baden and Bianconi (2006), we restrict the data to the subsample of 29 towns and cities and localities affected by the fishing industry only, but now include lagged dependent variables.³ Table AD.1.15 shows that original results are qualitative similar, we observe a negative impact of the policy intervention on the tax receipts of the 29 coastal towns where fishing is active. However, the effects are significantly smaller in magnitude, 4.4% and 2.8% lower average and total gross real sales tax receipts for all towns in this subsample as shown in the shaded rows. Taking into account the facts presented that the share of those fishing communities on the total gross state tax receipts for the whole State is about 6% for

³ The second lag is not significant in this case.

averages and about 19% for total, 4.4% and 2.8% loss in tax revenues for the fishing localities could still be a potentially significant amount for the State as a whole.

An approximate dollar value figure for the losses is as follows. For the averages, 4.4% represents an approximate \$1,247 monthly loss in average sales tax revenues for unit in each town in the fishing industry, measured in 1994-1995 US dollars.⁴ Given the confidence intervals of estimation, these measurements could be as large as \$2,168 monthly loss in average sales tax revenues for each unit in each town in the fishing industry, and as low as \$354 per month. For the total, 2.8% represents an approximate \$1,235,000 monthly loss in total gross sales tax revenues for each town in the fishing industry, measured in 1994-1995 dollars.⁵ Given the confidence intervals of estimation, these measurements could be as large as \$2,691,000 monthly loss in average sales tax revenues for each town in the fishing industry, and as low as \$176,468 per month.

Table AD.1.15

Linear regression						Number of obs = 700	
						F(13, 686) = 969.05	
						Prob > F = 0.0000	
						R-squared = 0.9386	
						Root MSE = .2187	

	lrmst1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

lrmst1							
L1.		.9678651	.0115822	83.56	0.000	.9451243	.990606
dfyear		-.0445035	.0162956	-2.73	0.006	-.0764988	-.0125083
dfeb		.271082	.0578522	4.69	0.000	.1574933	.3846707
dmar		.3451696	.0586529	5.88	0.000	.2300089	.4603304
dapr		.5197878	.0602827	8.62	0.000	.4014271	.6381485
dmay		.4091118	.0583584	7.01	0.000	.2945292	.5236944
djun		.4147587	.0577474	7.18	0.000	.3013758	.5281416
djul		.4343466	.0626624	6.93	0.000	.3113135	.5573797
daug		.3315244	.0561772	5.90	0.000	.2212246	.4418243
dsep		.1696356	.0607863	2.79	0.005	.0502861	.2889851
doct		.0203036	.0702322	0.29	0.773	-.1175923	.1581994
dnov		.1501365	.0652729	2.30	0.022	.0219778	.2782952
ddec		.3887553	.0640009	6.07	0.000	.2630941	.5144164
_cons		-.1259374	.0849331	-1.48	0.139	-.2926974	.0408226

⁴ This amount is estimated as 4.4% of the monthly average of gross real tax receipts per fishing town. The total sum of averages for the fishing towns was \$13,147,250 in 1982-1984 US dollars. This divided by 24 gives the monthly average of \$547,802, and divided by 29 given the monthly average per town of approximately \$18,889, multiplying by 0.044 gives \$831 in 1982-1984 dollars. For 1994-1995, the CPI index is about 1.5 larger giving the values in the text. For 2005 dollars, the CPI index is about 2 times larger giving a point estimate of \$1,662, with upper bound of \$2,890 and lower bound of \$472 approximately.

⁵ This amount is estimated as 2.8% of the monthly total of gross real tax receipts per fishing town as above. For 2005 dollars, gives a point estimate of \$1,647,000 with upper bound of \$3,588,000 and lower bound of \$2,353,000 approximately.

Table AD.1.15 cont.

Linear regression						Number of obs = 700	
						F(13, 686) = 3366.58	
						Prob > F = 0.0000	
						R-squared = 0.9804	
						Root MSE = .21926	

		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

lrtotst1							
L1.		.9915491	.0053815	184.25	0.000	.980983	1.002115
dfyear		-.028359	.0165113	-1.72	0.086	-.0607778	.0040597
dfeb		.2787191	.0564905	4.93	0.000	.1678041	.389634
dmar		.3533206	.0571365	6.18	0.000	.2411372	.465504
dapr		.5277049	.0595489	8.86	0.000	.4107849	.644625
dmay		.4055934	.0579141	7.00	0.000	.2918832	.5193037
djun		.4088393	.0572961	7.14	0.000	.2963424	.5213362
djul		.4259688	.0623036	6.84	0.000	.3036401	.5482975
daug		.3203066	.0558726	5.73	0.000	.2106048	.4300085
dsep		.1581025	.0606941	2.60	0.009	.0389339	.2772711
doct		.0122932	.0703206	0.17	0.861	-.1257763	.1503626
dnov		.149083	.0645999	2.31	0.021	.0222457	.2759204
ddec		.3913467	.0634142	6.17	0.000	.2668374	.515856
_cons		-.198506	.0854939	-2.32	0.021	-.3663672	-.0306449

Finally, in reference to Tables 2.12-2.13 in Baden and Bianconi (2006), Table AD.1.16 includes the lagged dependent variable and we find no statistically significant effect of Amendment 5 on Gloucester, New Bedford and Plymouth relative to all other fishing communities in the State.

Table AD.1.16

Linear regression						Number of obs = 700	
						F(19, 680) = 732.99	
						Prob > F = 0.0000	
						R-squared = 0.9388	
						Root MSE = .2193	

	lrmst1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	lrmst1						
	L1.	.9608389	.0125488	76.57	0.000	.9361999	.985478
	dfyear	-.0476221	.0172572	-2.76	0.006	-.081506	-.0137383
	dglouce	.0589188	.0956117	0.62	0.538	-.1288107	.2466484
	intgloufyear	-.0019097	.148486	-0.01	0.990	-.2934558	.2896365
	dnbed	.0395326	.0478359	0.83	0.409	-.0543913	.1334565
	intnbfyear	.0069673	.0561831	0.12	0.901	-.1033459	.1172805
	dply	-.0106288	.0318861	-0.33	0.739	-.0732358	.0519783
	intplyfyear	.0421625	.045472	0.93	0.354	-.04712	.1314449
	dfeb	.268958	.0581126	4.63	0.000	.1548563	.3830596
	dmar	.3428945	.0587735	5.83	0.000	.227495	.4582939
	dapr	.5181697	.0604235	8.58	0.000	.3995307	.6368087
	dmay	.4102986	.0581794	7.05	0.000	.2960658	.5245315
	djun	.4166578	.0576061	7.23	0.000	.3035506	.529765
	djul	.4369749	.0624419	7.00	0.000	.3143728	.559577
	daug	.334996	.0561037	5.97	0.000	.2248387	.4451533
	dsep	.1732009	.0605627	2.86	0.004	.0542885	.2921133
	doct	.0228222	.0698529	0.33	0.744	-.114331	.1599754
	dnov	.1505928	.0652632	2.31	0.021	.0224512	.2787343
	ddec	.3881277	.0641718	6.05	0.000	.2621291	.5141263
	_cons	-.0946608	.0894755	-1.06	0.290	-.2703423	.0810206

Linear regression						Number of obs = 700	
						F(19, 680) = 2620.58	
						Prob > F = 0.0000	
						R-squared = 0.9804	
						Root MSE = .22014	

	lrtotst1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

doct		.0126766	.0704118	0.18	0.857	-.125574	.1509272
dnov		.1491525	.0647254	2.30	0.022	.0220668	.2762382
ddec		.3912512	.0636576	6.15	0.000	.2662622	.5162402
_cons		-.187836	.0883394	-2.13	0.034	-.3612868	-.0143853

In summary, with the inclusion of lagged effects on sales tax receipts the main qualitative conclusions of Baden and Bianconi (2006) regarding Amendment 5 were not substantively altered: From a quantitative perspective, the best evidence is from Table AD.1.15 [Table 2.11 in Baden and Bianconi (2006)] where the introduction of the amendment is shown to reduce the average sales tax receipts of all fishing towns by about an upper bound of negative 7.6% to a lower bound of 1.2%, a much smaller effect than previously obtained, but still statistically significant.

AD.I.2. Amendment 7: July 1995 to July 1997 Period

First, we provide time series analysis and intervention analysis for the amendment, in Boston, Gloucester, New Bedford, Fairhaven, Plymouth, and Marshfield.

Boston

Figure AD.1.13 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment on July 1, 1996. No specific pattern is observed for Boston.

Figure AD1.13

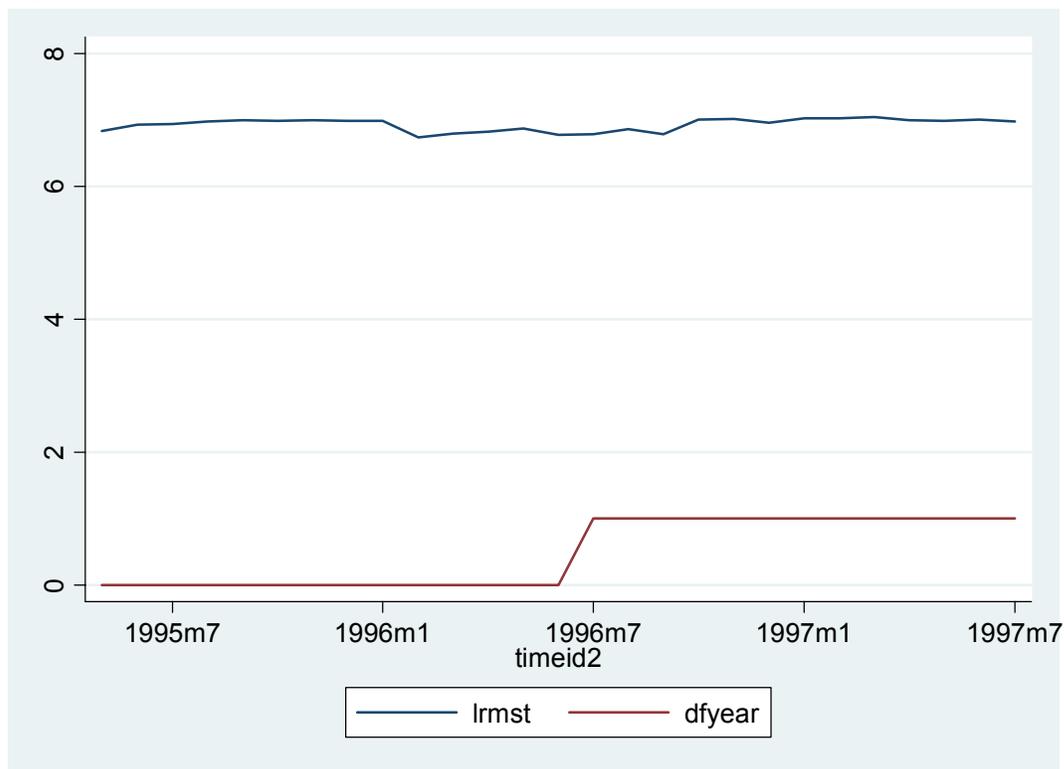


Table AD.1.16 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a positive effect, but not statistically significant for Boston.

Table AD.1.16

Linear regression		Number of obs =		26		
		F(2, 23) =		23.75		
		Prob > F =		0.0000		
		R-squared =		0.4270		
		Root MSE =		.07397		
lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrmst						
L1.	.5912233	.1625185	3.64	0.001	.2550282	.9274183
dfyear	.0321508	.0326564	0.98	0.335	-.035404	.0997056
_cons	2.822628	1.116103	2.53	0.019	.5137938	5.131463

Figure AD.1.14 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The dip in the year 1996 for the gross sales tax receipts in Boston is due to tax changes unrelated to Amendment 5.

Figure AD.1.14

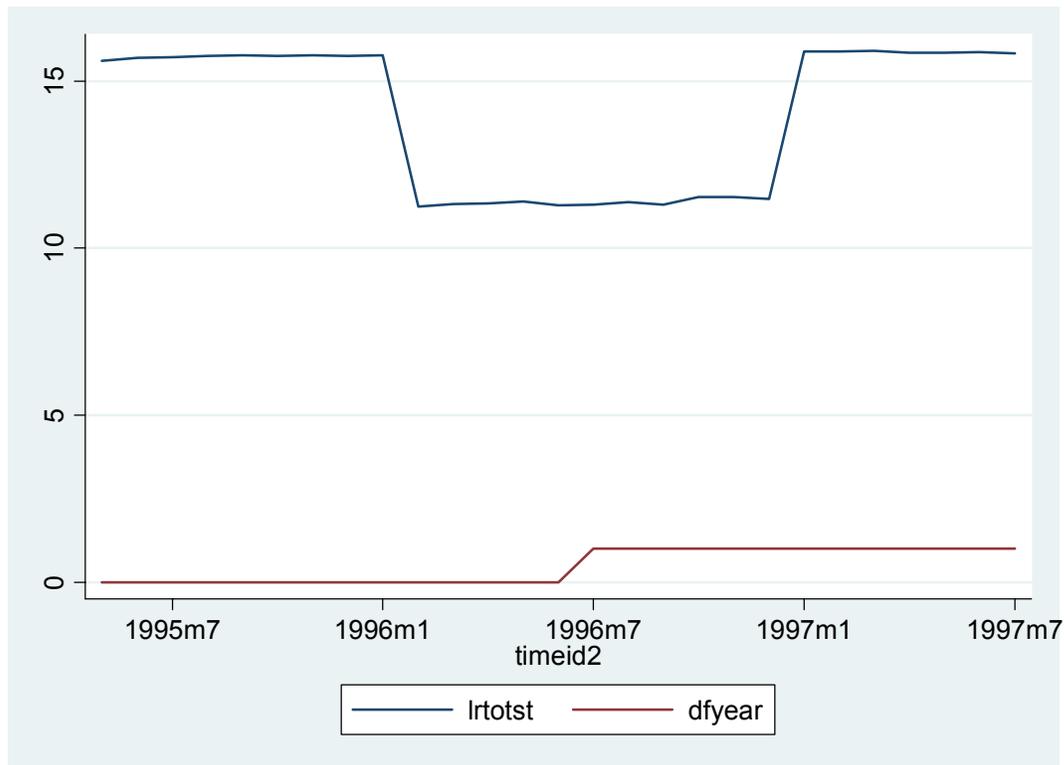


Table AD.1.17 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a positive effect, but not statistically significant.

Table AD.1.17

Linear regression		Number of obs =		26		
		F(2, 23) =		1024.91		
		Prob > F =		0.0000		
		R-squared =		0.7197		
		Root MSE =		1.2348		

		Robust				
		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]

lrtotst						
L1.		.8687836	.0907512	9.57	0.000	.6810504 1.056517
dfyear		.5619575	.4034394	1.39	0.177	-.2726205 1.396536
_cons		1.554684	1.135855	1.37	0.184	-.7950106 3.904378

Gloucester

Figure AD.1.15 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The dip in the year 1996 for the average and total gross sales tax receipts in Gloucester is due to tax changes unrelated to Amendment 5.

Figure AD.1.15

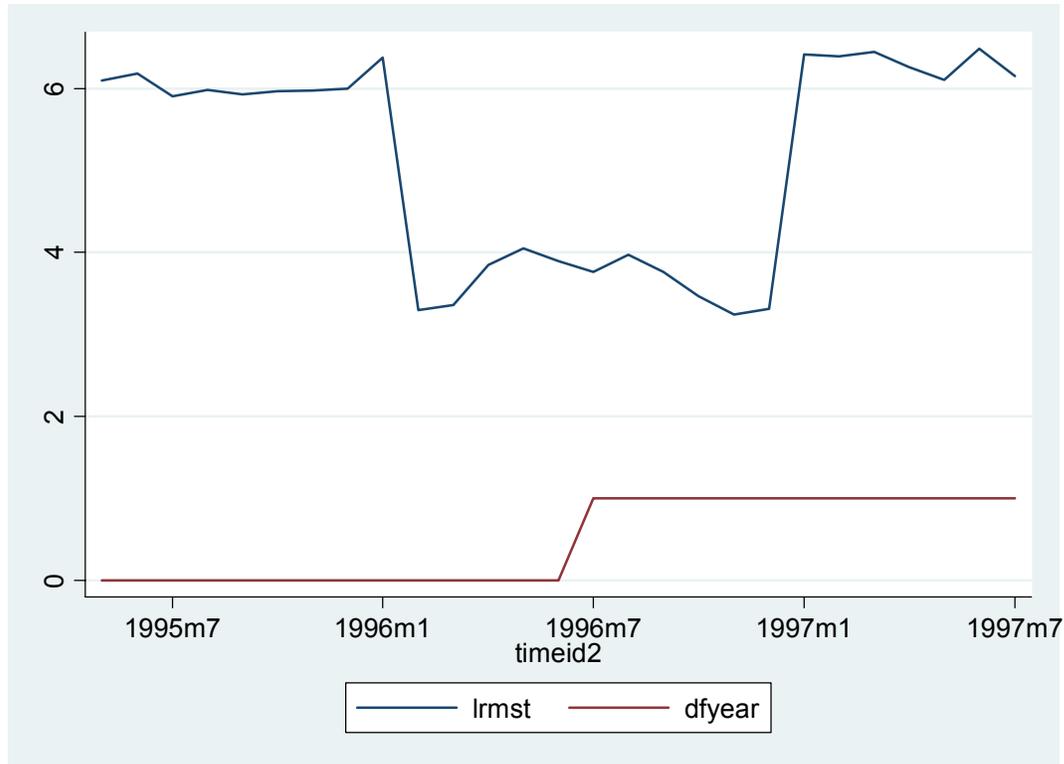


Table AD.1.16 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a positive effect, but not statistically significant.

Table AD.1.16

Linear regression							Number of obs =	26
							F(2, 23) =	86.30
							Prob > F =	0.0000
							R-squared =	0.5885
							Root MSE =	.87167

	lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]		

	lrmst							
	L1.	.7780136	.1398916	5.56	0.000	.4886258	1.067401	
	dfyear	.2510289	.2894342	0.87	0.395	-.3477114	.8497691	
	_cons	1.007527	.6502208	1.55	0.135	-.3375568	2.352612	

Figure AD.1.16 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The dip in the year 1996 for the average and total gross sales tax receipts in Gloucester is due to tax changes unrelated to Amendment 5.

Figure AD.1.16

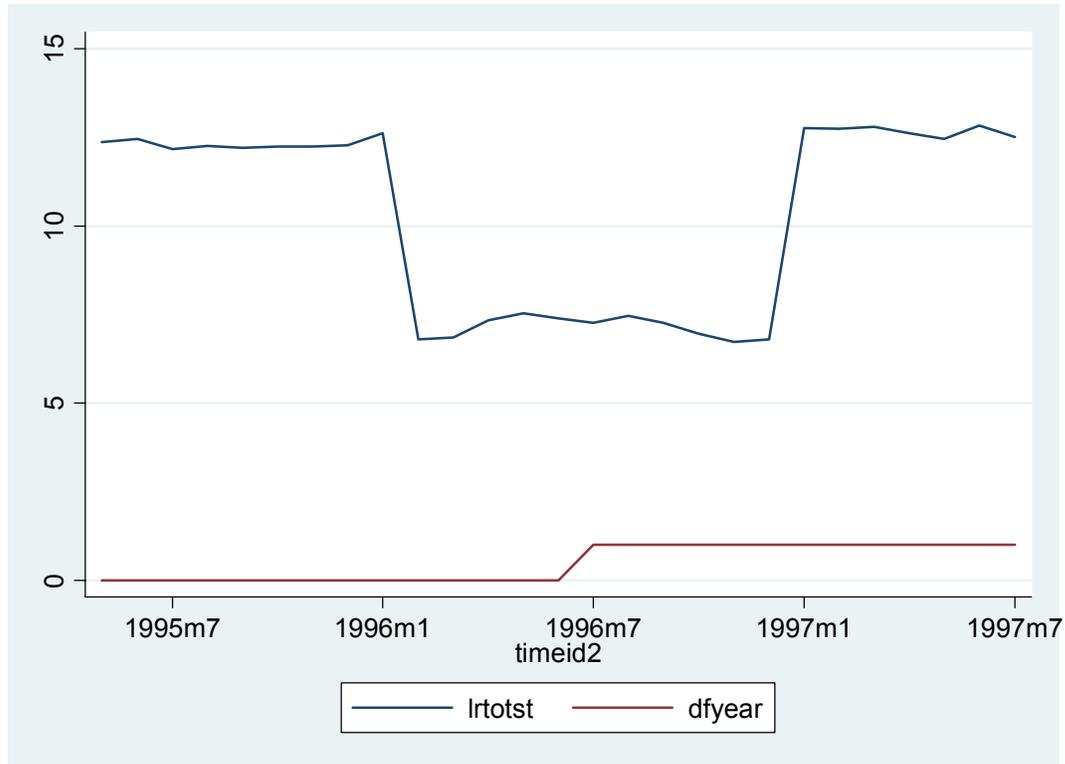


Table AD.1.17 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a positive effect, but not statistically significant.

Table AD.1.17

Linear regression		Number of obs =		26		
		F(2, 23) =		281.16		
		Prob > F		= 0.0000		
		R-squared		= 0.6641		
		Root MSE		= 1.6366		

		Robust				
lrtotst		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]

lrtotst						
L1.		.830398	.1143152	7.26	0.000	.5939191 1.066877
dfyear		.6039561	.5313726	1.14	0.267	-.495272 1.703184
_cons		1.434658	1.058521	1.36	0.188	-.7550583 3.624375

New Bedford

Figure AD.1.17 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The time series pattern is slightly upwards during the period.

Figure AD.1.17

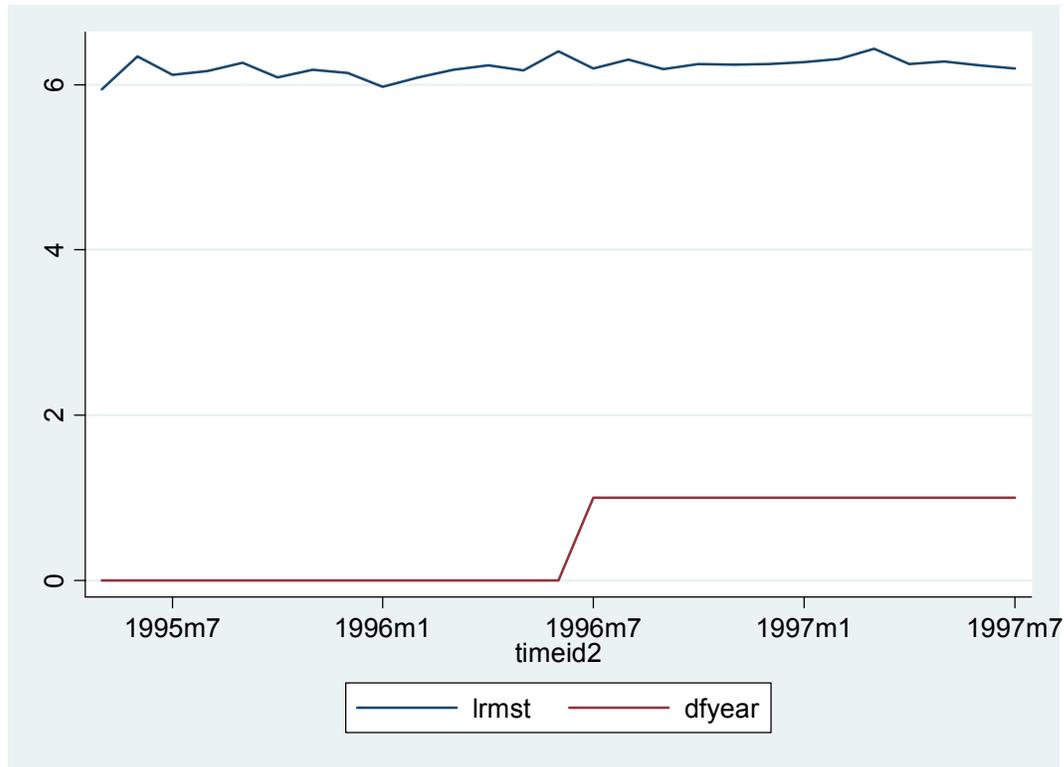


Table AD.1.18 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find some evidence of a positive effect, about 11% increase after the amendment, which is statistically significant.

Table AD.1.18

Linear regression						
					Number of obs =	26
					F(2, 23) =	3.90
					Prob > F	0.0347
					R-squared	0.2092
					Root MSE	.09058

	lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

	lrmst					
	L1.	-.1761298	.1905083	-0.92	0.365	-.5702263 .2179667
	dfyear	.1054451	.0379414	2.78	0.011	.0269572 .1839329
	_cons	7.266585	1.178586	6.17	0.000	4.828496 9.704675

Figure AD.1.18 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The time series pattern is slightly upwards during the period.

Figure AD.1.18

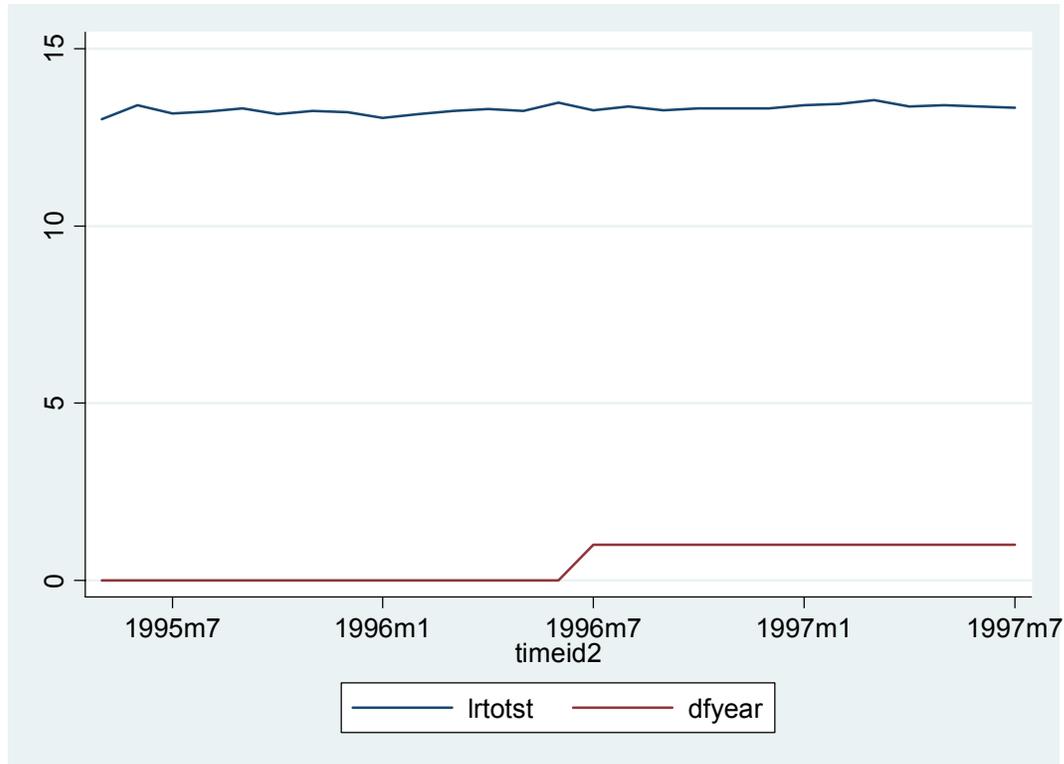


Table AD.1.19 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find some evidence of a positive effect, about 13% increase after the amendment, which is statistically significant.

Table AD.1.19

Linear regression		Number of obs =		26		
		F(2, 23) =		5.28		
		Prob > F =		0.0130		
		R-squared =		0.2881		
		Root MSE =		.09811		

		Robust				
lrtotst		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]

lrtotst						
L1.		-.0555895	.1987261	-0.28	0.782	-.4666857 .3555067
dfyear		.1260532	.0431399	2.92	0.008	.0368115 .2152949
_cons		13.98172	2.630406	5.32	0.000	8.54031 19.42313

Note that in Tables AD.1.18, and AD.1.19, the lagged dependent variable is not statistically significant. We also estimated models without lagged dependent variables and obtained similar results in Table AD.1.19a .

Table AD.1.19a

Linear regression	Number of obs =	27
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							F(1, 25) = 6.98
							Prob > F = 0.0140
							R-squared = 0.2110
							Root MSE = .09941

lrmst		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
dfyear		.0989933	.037462	2.64	0.014	.0218389 .1761478	
_cons		6.166722	.033144	186.06	0.000	6.09846 6.234983	

Linear regression							Number of obs = 27
							F(1, 25) = 11.35
							Prob > F = 0.0025
							R-squared = 0.3054
							Root MSE = .10501

lrtotst		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
dfyear		.1341148	.0398166	3.37	0.002	.052111 .2161187	
_cons		13.23013	.0332401	398.02	0.000	13.16167 13.29859	

Fairhaven

Figure AD.1.19 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The time series pattern now shows a hump in the year 1996, unrelated to the amendment.

Figure AD.1.19

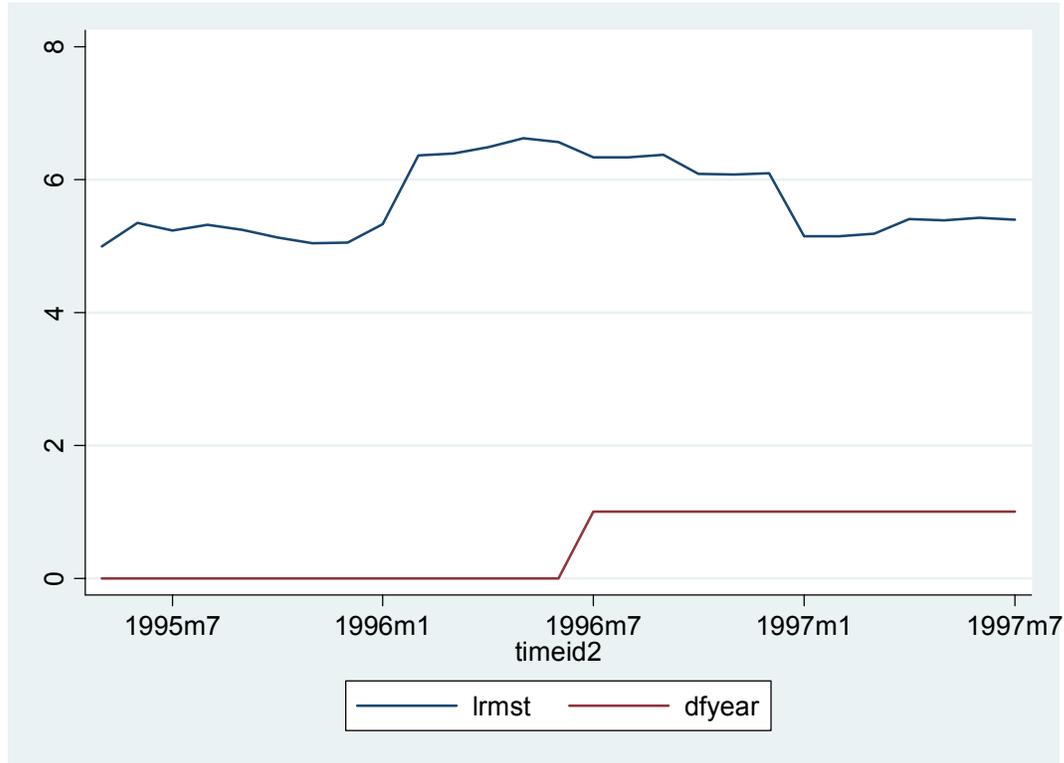


Table AD.1.20 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find some evidence of a negative effect, about 18% decline after the amendment, which is statistically significant at the 5% level.

Table AD.1.20

Linear regression		Number of obs = 26		F(2, 23) = 234.01		Prob > F = 0.0000		R-squared = 0.7551		Root MSE = .29296	
		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]					
lrmst											
lrmst	L1.	.8655349	.0611944	14.14	0.000	.7389447	.9921252				
dfyear		-.1795581	.1044767	-1.72	0.099	-.3956846	.0365685				
_cons		.8715897	.4001017	2.18	0.040	.0439162	1.699263				

Figure AD.1.20 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The time series pattern now shows a hump in the year 1996, unrelated to the amendment.

Figure AD.1.20

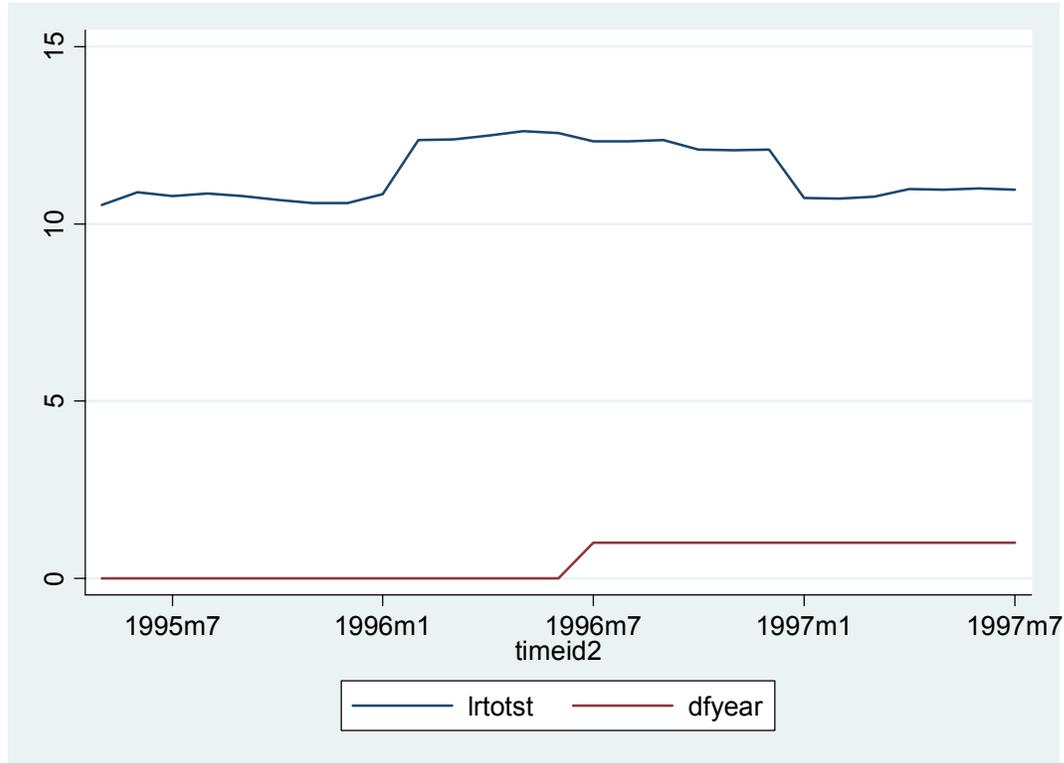


Table AD.1.21 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, about 23% decline after the amendment, but only statistically significant at the 6% level.

Table AD.1.21

Linear regression						
						Number of obs = 26
						F(2, 23) = 387.62
						Prob > F = 0.0000
						R-squared = 0.7492
						Root MSE = .41032

		Robust				
	lrtotst	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]

	lrtotst					
	L1.	.8701785	.0665211	13.08	0.000	.732569 1.007788
	dfyear	-.2327624	.1404679	-1.66	0.111	-.5233423 .0578175
	_cons	1.618184	.8350104	1.94	0.065	-.1091667 3.345535

Plymouth

Figure AD.1.21 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The time series pattern shows a slight upward trend.

Figure AD.1.21

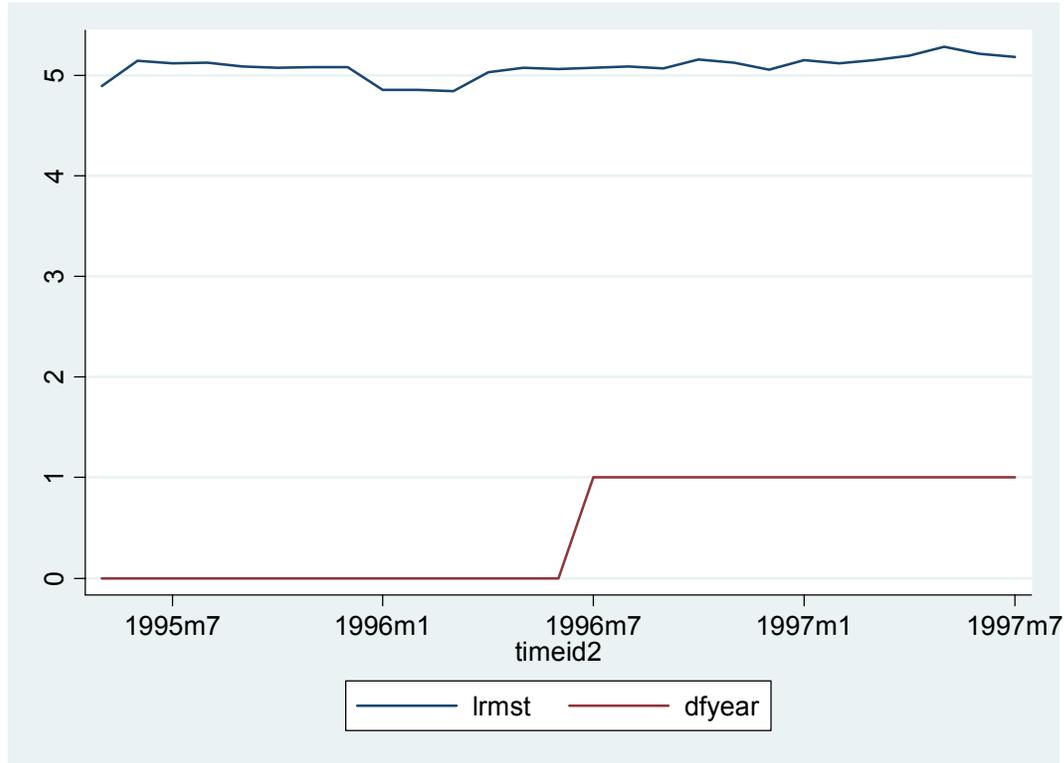


Table AD.1.22 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find some evidence of a positive effect, about 5% increase after the amendment, which is statistically significant at the 4.5% level.

Table AD.1.22

Linear regression						
					Number of obs =	26
					F(2, 23) =	7.55
					Prob > F =	0.0030
					R-squared =	0.4963
					Root MSE =	.07731

		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

lrmst						
lrmst						
L1.		.5045606	.1975244	2.55	0.018	.0959503 .913171
dfyear		.0537097	.0299753	1.79	0.086	-.0082989 .1157182
_cons		2.498752	1.00285	2.49	0.020	.4241989 4.573306

Figure AD.1.22 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The time series pattern shows a slight upward trend.

Figure AD.1.22

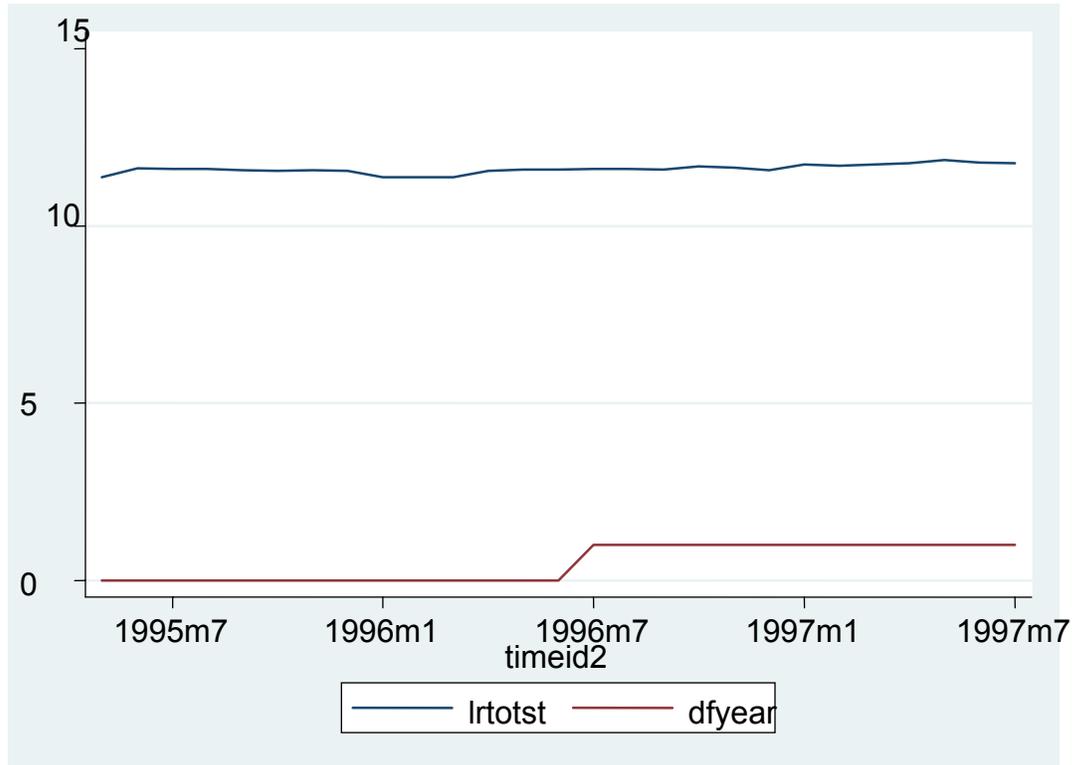


Table AD.1.23 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find some evidence of a positive effect, about 7.5% increase after the amendment, which is statistically significant.

Table AD.1.23

Linear regression		Number of obs = 26				
		F(2, 23) = 15.24				
		Prob > F = 0.0001				
		R-squared = 0.6383				
		Root MSE = .07644				

		Robust				
	lrtotst	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]

	l1	.5334875	.1811422	2.95	0.007	.1587664 .9082086
	dfyear	.0749162	.0330077	2.27	0.033	.0066346 .1431979
	_cons	5.390903	2.099974	2.57	0.017	1.046776 9.73503

Marshfield

Figure AD.1.23 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The time series pattern is oscillating in this case.

Figure AD.1.23

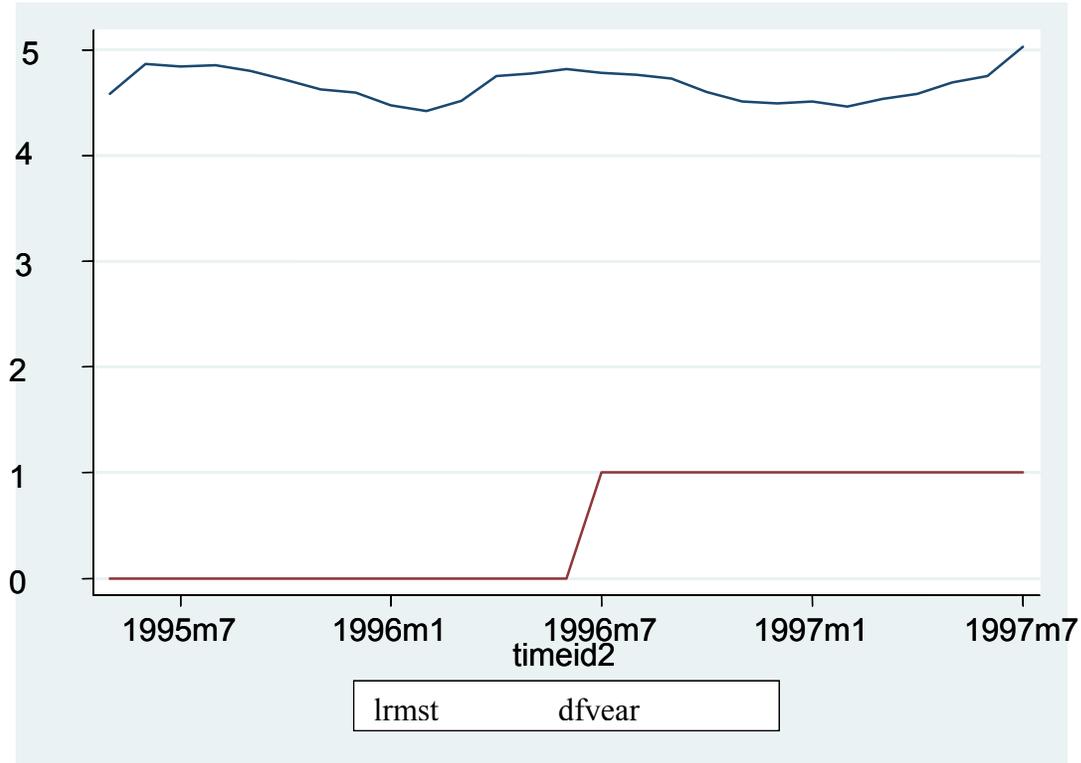


Table AD.1.24 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, but not statistically significant.

Table AD.1.24

Linear regression							Number of obs =	26
							F(2, 23) =	42.15
							Prob > F	= 0.0000
							R-squared	= 0.5285
							Root MSE	= .11148

	lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]		

	lrmst							
	L1.	.8080046	.1410264	5.73	0.000	.5162692	1.09974	
	dfyear	-.0105543	.0484963	-0.22	0.830	-.1108765	.0897679	
	_cons	.9164949	.6795051	1.35	0.191	-.4891685	2.322158	

Figure AD.1.24 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in July 1996. The time series pattern is again oscillating in this case.

Figure AD.1.24

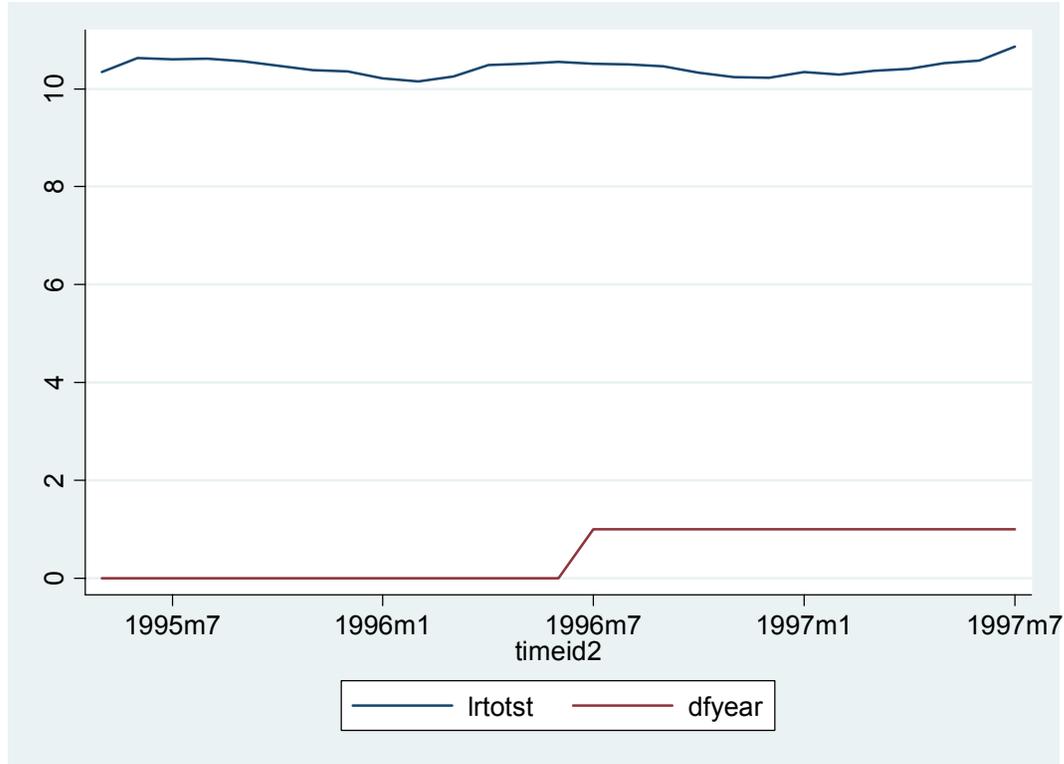


Table AD.1.25 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We do not find any statistically significant effect in this case.

Table AD.1.25

Linear regression		Number of obs =		26		
		F(2, 23) =		24.21		
		Prob > F =		0.0000		
		R-squared =		0.5245		
		Root MSE =		.1159		

			Robust			
lrtotst		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]

lrtotst						
L1.		.8488731	.1654079	5.13	0.000	.5067009 1.191045
dfyear		.0053244	.0472709	0.11	0.911	-.0924628 .1031116
_cons		1.592564	1.742665	0.91	0.370	-2.012414 5.197541

In summary, the time series intervention evidence for Amendment 7 is as follows:

Table AD.1.25a

	Policy Intervention Variable	Statistically Significant
Boston	Positive	No
Gloucester	Positive	No
New Bedford	Positive	Yes
Fairhaven	Negative	Yes
Plymouth	Positive	Yes
Marshfield	Inconclusive	No

AD.I.2.1. Amendment 7: July 1995 to July 1997 Period - Autocorrelation

We proceed now with estimates of the regressions in Baden and Bianconi (2006), but including lagged dependent variables, and with other assumptions as well.

In reference to Table 2.14 in Baden and Bianconi (2006), Table AD.1.26 shows the simple differences-in-differences estimator without further factors, fixed effects or controls, but with a lagged dependent variable. Note that original results are not changed in any dramatic fashion, but autocorrelation of taxes is statistically significant, with two month history in this case.⁶ For the average and total gross real sales receipts and the total gross real sales receipts we find no statistically significant effects of Amendment 7.

Table AD.1.26

Linear regression		Number of obs = 8355					
		F(5, 8349) =11109.16					
		Prob > F = 0.0000					
		R-squared = 0.9154					
		Root MSE = .3327					

		Robust					
lrmst	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]		

lrmst							
L1.	.8947312	.0149434	59.87	0.000	.8654385	.9240239	
L2.	.0618954	.0142605	4.34	0.000	.0339412	.0898496	
dfyear	-.0029536	.0075332	-0.39	0.695	-.0177205	.0118133	
dtreat	.004257	.0215278	0.20	0.843	-.0379428	.0464568	
dyeartreat	-.012471	.0297305	-0.42	0.675	-.0707501	.0458081	
_cons	.229948	.0227008	10.13	0.000	.1854487	.2744472	

⁶ We have performed several other econometric models and tests of fixed and random effects, available upon request. Results are unchanged.

Table AD.1.26 cont.

Linear regression						Number of obs =	8355
						F(5, 8349) =	17719.58
						Prob > F =	0.0000
						R-squared =	0.9435
						Root MSE =	.49327

		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

lrtotst							
L1.		.9410517	.0096248	97.77	0.000	.9221847	.9599187
L2.		.0290144	.009224	3.15	0.002	.0109331	.0470956
dfyear		.0031424	.0112291	0.28	0.780	-.0188694	.0251541
dtreat		.000656	.0311723	0.02	0.983	-.0604495	.0617615
dyeartreat		.0382929	.0428382	0.89	0.371	-.0456806	.1222663
_cons		.3124041	.0354667	8.81	0.000	.2428806	.3819277

In reference to Tables 2.19-2.22 in Baden and Bianconi (2006), we present in Table AD.1.27 models with separate difference-in-difference estimators for each town of the fishing industry group, reflecting changes in each town relative to all other towns in the state, excluding other fishing towns, and including lagged dependent variables. In Table AD.1.27, we note that the results previously obtained change only slightly reflecting the positive trend of sales taxes in the state during this period.

Table AD.1.27

Linear regression						Number of obs =	8697
						F(81, 8615) =	897.09
						Prob > F =	0.0000
						R-squared =	0.9158
						Root MSE =	.33369

		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

lrmst							
L1.		.9482882	.0041576	228.09	0.000	.9401383	.9564381
dfyear		-.016841	.0075533	-2.23	0.026	-.0316473	-.0020348
dtreat		-.0275852	.1520638	-0.18	0.856	-.3256667	.2704962
dyeartreat		-.2505017	.160537	-1.56	0.119	-.5651926	.0641892
dfeb		.0190607	.0390504	0.49	0.625	-.0574873	.0956088
dmar		.0965384	.0285027	3.39	0.001	.0406662	.1524105
dapr		.1589843	.0294699	5.39	0.000	.1012164	.2167523
dmay		.1104796	.0283961	3.89	0.000	.0548163	.1661428
djun		.1841384	.0286515	6.43	0.000	.1279746	.2403022
djul		.0577552	.0293241	1.97	0.049	.000273	.1152375
daug		.0851464	.0283838	3.00	0.003	.0295073	.1407855
dsep		.0462557	.0290676	1.59	0.112	-.0107238	.1032351
doct		.0331637	.0300298	1.10	0.269	-.025702	.0920294
dnov		.0323544	.0285142	1.13	0.257	-.0235403	.0882492
ddec		.076961	.0287898	2.67	0.008	.0205261	.1333959
dboston		-.1052317	.124483	-0.85	0.398	-.3492482	.1387848
dchatam		-.1681018	.1283599	-1.31	0.190	-.419718	.0835143
dfairhaven		-.0491751	.1416893	-0.35	0.729	-.32692	.2285698
dgloucester		-.3545473	.272042	-1.30	0.193	-.8878147	.1787201
dmarshfield		-.1989392	.1187687	-1.68	0.094	-.4317543	.0338759
dnewbedford		-.1056484	.1222705	-0.86	0.388	-.3453279	.134031

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dplymouth	-.1863631	.1187015	-1.57	0.116	-.4190466	.0463204
dprovincet~n	-.1490707	.1374178	-1.08	0.278	-.4184424	.120301
dscituate	-.2287004	.1191765	-1.92	0.055	-.4623149	.004914
dbeverly	-.0030692	.0928015	-0.03	0.974	-.1849823	.178844
deatham	.0784035	.2921083	0.27	0.788	-.4941987	.6510056
dedgardtown	-.0828296	.1963572	-0.42	0.673	-.4677368	.3020776
dharwich	-.2104031	.1220759	-1.72	0.085	-.4497011	.0288949
dwestport	-.1982855	.1186243	-1.67	0.095	-.4308176	.0342465
dmarblehead	-.1677714	.1226726	-1.37	0.171	-.408239	.0726962
dnantucket	-.1631159	.1247208	-1.31	0.191	-.4075986	.0813668
dnewburyport	-.1169423	.1325065	-0.88	0.378	-.3766867	.142802
drockport	-.2137384	.1304708	-1.64	0.101	-.4694924	.0420156
dsandwich	-.244809	.196418	-1.25	0.213	-.6298352	.1402173
dswamp	-.201532	.1205379	-1.67	0.095	-.4378151	.0347512
dwellfleet	-.200555	.1474076	-1.36	0.174	-.4895092	.0883991
dyarmouth	-.158799	.121361	-1.31	0.191	-.3966955	.0790976
dsalem	-.1680783	.1211131	-1.39	0.165	-.405489	.0693324
dmanchester	-.1899288	.1195972	-1.59	0.112	-.424368	.0445104
dorleans	-.1695441	.1188477	-1.43	0.154	-.4025141	.0634259
doakb	-.1642016	.14164	-1.16	0.246	-.44185	.1134468
dtisbury	-.1555063	.1212759	-1.28	0.200	-.3932361	.0822235
ddartmouth	-.058179	.1578495	-0.37	0.712	-.3676018	.2512437
dtboston	.295304	.1729244	1.71	0.088	-.0436693	.6342772
dtchatam	.2162007	.1770671	1.22	0.222	-.1308932	.5632946
dtfairhaven	.0753425	.1885943	0.40	0.690	-.2943474	.4450325
dtgloucester	.5964214	.3884933	1.54	0.125	-.1651186	1.357961
dtmarshfield	.2700214	.1655045	1.63	0.103	-.0544071	.5944499
dtnewbedford	.2302353	.1719965	1.34	0.181	-.1069189	.5673895
dtplymouth	.2765608	.1678077	1.65	0.099	-.0523825	.605504
dtprovince~n	.2375492	.1848817	1.28	0.199	-.1248632	.5999617
dtscituate	.2926132	.1665843	1.76	0.079	-.0339318	.6191582
dteatham	-.1393853	.4341623	-0.32	0.748	-.9904474	.7116768
dtedgardtown	.1262386	.2601978	0.49	0.628	-.3838114	.6362886
dtharwich	.2976411	.169414	1.76	0.079	-.0344508	.629733
dtwestport	.2570216	.1635436	1.57	0.116	-.063563	.5776063
dtmarblehead	.2238023	.1703344	1.31	0.189	-.110094	.5576985
dtnantucket	.2728596	.1729034	1.58	0.115	-.0660725	.6117917
dtnewburyp~t	.2059948	.1859135	1.11	0.268	-.1584402	.5704298
dtrockport	.2531954	.175777	1.44	0.150	-.0913697	.5977604
dtsandwich	.3206174	.2701631	1.19	0.235	-.208967	.8502018
dtswamp	.2367497	.1687364	1.40	0.161	-.094014	.5675135
dtwellfleet	.2558057	.2031116	1.26	0.208	-.1423418	.6539531
dtymouth	.2707676	.16934	1.60	0.110	-.0611793	.6027145
dtalem	.2525934	.1705466	1.48	0.139	-.0817187	.5869056
dtmanchester	.2578035	.1731347	1.49	0.137	-.081582	.5971891
dtorleans	.2579545	.1640459	1.57	0.116	-.0636146	.5795237
dtoakb	.2456666	.1978086	1.24	0.214	-.1420856	.6334188
dtisbury	.2739271	.1681411	1.63	0.103	-.0556698	.6035239
dtdartmouth	.1299687	.2089198	0.62	0.534	-.2795642	.5395015
dtreatmay	.2404329	.1101118	2.18	0.029	.0245875	.4562783
dtreatjun	.2594405	.1094439	2.37	0.018	.0449042	.4739768
dtreatjul	.297658	.114957	2.59	0.010	.0723147	.5230013
dtreataug	.199477	.1092585	1.83	0.068	-.0146957	.4136498
dtreatsep	.0763594	.1126448	0.68	0.498	-.1444513	.2971701
dtreatoct	.0630099	.1174081	0.54	0.592	-.167138	.2931578
dtreatnov	.1444138	.1108582	1.30	0.193	-.0728948	.3617224
dtreatdec	.1940715	.1104204	1.76	0.079	-.0223788	.4105219
dtreatfeb	.3205467	.1448424	2.21	0.027	.0366208	.6044725
dtreatmar	.1790839	.1093451	1.64	0.102	-.0352587	.3934265
dtreatapr	.2353698	.1151033	2.04	0.041	.0097397	.4609998
_cons	.2129967	.0354727	6.00	0.000	.1434618	.2825316

Table AD.1.27 cont.

Linear regression					Number of obs = 8697	
					F(81, 8615) = 1554.08	
					Prob > F = 0.0000	
					R-squared = 0.9449	
					Root MSE = .48968	

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

	lrtotst					
	L1.	.9673392	.003182	304.00	0.000	.9611017 .9735768
	dfyear	-.0118931	.0111832	-1.06	0.288	-.0338149 .0100287
	dtreat	-.0061984	.176387	-0.04	0.972	-.3519592 .3395624
	dyeartreat	.0668164	.1029086	0.65	0.516	-.1349091 .2685418
	dfeb	-.0078312	.0647578	-0.12	0.904	-.134772 .1191096
	dmar	.0636088	.0458731	1.39	0.166	-.0263134 .1535311
	dapr	.1253166	.0464884	2.70	0.007	.0341881 .216445
	dmay	.0749224	.0457983	1.64	0.102	-.0148531 .1646979
	djun	.1503679	.0452666	3.32	0.001	.0616345 .2391012
	djul	.0199459	.0470877	0.42	0.672	-.0723572 .112249
	daug	.0481124	.0457741	1.05	0.293	-.0416159 .1378407
	dsep	.0088529	.0461881	0.19	0.848	-.0816868 .0993926
	doct	-.0038477	.0468341	-0.08	0.935	-.0956537 .0879584
	dnov	-.0040364	.0458693	-0.09	0.930	-.0939512 .0858784
	ddec	.041174	.0460312	0.89	0.371	-.0490582 .1314063
	dboston	-.2458017	.3393882	-0.72	0.469	-.9110838 .4194803
	dchatam	-.0356462	.0954167	-0.37	0.709	-.2226857 .1513934
	dfairhaven	.1402311	.1435819	0.98	0.329	-.1412238 .4216861
	dgloucester	-.4170012	.4388835	-0.95	0.342	-1.277318 .4433156
	dmarshfield	-.026918	.0816722	-0.33	0.742	-.187015 .133179
	dnewbedford	.0837095	.0885925	0.94	0.345	-.089953 .257372
	dplymouth	.0089594	.0829666	0.11	0.914	-.1536751 .1715938
	dprovincetown	.0297076	.1074413	0.28	0.782	-.180903 .2403183
	dscituate	-.0578512	.08753	-0.66	0.509	-.2294309 .1137285
	dbeverly	.0082654	.2163865	0.04	0.970	-.415904 .4324348
	deastham	.3047961	.3578657	0.85	0.394	-.3967064 1.006299
	dedgartown	.1305805	.2253241	0.58	0.562	-.3111087 .5722696
	dharwich	-.1101575	.102071	-1.08	0.281	-.3102411 .0899262
	dwestport	-.0211736	.0808759	-0.26	0.793	-.1797097 .1373625
	dmarblehead	.0146843	.0879306	0.17	0.867	-.1576809 .1870494
	dnantucket	.0194703	.090652	0.21	0.830	-.1582294 .19717
	dnewburyport	.0405832	.0971403	0.42	0.676	-.149835 .2310015
	drockport	-.0419096	.1075886	-0.39	0.697	-.2528091 .1689898
	dsandwich	-.0683393	.1718742	-0.40	0.691	-.4052538 .2685753
	dswamp	-.0373036	.0854158	-0.44	0.662	-.2047391 .1301319
	dwellfleet	-.0363734	.1222474	-0.30	0.766	-.2760076 .2032608
	dyarmouth	.0155783	.0906195	0.17	0.864	-.1620577 .1932142
	dsalem	.0071732	.0876562	0.08	0.935	-.1646538 .1790002
	dmanchester	-.0399721	.0849846	-0.47	0.638	-.2065623 .1266181
	dorleans	-.0000715	.0842049	-0.00	0.999	-.1651332 .1649902
	doakb	-.0224733	.1178902	-0.19	0.849	-.2535664 .2086198
	dtisbury	.0053381	.0857136	0.06	0.950	-.162681 .1733572
	ddartmouth	.1182555	.148977	0.79	0.427	-.1737752 .4102861
	dtboston	.605537	.477565	1.27	0.205	-.3306047 1.541679
	dtchatam	-.0172885	.1307069	-0.13	0.895	-.2735054 .2389284
	dtfairhaven	-.3135127	.186682	-1.68	0.093	-.6794541 .0524288
	dtgloucester	.6972392	.6301808	1.11	0.269	-.538066 1.932544
	dtmarshfield	-.0392394	.1082209	-0.36	0.717	-.2513784 .1728996
	dtnewbedford	-.0879736	.1195699	-0.74	0.462	-.3223592 .1464121
	dtplymouth	-.0436674	.1120503	-0.39	0.697	-.2633128 .1759779

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dtprovince~n	-.0780596	.1366148	-0.57	0.568	-.3458574	.1897381
dtscituate	-.0114889	.1139693	-0.10	0.920	-.2348959	.2119182
dteastham	-.5971017	.5218788	-1.14	0.253	-1.620109	.4259057
dtedgardtown	-.2871988	.3002546	-0.96	0.339	-.8757697	.3013721
dtharwich	.1109989	.1503215	0.74	0.460	-.1836672	.4056651
dtwestport	-.0566638	.1058648	-0.54	0.592	-.2641841	.1508566
dtmarblehead	-.08423	.1165432	-0.72	0.470	-.3126824	.1442225
dtnantucket	-.0413389	.1209429	-0.34	0.733	-.2784159	.195738
dtnewburyp~t	-.0987766	.1360508	-0.73	0.468	-.3654687	.1679155
dtrockport	-.0630208	.1337877	-0.47	0.638	-.3252767	.1992351
dtsandwich	.0085781	.236886	0.04	0.971	-.4557753	.4729314
dtswamp	-.0806082	.116326	-0.69	0.488	-.3086349	.1474185
dtwellfleet	-.0609706	.1629169	-0.37	0.708	-.3803267	.2583854
dtymouth	-.0457852	.1216496	-0.38	0.707	-.2842475	.1926771
dtsealem	-.0601018	.119641	-0.50	0.615	-.2946268	.1744231
dtmanchester	-.0578276	.1198316	-0.48	0.629	-.2927262	.177071
dtorleans	-.055192	.1089223	-0.51	0.612	-.2687058	.1583219
dtoakb	-.0637219	.1576914	-0.40	0.686	-.3728348	.245391
dttisbury	-.0397855	.1135345	-0.35	0.726	-.2623404	.1827693
dtdartmouth	-.2004043	.1886795	-1.06	0.288	-.5702613	.1694528
dtreatmay	.074441	.1723753	0.43	0.666	-.2634559	.4123379
dtreatjun	.1020969	.1691821	0.60	0.546	-.2295405	.4337344
dtreatjul	.1229843	.1779827	0.69	0.490	-.2259043	.4718729
dtreataug	.0278946	.1718665	0.16	0.871	-.3090049	.3647942
dtreatsep	-.095241	.1745428	-0.55	0.585	-.4373867	.2469047
dtreatoct	-.1062627	.1773116	-0.60	0.549	-.4538359	.2413106
dtreatnov	-.0223965	.1731167	-0.13	0.897	-.3617466	.3169536
dtreatdec	.0280451	.1726882	0.16	0.871	-.3104652	.3665554
dtreatfeb	.0041307	.2347749	0.02	0.986	-.4560842	.4643457
dtreatmar	.0140555	.1717758	0.08	0.935	-.3226662	.3507772
dtreatapr	.0703974	.175925	0.40	0.689	-.2744577	.4152524
_cons	.3136446	.0561638	5.58	0.000	.2035502	.4237391

In reference to Tables 2-23-2.25 in Baden and Bianconi (2006), Table AD.1.28 presents equivalent regressions with two lags of the dependent variable. Results continue to show that Amendment 7 no significant impact on sales tax receipts, for all towns in the sample.

Table AD.1.28

Linear regression		Number of obs = 700				
		F(14, 685) = 140.53				
		Prob > F = 0.0000				
		R-squared = 0.8342				
		Root MSE = .35944				

lrmst1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

lrmst1						
L1.	.9610661	.0349878	27.47	0.000	.89237	1.029762
L2.	-.0583819	.0285163	-2.05	0.041	-.1143718	-.0023919
dfyear	-.0148705	.026716	-0.56	0.578	-.0673256	.0375847
dfeb	.3442359	.1465633	2.35	0.019	.0564685	.6320032
dmar	.264477	.1067553	2.48	0.013	.05487	.474084
dapr	.3888953	.1108338	3.51	0.000	.1712805	.60651
dmay	.3455933	.1077146	3.21	0.001	.134103	.5570837
djun	.3508203	.1104889	3.18	0.002	.1338827	.5677579
djul	.3515277	.1139129	3.09	0.002	.1278673	.575188
daug	.2955607	.1100827	2.68	0.007	.0794206	.5117007
dsep	.138962	.1134599	1.22	0.221	-.083809	.3617329

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Table AD.1.28 Cont.

doct		.1155303	.1167435	0.99	0.323	-.1136877	.3447484
dnov		.1898587	.1095369	1.73	0.083	-.0252097	.4049271
ddec		.2756072	.1081004	2.55	0.011	.0633592	.4878552
_cons		.2482427	.1528991	1.62	0.105	-.0519645	.5484499

Linear regression						Number of obs =	700
						F(14, 685) =	157.45
						Prob > F =	0.0000
						R-squared =	0.8637
						Root MSE =	.53891

			Robust				
lrtotst1		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

lrtotst1							
L1.		.9516706	.0252299	37.72	0.000	.9021333	1.001208
L2.		-.0274487	.0192158	-1.43	0.154	-.0651776	.0102803
dfyear		.0354684	.0412897	0.86	0.391	-.0456013	.116538
dfeb		-.003712	.2371943	-0.02	0.988	-.4694272	.4620032
dmar		.0756838	.1732492	0.44	0.662	-.2644794	.415847
dapr		.1930019	.175494	1.10	0.272	-.1515689	.5375727
dmay		.1500268	.1725216	0.87	0.385	-.1887078	.4887614
djun		.1514828	.1733448	0.87	0.382	-.1888668	.4918337
djul		.1478963	.1778629	0.83	0.406	-.2013257	.4971183
daug		.0883436	.1717518	0.51	0.607	-.2488795	.4255668
dsep		-.0710678	.1742004	-0.41	0.683	-.4130986	.270963
doct		-.0964692	.1767651	-0.55	0.585	-.4435356	.2505972
dnov		-.0192453	.1734913	-0.11	0.912	-.3598839	.3213932
ddec		.070824	.1731825	0.41	0.683	-.2692082	.4108562
_cons		.7537991	.3510737	2.15	0.032	.0644894	1.443109

Finally, in reference to Tables 2.24-2.25, we present in Table AD.1.29 results including lagged dependent variables. Results are practically unchanged as well.

Table AD.1.29

Linear regression						Number of obs =	700
						F(20, 679) =	229.48
						Prob > F =	0.0000
						R-squared =	0.8360
						Root MSE =	.3591

			Robust				
lrmst1		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

lrmst1							
L1.		.9501785	.0414349	22.93	0.000	.8688227	1.031534
L2.		-.0512124	.033713	-1.52	0.129	-.1174067	.0149819
dfyear		-.0292448	.0254185	-1.15	0.250	-.079153	.0206635
dglouce		-.1795685	.2708436	-0.66	0.508	-.7113602	.3522232
intgloufyear		.3329823	.3706692	0.90	0.369	-.3948132	1.060778
dnbed		.1008897	.0542654	1.86	0.063	-.0056585	.2074379
intnbfyear		.0134089	.0676648	0.20	0.843	-.1194485	.1462662
dply		-.0257251	.030131	-0.85	0.394	-.0848862	.033436
intplyfyear		.048667	.0530081	0.92	0.359	-.0554125	.1527465
dfeb		.3413644	.1455638	2.35	0.019	.0555552	.6271737

Table AD.1.29 Cont.

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dmar	.2645233	.1049882	2.52	0.012	.0583828	.4706638
dapr	.388577	.1098944	3.54	0.000	.1728033	.6043506
dmay	.3466807	.1064277	3.26	0.001	.1377137	.5556477
djun	.3519165	.109267	3.22	0.001	.1373746	.5664583
djul	.3536409	.1128805	3.13	0.002	.1320042	.5752776
daug	.2973103	.1097366	2.71	0.007	.0818466	.5127741
dsep	.1402366	.1135555	1.23	0.217	-.0827255	.3631986
doct	.1151052	.1173559	0.98	0.327	-.1153188	.3455292
dnov	.1886953	.1097678	1.72	0.086	-.0268299	.4042204
ddec	.2748107	.1076785	2.55	0.011	.0633878	.4862336
_cons	.2709838	.1499503	1.81	0.071	-.0234381	.5654058

Linear regression				Number of obs = 700		
				F(20, 679) = 459.39		
				Prob > F = 0.0000		
				R-squared = 0.8665		
				Root MSE = .53572		

		Robust				
lrtotst1	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

lrtotst1						
L1.	.9316355	.0373689	24.93	0.000	.858263	1.005008
L2.	-.0132344	.0264081	-0.50	0.616	-.0650858	.038617
dfyear	.0103784	.0374273	0.28	0.782	-.0631088	.0838657
dglouce	-.4346135	.4750047	-0.91	0.361	-1.367268	.498041
intgloufyear	.7157849	.6451185	1.11	0.268	-.5508821	1.982452
dnbed	.2041716	.0894727	2.28	0.023	.0284952	.3798481
intnbfyear	-.0223763	.0716309	-0.31	0.755	-.163021	.1182683
dply	.0568155	.0446909	1.27	0.204	-.0309333	.1445644
intplyfyear	.0118402	.055252	0.21	0.830	-.096645	.1203255
dfeb	-.0049687	.2329898	-0.02	0.983	-.4624358	.4524985
dmar	.0741212	.1714365	0.43	0.666	-.2624881	.4107305
dapr	.1928262	.1737161	1.11	0.267	-.1482592	.5339116
dmay	.1524012	.1702899	0.89	0.371	-.1819568	.4867592
djun	.1537175	.1713871	0.90	0.370	-.1827948	.4902299
djul	.1521271	.1762083	0.86	0.388	-.1938517	.4981058
daug	.0912979	.1712145	0.53	0.594	-.2448756	.4274713
dsep	-.069087	.1744134	-0.40	0.692	-.4115414	.2733675
doct	-.097637	.1777943	-0.55	0.583	-.4467296	.2514557
dnov	-.0216259	.1738468	-0.12	0.901	-.3629678	.319716
ddec	.0692839	.1724443	0.40	0.688	-.2693041	.407872
_cons	.8230721	.337338	2.44	0.015	.160721	1.485423

The conclusions from Baden and Bianconi (2006) are unchanged even when we take into account correlations across time in this case. For July 1995 to July 1997 period, the effects of Amendment 7 on average and total gross real sales taxes are not statistically significant. However, we note a general positive trend in sales tax receipts in this period, and this is reflected in some of the regressions above.

AD.I.3. Amendment 13: May 2003 to December 2004 Period

First, we provide time series analysis and intervention analysis for the amendment, in Boston, Gloucester, New Bedford, Fairhaven, Plymouth, and Marshfield.

Boston

Figure AD.1.25 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment on May 1, 2004. We observe a slight hump after the amendment.

Figure AD1.25

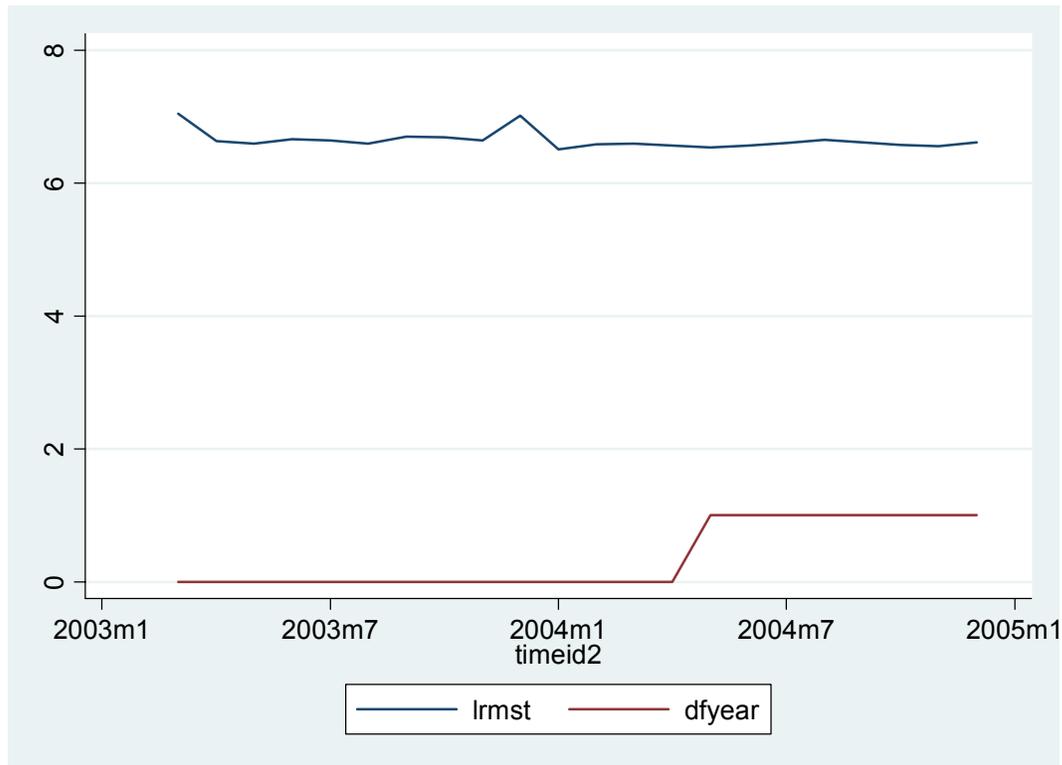


Table AD.1.30 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. The lagged value is not statistically significant in this case. We find evidence of a negative effect of the amendment, about 7.3% decline, with a 6% significance level.

Table AD.1.30

Linear regression						Number of obs =	21
						F(2, 18) =	1.42
						Prob > F =	0.2677
						R-squared =	0.1137
						Root MSE =	.10036

	lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	lrmst						
	L1.	-.136279	.1469531	-0.93	0.366	-.4450161	.1724581
	dfyear	-.0730222	.0433525	-1.68	0.109	-.1641023	.018058
	_cons	7.56156	.9927025	7.62	0.000	5.47597	9.647151

Figure AD.1.26 presents the (logarithm of) total gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe a slight hump after the amendment.

Figure AD1.26

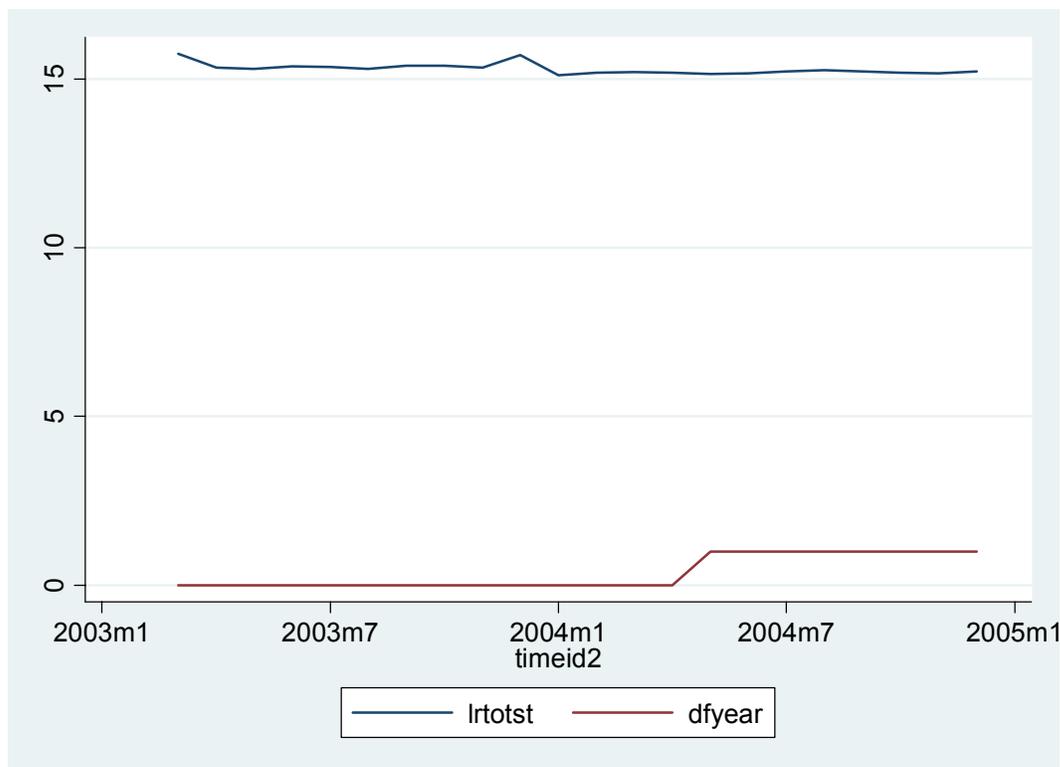


Table AD.1.31 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. The own lagged value is not significant as well. We find some evidence of a negative effect, about 12% decline, with appropriate statistical significance.

Table AD.1.31

Linear regression						Number of obs = 21	
						F(2, 18) = 3.94	
						Prob > F = 0.0380	
						R-squared = 0.2213	
						Root MSE = .12437	

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	lrtotst						
	L1.	-.0038092	.2324067	-0.02	0.987	-.4920776	.4844592
	dfyear	-.1270651	.0596788	-2.13	0.047	-.2524456	-.0016846
	_cons	15.39165	3.572212	4.31	0.000	7.886715	22.89659

Gloucester

Figure AD.1.27 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe a decline in average sales tax receipts since January of 2004, but with a slight increase afterwards.

Figure AD.1.27

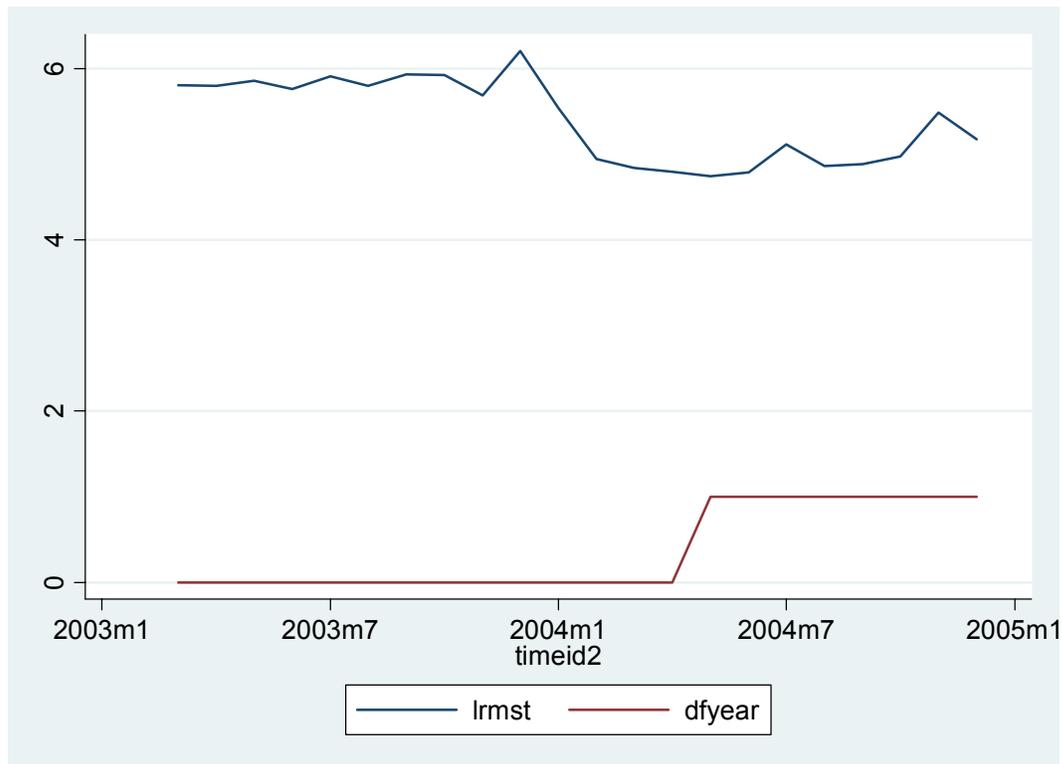


Table AD.1.31 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a negative effect, but not statistically significant.

Table AD.1.31

Linear regression		Number of obs = 21				
		F(2, 18) = 22.88				
		Prob > F = 0.0000				
		R-squared = 0.6766				
		Root MSE = .29385				
lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrmst						
L1.	.791294	.1729192	4.58	0.000	.4280042	1.154584
dfyear	-.0282787	.1807343	-0.16	0.877	-.4079874	.35143
_cons	1.110587	.9780127	1.14	0.271	-.9441418	3.165315

Figure AD.1.28 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe a sharp decline in total sales tax receipts since January of 2004, but with a slight increase afterwards.

Figure AD.1.28

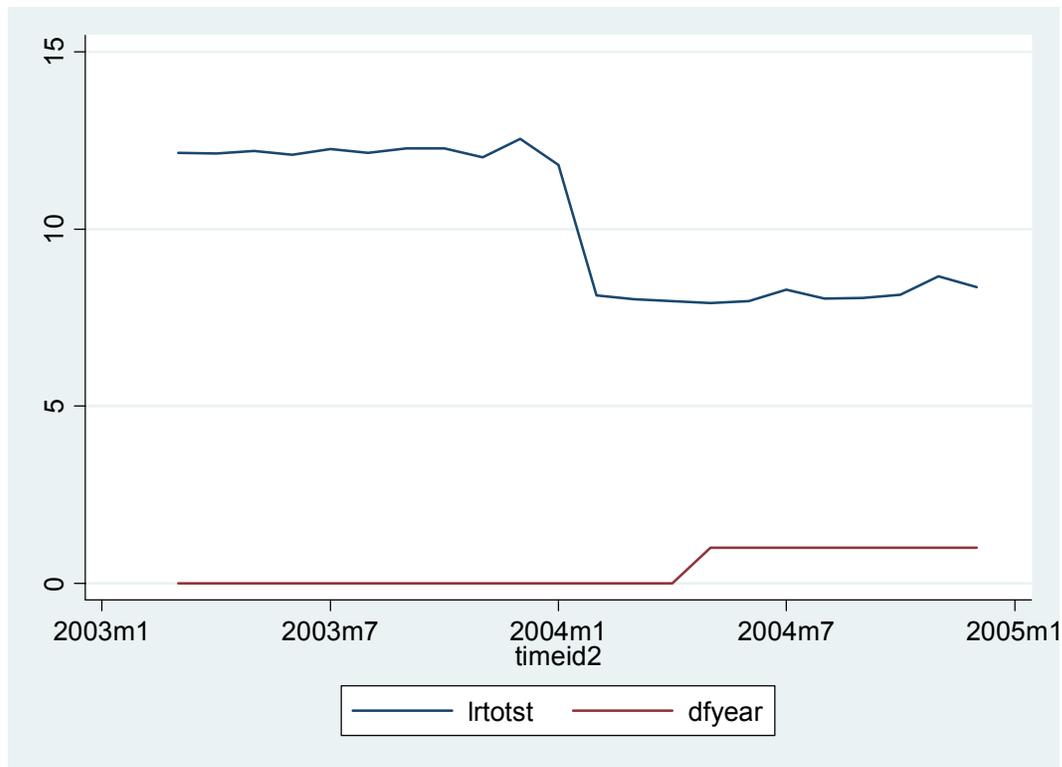


Table AD.1.32 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a positive effect, but not statistically significant.

Table AD.1.32

Linear regression		Number of obs =		21		
		F(2, 18) =		237.67		
		Prob > F =		0.0000		
		R-squared =		0.8417		
		Root MSE =		.86983		

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

	lrtotst					
	L1.	.9549192	.0483907	19.73	0.000	.8532541 1.056584
	dfyear	.2155243	.2135203	1.01	0.326	-.2330653 .6641138
	_cons	.199071	.3458614	0.58	0.572	-.5275569 .9256989

New Bedford

Figure AD.1.29 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe hump shape in average sales tax receipts since January of 2004, but with a slight increase afterwards.

Figure AD.1.29

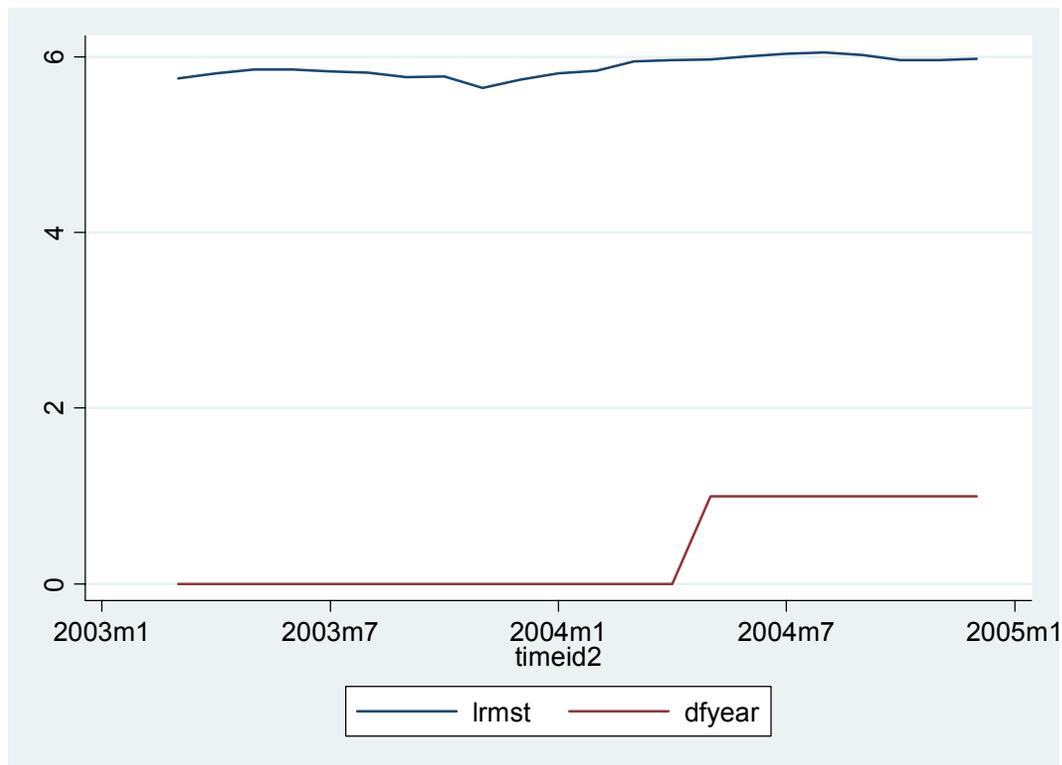


Table AD.1.33 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a positive effect, but not statistically significant.

Table AD.1.33

Linear regression		Number of obs = 21					
		F(2, 18) = 42.19					
		Prob > F = 0.0000					
		R-squared = 0.7931					
		Root MSE = .05332					
lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]		
lrmst							
L1.	.7451552	.1582435	4.71	0.000	.4126981	1.077612	
dfyear	.034077	.0325498	1.05	0.309	-.0343077	.1024616	
_cons	1.495806	.9222394	1.62	0.122	-.4417475	3.433359	

Figure AD.1.30 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe a slight downward trend in total sales tax receipts, but since about January of 2004 a slight increase.

Figure AD.1.30

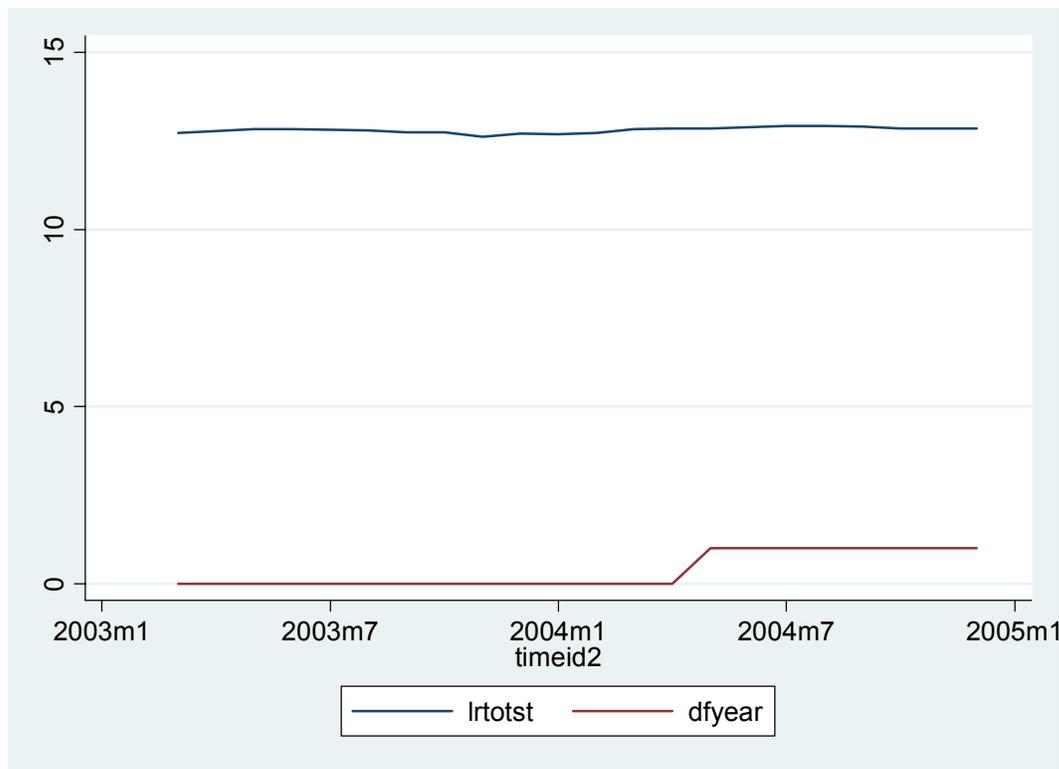


Table AD.1.34 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find evidence of a positive effect of about 4.2%, which is statistically significant.

Table AD.1.34

Linear regression					Number of obs =	21
					F(2, 18) =	21.25
					Prob > F	= 0.0000
					R-squared	= 0.6548
					Root MSE	= .04865

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

	lrtotst					
	l1.	.5805591	.1286029	4.51	0.000	.3103745 .8507437
	dfyear	.0423922	.0197089	2.15	0.045	.0009855 .083799
	_cons	5.361389	1.645971	3.26	0.004	1.903331 8.819446

Fairhaven

Figure AD.1.31 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe a downward trend in average sales tax receipts, but an increase since January of 2004.

Figure AD.1.31

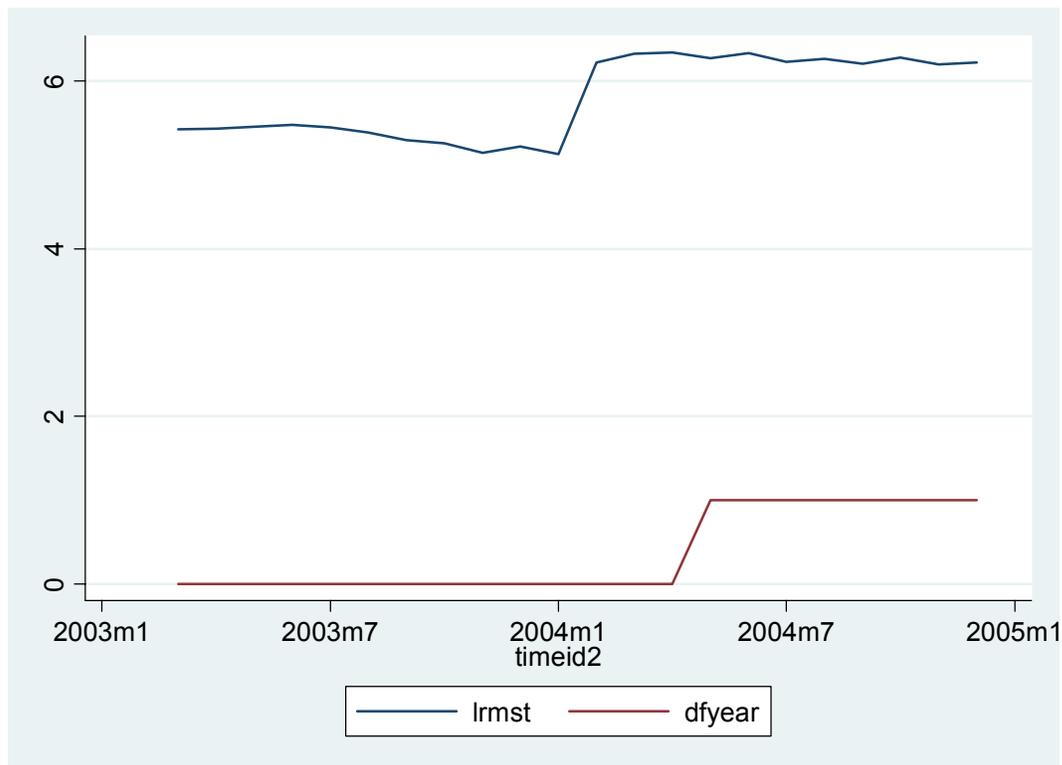


Table AD.1.35 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a positive effect, but not statistically significant.

Table AD.1.35

Linear regression		Number of obs = 21				
		F(2, 18) = 39.55				
		Prob > F = 0.0000				
		R-squared = 0.7560				
		Root MSE = .25534				
lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrmst						
L1.	.8379511	.23585	3.55	0.002	.3424486	1.333454
dfyear	.0422137	.1237883	0.34	0.737	-.2178559	.3022832
_cons	.9581319	1.367101	0.70	0.492	-1.914041	3.830305

Figure AD.1.32 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe a downward trend in total sales tax receipts, but an increase since January of 2004.

Figure AD.1.32

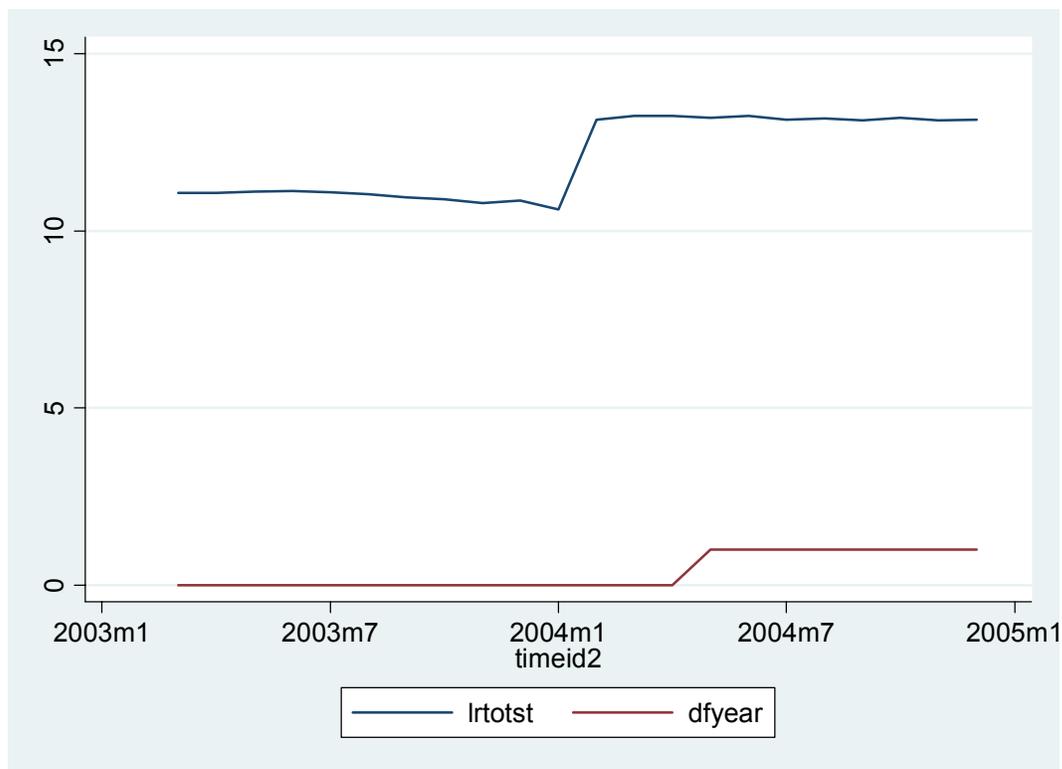


Table AD.1.36 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find a positive effect, but not statistically significant.

Table AD.1.36

Linear regression		Number of obs =		21		
		F(2, 18) =		60.82		
		Prob > F =		0.0000		
		R-squared =		0.7729		
		Root MSE =		.57554		

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]

	l1.	.8375206	.210118	3.99	0.001	.3960792 1.278962
	dfyear	.1215924	.2339761	0.52	0.610	-.3699731 .6131579
	_cons	2.005113	2.557376	0.78	0.443	-3.367736 7.377961

Plymouth

Figure AD.1.33 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe a downward trend in average sales tax receipts, but an increase since about March of 2004.

Figure AD.1.33

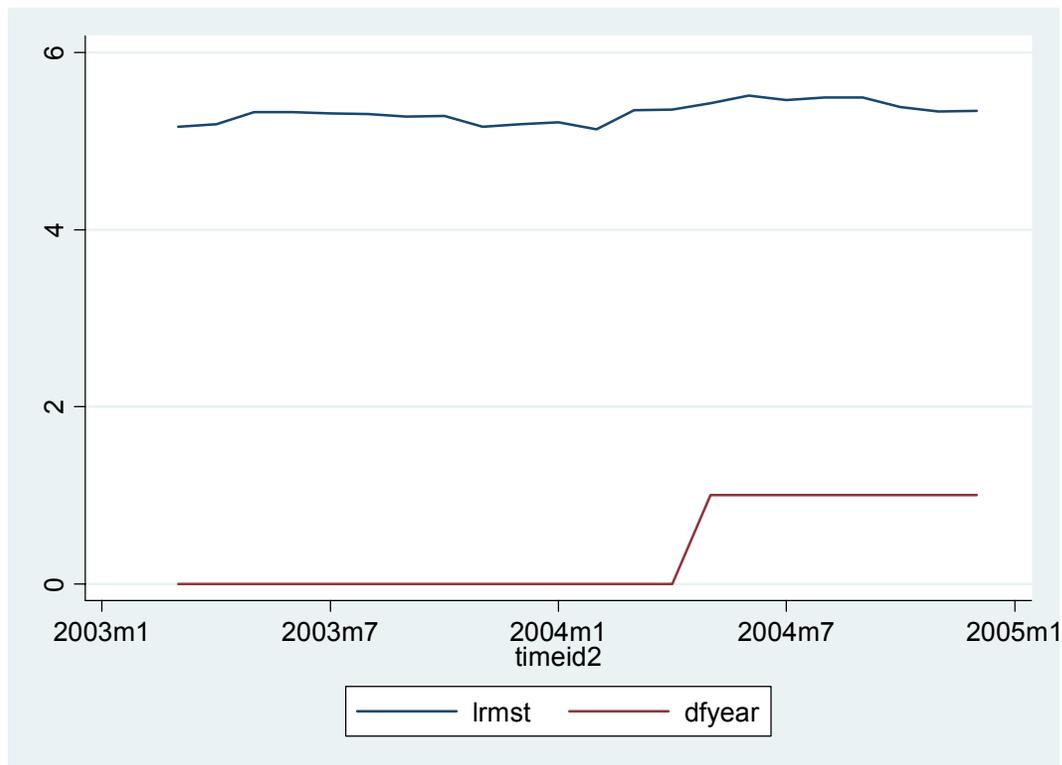


Table AD.1.37 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find evidence of a positive effect of about 8.7%, which is statistically significant.

Table AD.1.37

Linear regression		Number of obs =		21		
		F(2, 18) =		16.27		
		Prob > F =		0.0001		
		R-squared =		0.6497		
		Root MSE =		.0685		
lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrmst						
L1.	.4363274	.2207435	1.98	0.064	-.0274376	.9000924
dfyear	.0866301	.0414839	2.09	0.051	-.0005244	.1737846
_cons	2.973518	1.16915	2.54	0.020	.5172238	5.429811

Figure AD.1.34 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe a downward trend in total sales tax receipts, but an increase since about March of 2004.

Figure AD.1.34

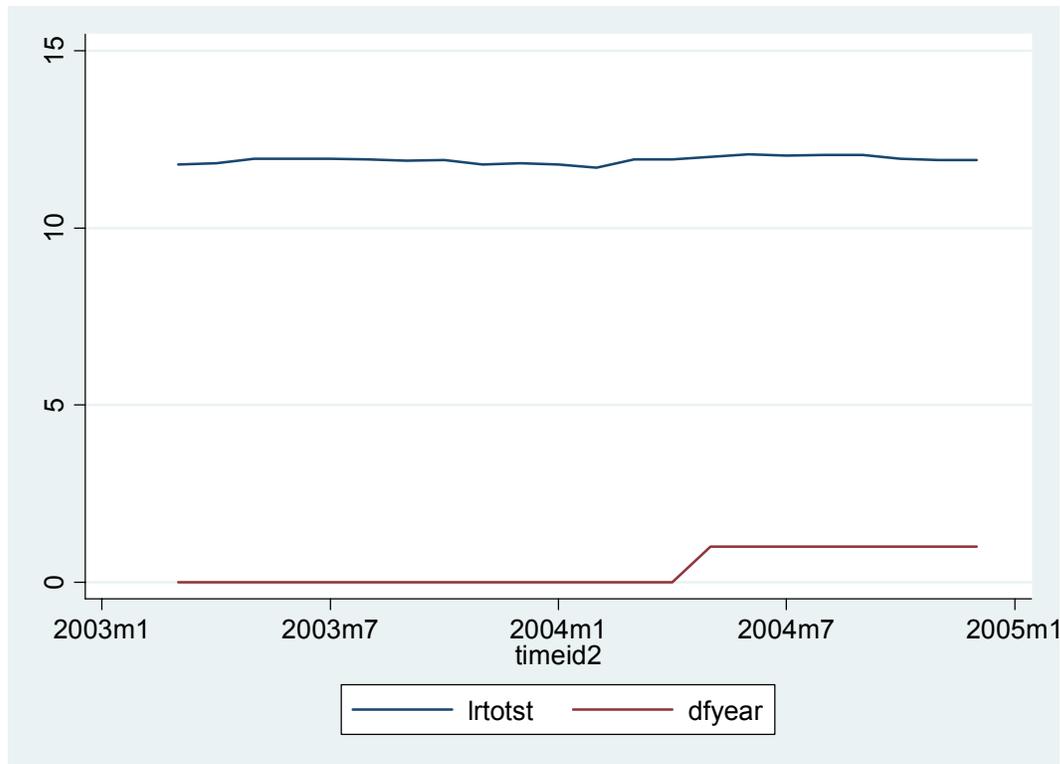


Table AD.1.38 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find evidence of a positive effect of about 6.2%, which is statistically significant at the 4% significance level.

Table AD.1.38

Linear regression						Number of obs =	21
						F(2, 18) =	9.65
						Prob > F =	0.0014
						R-squared =	0.5543
						Root MSE =	.06943

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	lrtotst						
	L1.	.4740119	.2394405	1.98	0.063	-.0290339	.9770577
	dfyear	.0619596	.0348649	1.78	0.092	-.0112889	.135208
	_cons	6.252633	2.853785	2.19	0.042	.257053	12.24821

Marshfield

Figure AD.1.35 presents the (logarithm of) average real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe a flat average sales tax receipts, but an increase since about January of 2004.

Figure AD.1.35

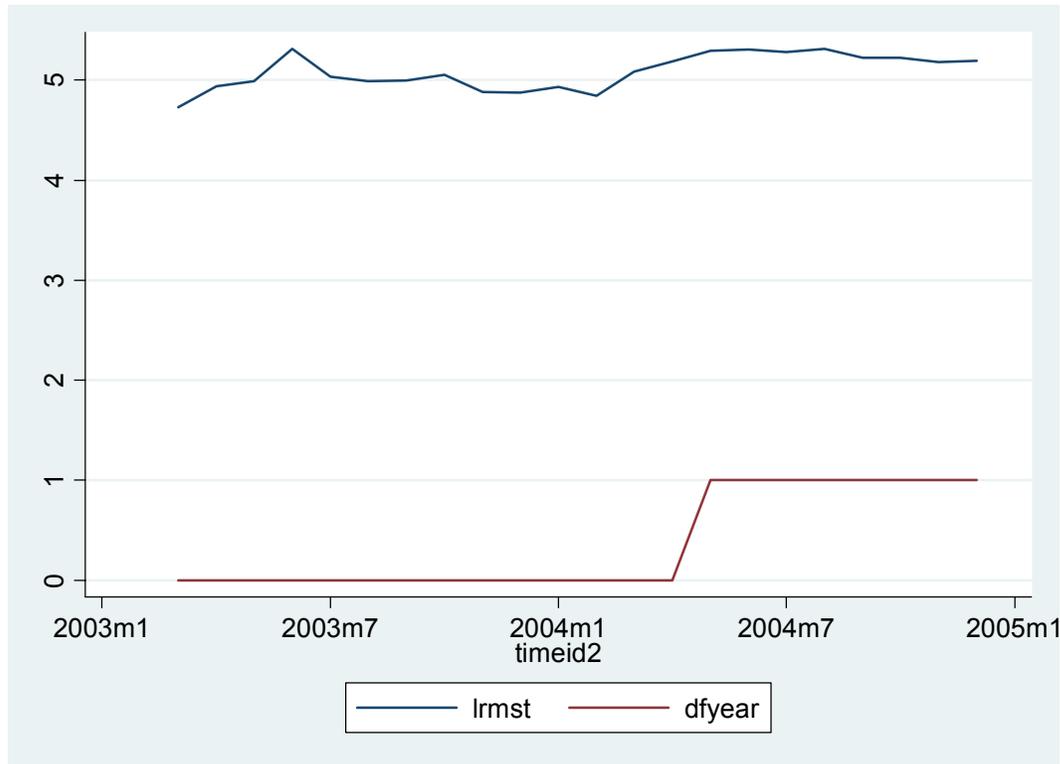


Table AD.1.39 presents the time series linear regression of the average real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find evidence of a positive effect of about 17%, which is statistically significant.

Table AD.1.39

Linear regression		Number of obs =		21		
		F(2, 18) =		22.49		
		Prob > F =		0.0000		
		R-squared =		0.6003		
		Root MSE =		.10774		
lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrmst						
L1.	.2632378	.1506116	1.75	0.098	-.0531854	.579661
dfyear	.1709887	.0633141	2.70	0.015	.0379708	.3040067
_cons	3.699594	.7430923	4.98	0.000	2.138415	5.260773

Figure AD.1.36 presents the (logarithm of) total real gross sales tax receipts (line above), the line below represents the introduction of the amendment in May 2004. We observe an increase since about January of 2004.

Figure AD.1.36

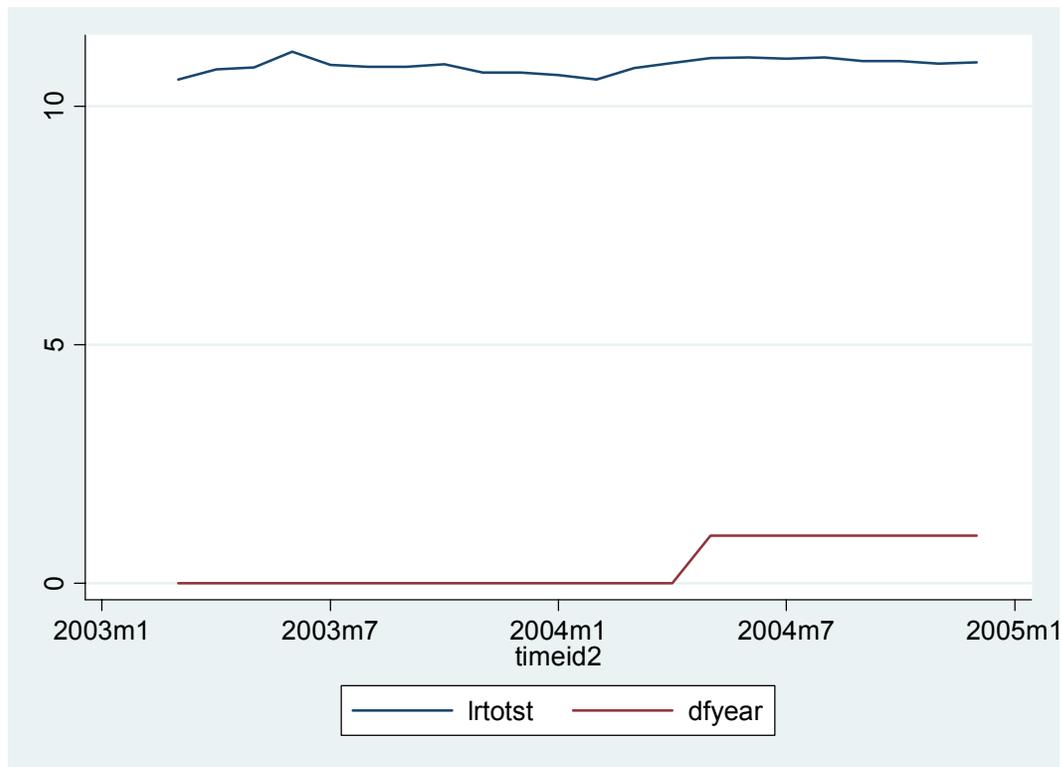


Table AD.1.40 presents the time series linear regression of the total real gross sales tax receipts taking into account the own lagged value and the intervention dummy variable of the amendment. We find evidence of a positive effect of about 10%, which is statistically significant at the 4% significance level.

Table AD.1.40

Linear regression		Number of obs =		21		
		F(2, 18) =		10.20		
		Prob > F =		0.0011		
		R-squared =		0.4362		
		Root MSE =		.11168		

		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
lrtotst						
L1.		.3399782	.1615556	2.10	0.050	.0005625 .6793938
dfyear		.1016675	.0517821	1.96	0.065	-.0071227 .2104577
_cons		7.143547	1.743231	4.10	0.001	3.481154 10.80594

In summary, the time series intervention evidence for Amendment 13 is as follows:

Table AD.1.40a

	Policy Intervention Variable	Statistically Significant
Boston	Negative	Yes
Gloucester	Inconclusive	No
New Bedford	Positive	Yes*
Fairhaven	Positive	No
Plymouth	Positive	Yes
Marshfield	Positive	Yes

* For total gross tax receipts only

AD.I.3.1. Amendment 13: May 2003 to December 2004 Period - Autocorrelation

We proceed now with estimates of the regressions in Baden and Bianconi (2006), but including lagged dependent variables, and with other assumptions as well.

In reference to Table 2.26 from Baden and Bianconi (2006), Table AD.1.41 shows linear regressions with lagged dependent variables. Results show no statistically significant effect of Amendment 13, in accordance with the original results of Table 2.26.⁷

⁷ We have performed several other econometric models and tests of fixed and random effects, available upon request. Results are unchanged.

Table AD.1.41

Linear regression							Number of obs = 7179
							F(4, 7174) = 1991.64
							Prob > F = 0.0000
							R-squared = 0.9277
							Root MSE = .30278

	lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	lrmst						
	L1.	.9567674	.0116578	82.07	0.000	.9339148	.9796201
	dfyear	-.007148	.00641	-1.12	0.265	-.0197135	.0054176
	dtreat	-.0058325	.0133616	-0.44	0.662	-.0320252	.0203601
	dyeartreat	.0154577	.0198179	0.78	0.435	-.0233911	.0543066
	_cons	.2439376	.0647044	3.77	0.000	.1170978	.3707773

Linear regression							Number of obs = 7179
							F(4, 7174) = 9662.35
							Prob > F = 0.0000
							R-squared = 0.9663
							Root MSE = .37304

	lrtotst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

In reference to Tables 2.31-2.35 in Baden and Bianconi (2006), we present in Table AD.1.42 models with separate difference-in-difference estimators for each town of the fishing industry group, reflecting changes in each town relative to all other towns in the state, excluding other fishing towns, and including lagged dependent variables. In Table AD.1.42, we note that the results previously obtained change substantively. We note that the following results are not robust to the change in econometric specification for the effects of Amendment 13:

- i. For Boston, Gloucester, Edgartown and Harwich, no statistically significant decline in average gross sales tax receipts is observed when lagged dependent variables are included;
- ii. For Fairhaven and Eastham, no statistically significant increase is observed in average gross sales tax receipts when lagged dependent variables are included;

iii. For Sandwich, a statistically significant increase is observed for average gross sales tax receipts when lagged dependent variables are included.⁸

Hence, we cannot corroborate the previous results of a negative impact of Amendment 13 on Gloucester, Boston, Edgartown and Harwich, with the alternative econometric specifications presented here.

Table AD.1.42

Linear regression						Number of obs = 6832	
						F(82, 6749) = 1111.38	
						Prob > F = 0.0000	
						R-squared = 0.9499	
						Root MSE = .25129	

lrmst	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]		

lrmst							
L1.	.8875913	.0234331	37.88	0.000	.8416551	.9335274	
L2.	.0907586	.0228413	3.97	0.000	.0459825	.1355348	
dfyear	-.0004561	.0055365	-0.08	0.934	-.0113093	.0103971	
dtreat	-.0823427	.0500473	-1.65	0.100	-.1804512	.0157657	
dyeartreat	-.0177289	.0273777	-0.65	0.517	-.0713978	.03594	
dfeb	-.0351471	.0444798	-0.79	0.429	-.1223415	.0520473	
dmar	.1678365	.0271835	6.17	0.000	.1145481	.2211248	
dapr	.1489909	.0280719	5.31	0.000	.0939611	.2040206	
dmay	.1526486	.0263295	5.80	0.000	.1010346	.2042627	
djun	.0918558	.0265764	3.46	0.001	.0397578	.1439539	
djul	.0557065	.0263219	2.12	0.034	.0041072	.1073058	
daug	.0506193	.0260753	1.94	0.052	-.0004966	.1017352	
dsep	.0519211	.0262079	1.98	0.048	.0005453	.1032969	
doct	.0495282	.0264705	1.87	0.061	-.0023624	.1014187	
dnov	-.0120251	.0262434	-0.46	0.647	-.0634705	.0394202	
ddec	.1030404	.0267128	3.86	0.000	.0506749	.155406	
dboston	-.0020632	.0550633	-0.04	0.970	-.1100046	.1058782	
dchatam	-.0283026	.0603005	-0.47	0.639	-.1465106	.0899053	
dfairhaven	.0635	.091148	0.70	0.486	-.1151788	.2421788	
dgloucester	-.1057972	.0866964	-1.22	0.222	-.2757494	.0641551	
dmarshfield	-.006033	.0542558	-0.11	0.911	-.1123915	.1003255	
dnewbedford	.0024659	.0377438	0.07	0.948	-.0715238	.0764556	
dplymouth	-.0090719	.0379938	-0.24	0.811	-.0835518	.065408	
dprovincetown	-.0404712	.0621169	-0.65	0.515	-.16224	.0812976	
dscituate	-.0329913	.033873	-0.97	0.330	-.099393	.0334104	
dbeverly	.0081747	.0163866	0.50	0.618	-.0239482	.0402975	
deastham	.079516	.1012961	0.78	0.432	-.1190564	.2780884	
dedgartown	-.1094412	.0764138	-1.43	0.152	-.2592363	.0403538	
dharwich	-.0347712	.0546292	-0.64	0.524	-.1418618	.0723193	
dwestport	-.007271	.0360338	-0.20	0.840	-.0779087	.0633666	
dmarblehead	-.0245062	.0405506	-0.60	0.546	-.1039981	.0549857	
dnantucket	-.0089522	.0473664	-0.19	0.850	-.1018053	.083901	
dnewburyport	.036544	.0509431	0.72	0.473	-.0633205	.1364085	
drockport	-.0782475	.0581344	-1.35	0.178	-.1922094	.0357143	
dsandwich	-.0273355	.0391275	-0.70	0.485	-.1040377	.0493667	
dswamp	-.0222183	.0442271	-0.50	0.615	-.1089173	.0644807	

⁸ This result is consistent with previously obtained results. We also found the result for Sandwich to hold in alternative specifications not presented here, but available upon request.

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dwellfleet	-.0520713	.1159226	-0.45	0.653	-.2793161	.1751736
dyarmouth	-.0179285	.0528323	-0.34	0.734	-.1214965	.0856394
dsalem	-.0190258	.0347246	-0.55	0.584	-.0870969	.0490453
dmanchester	-.0199126	.0509752	-0.39	0.696	-.1198401	.0800149
dorleans	-.0088085	.0316271	-0.28	0.781	-.0708076	.0531907
doakb	-.0162866	.0711957	-0.23	0.819	-.1558527	.1232795
dtisbury	-.005837	.0408759	-0.14	0.886	-.0859668	.0742927
ddartmouth	-.0042674	.0443349	-0.10	0.923	-.0911779	.082643
dtboston	.0329791	.0653956	0.50	0.614	-.095217	.1611752
dtchatam	.0167871	.0788009	0.21	0.831	-.1376876	.1712618
dtfairhaven	-.0612421	.1023748	-0.60	0.550	-.261929	.1394448
dtgloucester	.1513531	.1363373	1.11	0.267	-.1159109	.4186172
dtmarshfield	.0047784	.0614298	0.08	0.938	-.1156433	.1252001
dtnewbedford	.0120416	.0502558	0.24	0.811	-.0864755	.1105588
dtplymouth	.0078547	.0508169	0.15	0.877	-.0917625	.107472
dtprovince~n	.035588	.104666	0.34	0.734	-.1695904	.2407664
dtscituate	.0351861	.044942	0.78	0.434	-.0529143	.1232865
dteastham	-.0816473	.1274997	-0.64	0.522	-.331587	.1682923
dtedgartown	.1880899	.1612913	1.17	0.244	-.128092	.5042718
dtharwich	.0548597	.0879561	0.62	0.533	-.117562	.2272815
dtwestport	.0055966	.0431608	0.13	0.897	-.0790122	.0902054
dtmarblehead	.0559152	.0618926	0.90	0.366	-.0654138	.1772441
dtnantucket	.0244406	.075732	0.32	0.747	-.124018	.1728992
dtnewburyp~t	-.039069	.1036391	-0.38	0.706	-.2422342	.1640963
dtrockport	.0752852	.0717251	1.05	0.294	-.0653185	.215889
dtsandwich	.2824404	.1652936	1.71	0.088	-.0415872	.6064681
dtswamp	.0309843	.0607125	0.51	0.610	-.0880313	.1499998
dwellfleet	.0568438	.184885	0.31	0.759	-.3055891	.4192767
dyarmouth	.0735282	.0734351	1.00	0.317	-.0704277	.2174842
dsalem	.0360986	.0510102	0.71	0.479	-.0638975	.1360947
dmanchester	-.0340625	.0849498	-0.40	0.688	-.2005908	.1324659
dorleans	.0073202	.0423428	0.17	0.863	-.075685	.0903254
dtoakb	.0083733	.1062242	0.08	0.937	-.1998597	.2166063
dtisbury	.0086161	.0553323	0.16	0.876	-.0998526	.1170849
dtdartmouth	-.0086213	.0642751	-0.13	0.893	-.1346208	.1173782
dtreatmay	.1488443	.052502	2.84	0.005	.045924	.2517647
dtreatjun	.1527355	.0533214	2.86	0.004	.0482088	.2572622
dtreatjul	.2193871	.0567139	3.87	0.000	.10821	.3305643
dtreataug	.1088867	.0497334	2.19	0.029	.0113936	.2063798
dtreatsep	-.0580703	.0601934	-0.96	0.335	-.1760683	.0599277
dtreatoct	.0173202	.0524626	0.33	0.741	-.085523	.1201634
dtreatnov	.0542545	.0526102	1.03	0.302	-.0488781	.1573871
dtreatdec	.1258187	.0511339	2.46	0.014	.0255802	.2260572
dtreatfeb	.2284885	.0823521	2.77	0.006	.0670525	.3899245
dtreatmar	.0587426	.0529576	1.11	0.267	-.045071	.1625561
dtreatapr	.1173108	.053613	2.19	0.029	.0122123	.2224093
_cons	.0534427	.0318951	1.68	0.094	-.0090819	.1159672

Linear regression				Number of obs = 6832		
				F(82, 6749) = 3082.23		
				Prob > F = 0.0000		
				R-squared = 0.9740		
				Root MSE = .32798		

		Robust				
lrtotst	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

lrtotst						
L1.	.9451703	.0183559	51.49	0.000	.9091869	.9811536
L2.	.0425984	.0181066	2.35	0.019	.0071038	.078093

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dfyear	-.0019243	.0055559	-0.35	0.729	-.0128157	.0089671
dtreat	-.0519914	.0541941	-0.96	0.337	-.158229	.0542462
dyeartreat	-.0249804	.0331985	-0.75	0.452	-.0900599	.0400991
dfeb	.0327122	.0696849	0.47	0.639	-.1038923	.1693167
dmar	.2389692	.0275998	8.66	0.000	.1848648	.2930735
dapr	.209523	.0286237	7.32	0.000	.1534115	.2656346
dmay	.2156755	.0268551	8.03	0.000	.163031	.2683199
djun	.1539486	.0271283	5.67	0.000	.1007686	.2071286
djul	.1205343	.0268491	4.49	0.000	.0679017	.173167
daug	.1171054	.0266271	4.40	0.000	.0649079	.1693029
dsep	.1186141	.0267477	4.43	0.000	.0661802	.1710479
doct	.1162353	.0270106	4.30	0.000	.0632861	.1691845
dnov	.0549138	.0267905	2.05	0.040	.0023961	.1074315
ddec	.1735094	.0272163	6.38	0.000	.1201568	.226862
dboston	.0120816	.0599124	0.20	0.840	-.1053657	.1295289
dchatam	-.0278518	.0587004	-0.47	0.635	-.1429231	.0872195
dfairhaven	.1673062	.214716	0.78	0.436	-.253605	.5882173
dgloucester	-.3819707	.2947879	-1.30	0.195	-.959848	.1959065
dmarshfield	-.0159061	.0549059	-0.29	0.772	-.1235389	.0917268
dnewbedford	.0009881	.0399624	0.02	0.980	-.0773508	.0793269
dplymouth	-.0058865	.0396145	-0.15	0.882	-.0835433	.0717704
dprovincetown	-.0400635	.0621672	-0.64	0.519	-.1619308	.0818038
dscituate	-.0384964	.0357698	-1.08	0.282	-.1086166	.0316238
dbeverly	.0179738	.0182849	0.98	0.326	-.0178703	.053818
deastham	.1251446	.1526453	0.82	0.412	-.1740884	.4243776
dedgartown	-.2582679	.1367572	-1.89	0.059	-.5263552	.0098194
dharwich	-.1633683	.0811562	-2.01	0.044	-.3224601	-.0042765
dwestport	-.0083064	.0401939	-0.21	0.836	-.0870992	.0704863
dmarblehead	-.0291007	.0412262	-0.71	0.480	-.109917	.0517156
dnantucket	-.0112155	.0472172	-0.24	0.812	-.1037762	.0813452
dnewburyport	.0267109	.0521356	0.51	0.608	-.0754912	.1289131
drockport	-.0702711	.0545774	-1.29	0.198	-.17726	.0367179
dsandwich	-.0222645	.0411093	-0.54	0.588	-.1028518	.0583227
dswamp	-.0300909	.0473669	-0.64	0.525	-.122945	.0627632
dwellfleet	-.0607702	.114654	-0.53	0.596	-.2855283	.1639878
dyarmouth	-.018342	.0568305	-0.32	0.747	-.1297476	.0930637
dsalem	-.0223727	.035811	-0.62	0.532	-.0925735	.047828
dmanchester	-.0322864	.0558629	-0.58	0.563	-.1417952	.0772224
dorleans	-.0237725	.038613	-0.62	0.538	-.0994662	.0519211
doakb	-.0229701	.0704619	-0.33	0.744	-.1610977	.1151575
dtisbury	-.0154875	.0427597	-0.36	0.717	-.09931	.068335
ddartmouth	-.0057357	.0462147	-0.12	0.901	-.096331	.0848596
dtboston	.0304997	.0688229	0.44	0.658	-.1044149	.1654143
dtchatam	.0074234	.0764719	0.10	0.923	-.1424857	.1573325
dtfairhaven	-.1706495	.2196443	-0.78	0.437	-.6012216	.2599226
dtgloucester	.3839028	.3132618	1.23	0.220	-.2301892	.9979947
dtmarshfield	.0026913	.0616955	0.04	0.965	-.1182513	.123634
dtnewbedford	.0092273	.0517265	0.18	0.858	-.0921729	.1106276
dtplymouth	.0021085	.052181	0.04	0.968	-.1001827	.1043997
dtprovincetown	.026515	.10293	0.26	0.797	-.1752604	.2282903
dtscituate	.0309962	.0462132	0.67	0.502	-.0595961	.1215886
dteastham	-.1412109	.1722728	-0.82	0.412	-.47892	.1964982
dtedgartown	.3084098	.1996105	1.55	0.122	-.0828897	.6997094
dtharwich	.1621274	.1066387	1.52	0.128	-.0469181	.3711728
dtwestport	-.002827	.0464955	-0.06	0.952	-.0939728	.0883189
dtmarblehead	.0521814	.0616162	0.85	0.397	-.0686058	.1729685
dtnantucket	.0163817	.0743303	0.22	0.826	-.129329	.1620925
dtnewburyport	-.0456197	.1037601	-0.44	0.660	-.2490222	.1577827
dtrockport	.0637771	.0688	0.93	0.354	-.0710925	.1986467
dtsandwich	.2564825	.1669249	1.54	0.124	-.070743	.583708
dtswamp	.0264643	.0616417	0.43	0.668	-.0943728	.1473014
dtwellfleet	.0503726	.180777	0.28	0.781	-.3040074	.4047526
dyarmouth	.0613689	.076938	0.80	0.425	-.0894539	.2121916

dtsalem	.0316396	.0512508	0.62	0.537	-.0688282	.1321074
dtmanchester	-.0413037	.0879496	-0.47	0.639	-.2137127	.1311053
dtorleans	.0056678	.0480624	0.12	0.906	-.0885496	.0998852
dtoakb	-.0056789	.1035729	-0.05	0.956	-.2087145	.1973567
dttisbury	.0007192	.0561384	0.01	0.990	-.1093299	.1107682
dtdartmouth	-.0187407	.0666943	-0.28	0.779	-.1494825	.1120011
dtreatmay	.1474071	.0528636	2.79	0.005	.0437778	.2510364
dtreatjun	.1496642	.0532379	2.81	0.005	.0453012	.2540272
dtreatjul	.2158255	.0567726	3.80	0.000	.1045334	.3271177
dtreataug	.1009294	.0497826	2.03	0.043	.0033397	.1985191
dtreatsep	-.0603359	.0615174	-0.98	0.327	-.1809295	.0602577
dtreatoct	.023999	.0523368	0.46	0.647	-.0785977	.1265957
dtreatnov	.0570073	.0524608	1.09	0.277	-.0458324	.159847
dtreatdec	.1275086	.0523856	2.43	0.015	.0248164	.2302008
dtreatfeb	.0889844	.1834353	0.49	0.628	-.2706067	.4485754
dtreatmar	.0437522	.0515939	0.85	0.396	-.0573881	.1448924
dtreatapr	.1178678	.0536151	2.20	0.028	.0127652	.2229704
_cons	-.0012767	.0389509	-0.03	0.974	-.0776328	.0750794

In reference to the regressions in Tables 2.35-2.37 in Baden and Bianconi (2006) where we restrict the data to the subsample of towns affected by the fishing industry only, we present results below including lagged dependent variables. Table AD.1.43 indicates that original results from Table 2.35 are unchanged.

Table AD.1.43

Linear regression		Number of obs =		560		
		F(14, 545) =		528.75		
		Prob > F =		0.0000		
		R-squared =		0.9231		
		Root MSE =		.19248		

		Robust				
lrmst1	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

lrmst						
L1.	1.120841	.0595882	18.81	0.000	1.00379	1.237892
L2.	-.1676363	.0626976	-2.67	0.008	-.290795	-.0444777
dfyear	.0150885	.0176966	0.85	0.394	-.0196734	.0498503
dfeb	.2630942	.0776244	3.39	0.001	.1106146	.4155738
dmar	.2401243	.0538378	4.46	0.000	.1343693	.3458793
dapr	.278494	.0503804	5.53	0.000	.1795304	.3774576
dmay	.3031628	.0487379	6.22	0.000	.2074258	.3988999
djun	.2432679	.0501621	4.85	0.000	.1447332	.3418026
djul	.2918845	.052975	5.51	0.000	.1878243	.3959446
daug	.1702854	.0484427	3.52	0.000	.0751281	.2654427
dsep	.0357809	.0629622	0.57	0.570	-.0878974	.1594593
doct	.1446539	.0503478	2.87	0.004	.0457543	.2435534
dnov	.0948735	.0515439	1.84	0.066	-.0063756	.1961226
ddec	.2870309	.0515534	5.57	0.000	.1857632	.3882986
_cons	.0500029	.0798756	0.63	0.532	-.1068989	.2069048

Linear regression		Number of obs =		560		
		F(14, 545) =		1530.16		
		Prob > F =		0.0000		
		R-squared =		0.9647		
		Root MSE =		.28037		

lrtotst1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrtotst						
L1.	1.152485	.0668278	17.25	0.000	1.021213	1.283756
L2.	-.1680243	.0703957	-2.39	0.017	-.3063044	-.0297442
dfyear	.0085523	.0175477	0.49	0.626	-.0259171	.0430217
dfeb	.1940132	.1788231	1.08	0.278	-.1572537	.5452801
dmar	.3254123	.0552459	5.89	0.000	.2168912	.4339333
dapr	.3374103	.0514909	6.55	0.000	.2362653	.4385553
dmay	.3651465	.0499651	7.31	0.000	.2669988	.4632942
djun	.3007874	.0511082	5.89	0.000	.2003942	.4011805
djul	.346921	.0538625	6.44	0.000	.2411175	.4527245
daug	.2217256	.049252	4.50	0.000	.1249787	.3184726
dsep	.0874734	.0634904	1.38	0.169	-.0372426	.2121893
doct	.2014407	.0531176	3.79	0.000	.0971003	.305781
dnov	.1538309	.0526576	2.92	0.004	.0503941	.2572677
ddec	.3491844	.0530047	6.59	0.000	.245066	.4533028
_cons	-.0765561	.0973605	-0.79	0.432	-.2678039	.1146917

Table AD.1.44 presents the results with lagged dependent variables, for the same subsample, or differences-in-differences estimates for the ports of Gloucester, New Bedford/Fairhaven and Plymouth/Marshfield in reference to Tables 2-36-2.37 in Baden and Bianconi (2006). We find that the negative impact for Gloucester and the positive impact New Bedford/Fairhaven are not robust to this alternative specification.

Table AD.1.44

Linear regression				Number of obs = 560		
				F(20, 539) = 392.09		
				Prob > F = 0.0000		
				R-squared = 0.9233		
				Root MSE = .1933		
lrmst1	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lrmst						
L1.	1.118025	.0594067	18.82	0.000	1.001328	1.234723
L2.	-.1649707	.0625755	-2.64	0.009	-.2878924	-.0420489
dfyear	.0143215	.0190784	0.75	0.453	-.0231556	.0517986
dglouce	-.0507532	.0830029	-0.61	0.541	-.2138021	.1122956
intgloufyear	.0563427	.1412282	0.40	0.690	-.2210825	.3337679
dnbed	.0358178	.0321228	1.12	0.265	-.0272834	.098919
intnbfyear	-.0146292	.0449813	-0.33	0.745	-.1029892	.0737309
dply	.0104679	.035108	0.30	0.766	-.0584973	.0794332
intplyfyear	-.0185107	.0506883	-0.37	0.715	-.1180815	.0810602
dfeb	.2623061	.0773392	3.39	0.001	.1103829	.4142294
dmar	.2399389	.0530984	4.52	0.000	.1356336	.3442441
dapr	.2783501	.0496631	5.60	0.000	.1807932	.3759071
dmay	.3031289	.0480658	6.31	0.000	.2087097	.397548
djun	.2433232	.0493883	4.93	0.000	.1463061	.3403402
djul	.2917857	.0523417	5.57	0.000	.1889669	.3946044
daug	.1702955	.047516	3.58	0.000	.0769562	.2636348
dsep	.035467	.0625226	0.57	0.571	-.0873508	.1582848
doct	.143904	.0495222	2.91	0.004	.0466239	.241184
dnov	.0943543	.0506818	1.86	0.063	-.0052038	.1939123

ddec		.2864129	.0511053	5.60	0.000	.186023	.3868028
_cons		.0512395	.0823285	0.62	0.534	-.1104846	.2129636

Linear regression						Number of obs =	560
						F(20, 539) =	1306.82
						Prob > F =	0.0000
						R-squared =	0.9655
						Root MSE =	.27864

		Robust					
lrtotst1		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

lrtotst							
L1.		1.126326	.0641063	17.57	0.000	1.000397	1.252255
L2.		-.1411599	.0657189	-2.15	0.032	-.2702565	-.0120633
dfyear		.0006379	.0214155	0.03	0.976	-.0414301	.042706
dglouce		-.2929291	.2889098	-1.01	0.311	-.8604562	.274598
intgloufyear		.27163	.3055986	0.89	0.374	-.3286803	.8719402
dnbed		.0285264	.0331028	0.86	0.389	-.0364999	.0935527
intnbfyear		-.0145938	.0460106	-0.32	0.751	-.1049759	.0757883
dply		.0188489	.0335074	0.56	0.574	-.0469722	.0846701
intplyfyear		-.0206073	.0499533	-0.41	0.680	-.1187343	.0775198
dfeb		.1848542	.1788174	1.03	0.302	-.1664102	.5361185
dmar		.3200572	.0517025	6.19	0.000	.2184941	.4216202
dapr		.3362032	.048279	6.96	0.000	.241365	.4310413
dmay		.3648949	.0469097	7.78	0.000	.2727467	.457043
djun		.3011189	.0478115	6.30	0.000	.2071992	.3950385
djul		.3455223	.0515381	6.70	0.000	.2442822	.4467623
daug		.2211706	.0458009	4.83	0.000	.1312006	.3111407
dsep		.0836672	.0618485	1.35	0.177	-.0378264	.2051608
doct		.1935973	.0503431	3.85	0.000	.0947045	.2924901
dnov		.1484688	.0496915	2.99	0.003	.050856	.2460817
ddec		.3430533	.0513065	6.69	0.000	.2422681	.4438384
_cons		-.0723809	.0990117	-0.73	0.465	-.266877	.1221153

Finally, in Table AD.1.45 we present an alternative econometric estimation for the group of coastal towns with fishing activities. We estimate a fixed-effects econometric model for the group of coastal towns with fishing activity, an alternative model from Table AD.1.44 in reference to Tables 2-36-2.37 in Baden and Bianconi (2006). In this case, we find that the negative impact for Gloucester is statistically significant with magnitudes of -17% for average gross sales tax receipts, with upper/lower bounds of -35%/0.0% and -26% for total gross sales tax receipts, with upper and lower bounds of -57%/0.05%.

Given that average and total gross sales tax receipts in Gloucester have an approximate order of magnitude of \$49,539 and \$20,000,000 respectively on average per month in 2005 US dollars, a %17 loss in average gross sales taxes represents an average loss of \$8,422 per unit, with an upper bound of \$17,339 and a lower bound of no loss, or zero. For total gross sales tax receipts, the loss of %26 is of the order of \$5,200,000 with an upper bound of \$11,400,000 and a lower bound of a gain of \$1,000,000 reflecting the wider confidence interval in the estimation.

Table AD.1.45

ADDENDUM: Marine Fisheries Industry Economic Impact Study - Bid No. 06-41
 Baden and Bianconi, August 27, 2006

Fixed-effects (within) regression		Number of obs	=	560		
Group variable (i): townidl		Number of groups	=	28		
R-sq: within	= 0.7301	Obs per group: min	=	20		
between	= 0.9991	avg	=	20.0		
overall	= 0.9149	max	=	20		
		F(17,515)	=	81.95		
corr(u_i, Xb)	= 0.8237	Prob > F	=	0.0000		

lrmst1	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

lrmst						
L1.	1.013167	.042195	24.01	0.000	.9302713	1.096062
L2.	-.2826074	.0434574	-6.50	0.000	-.3679829	-.1972319
dfyear	.0444369	.0187681	2.37	0.018	.0075654	.0813083
intgloufyear	-.1705568	.091496	-1.86	0.063	-.3503082	.0091946
intnbfyear	-.0167058	.0860948	-0.19	0.846	-.185846	.1524345
intplyfyear	-.0229433	.0860974	-0.27	0.790	-.1920886	.1462021
dfeb	.2587547	.0507478	5.10	0.000	.1590565	.3584528
dmar	.2225425	.0495462	4.49	0.000	.1252049	.31988
dapr	.2745721	.0494748	5.55	0.000	.1773748	.3717694
dmay	.2952605	.0437043	6.76	0.000	.2093998	.3811212
djun	.2641684	.0437814	6.03	0.000	.1781563	.3501805
djul	.3378062	.0441654	7.65	0.000	.2510397	.4245727
daug	.2377721	.044611	5.33	0.000	.1501302	.325414
dsep	.1157406	.0453319	2.55	0.011	.0266824	.2047988
doct	.2065156	.0460148	4.49	0.000	.1161158	.2969154
dnov	.1306181	.0446909	2.92	0.004	.0428192	.218417
ddec	.3036934	.044721	6.79	0.000	.2158353	.3915515
_cons	1.165055	.1524511	7.64	0.000	.8655523	1.464557

sigma_u	.16491157					
sigma_e	.1849485					
rho	.44291595	(fraction of variance due to u_i)				

F test that all u_i=0:	F(27, 515) =	2.78			Prob > F =	0.0000

Fixed-effects (within) regression		Number of obs	=	560		
Group variable (i): townidl		Number of groups	=	28		
R-sq: within	= 0.8174	Obs per group: min	=	20		
between	= 0.9996	avg	=	20.0		
overall	= 0.9637	max	=	20		
		F(17,515)	=	135.58		
corr(u_i, Xb)	= 0.8891	Prob > F	=	0.0000		

lrtotst1	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

lrtotst						
L1.	1.063128	.0427124	24.89	0.000	.9792162	1.14704
L2.	-.2213974	.0452011	-4.90	0.000	-.3101987	-.1325962
dfyear	.0041678	.027083	0.15	0.878	-.0490389	.0573745
intgloufyear	-.26196	.1590624	-1.65	0.100	-.5744511	.050531
intnbfyear	-.015094	.127175	-0.12	0.906	-.2649395	.2347516
intplyfyear	-.0179733	.1271783	-0.14	0.888	-.2678255	.2318788
dfeb	.1799881	.0744532	2.42	0.016	.0337187	.3262575
dmar	.2912896	.0736667	3.95	0.000	.1465653	.4360138

dapr		.3045784	.0732452	4.16	0.000	.1606821	.4484746
dmay		.3598683	.064546	5.58	0.000	.2330624	.4866741
djun		.314491	.0645824	4.87	0.000	.1876136	.4413684
djul		.3753983	.0647894	5.79	0.000	.2481143	.5026824
daug		.2647273	.0649888	4.07	0.000	.1370514	.3924031
dsep		.136007	.0654754	2.08	0.038	.0073753	.2646387
doct		.2352311	.0661592	3.56	0.000	.105256	.3652061
dnov		.1727186	.0652477	2.65	0.008	.0445343	.300903
ddec		.3552028	.0653111	5.44	0.000	.2268938	.4835118
_cons		1.491582	.2781164	5.36	0.000	.9452	2.037964

sigma_u		.22337473					
sigma_e		.27320664					
rho		.40065051	(fraction of variance due to u_i)				

F test that all u_i=0:		F(27, 515) =	2.18			Prob > F =	0.0006

Hence, the main conclusion for the period of Amendment 13, May 2003-December 2004, is that we find some statistically significant evidence of a decline in sales tax receipts in the port city of Gloucester in the North Shore relative to all towns in the fishing industry.

AD.I.4. Summary of Results for Sales Tax Data

The time series properties and intervention analysis of policy interventions for sales tax receipts data were presented and complement the previous analysis of Baden and Bianconi (2006). There is evidence of important correlations of monthly sales tax receipts over time. Measuring those effects and including those effects in the regressions do strengthen some of our previous results, and also change our previous results in some important dimensions:

- Separate time series and intervention analysis for Boston, Gloucester, New Bedford, Fairhaven, Plymouth and Marshfield show the following results for the effects of statistical evidence of Amendments 5, 7 and 13:

Table AD.1.46

<i>Amendment 5</i>	Policy Intervention Variable	Statistically Significant
Boston	Negative	Yes
Gloucester	Negative	No
New Bedford	Negative	No
Fairhaven	Negative	No
Plymouth	Negative	No
Marshfield	Negative	No
<i>Amendment 7</i>	Policy Intervention Variable	Statistically Significant
Boston	Positive	No
Gloucester	Positive	No
New Bedford	Positive	Yes
Fairhaven	Negative	Yes
Plymouth	Positive	Yes
Marshfield	Inconclusive	No
<i>Amendment 13</i>	Policy Intervention Variable	Statistically Significant
Boston	Negative	Yes
Gloucester	Inconclusive	No
New Bedford	Positive	Yes*
Fairhaven	Positive	No
Plymouth	Positive	Yes
Marshfield	Positive	Yes

While the effects of Amendment 5 are all negative in magnitude, only Boston has a statistically significant effect.

- Amendment 5 had a negative effect for all towns in the fishing industry as shown in Table 2.11 of Baden and Bianconi (2006). Table AD.1.15 confirms those results, but magnitudes are smaller when autocorrelation of dependent variables is taken into account. The negative impact of the policy intervention is -4.4% and -2.8% lower average and total gross real sales tax receipts respectively. An approximate dollar value figure for those losses is a \$1,247 monthly loss in average sales tax revenues for each town in the fishing industry, the %4.4 point estimate measured in 1994-1995 dollars with an upper/lower bound of \$2,168/\$354 monthly loss in average sales tax revenues. For the totals, 2.8% represents an approximate \$1,235,000 monthly loss in total gross sales tax revenues for each town in the fishing industry, measured in 1994-1995 dollars with an upper/lower bound of \$2,691,000/\$176,468 per month quoted in 1994-1995 US dollars.
- Amendment 7 had no identified statistically significant effects.
- Amendment 13: Under the alternative specifications, relative to all other towns in the State, Boston, Gloucester, Edgartown and Harwich, no statistically significant decline in average gross sales tax receipts is observed when lagged

dependent variables are included; and in Fairhaven and Eastham, no statistically significant increase is observed in average gross sales tax receipts when lagged dependent variables are included. Table AD.1.45 shows that Amendment 13 had a negative effect in Gloucester in the North Shore relative to the other fishing towns in the State only, the magnitude of the effect is of the order of -17% for average gross sales tax receipts and -26% for total gross sales tax receipts, smaller than previously obtained. Given that average and total gross sales tax receipts in Gloucester have an approximate order of magnitude of $\$49,539$ and $\$20,000,000$ respectively on average per month in 2005 US dollars, a 17% loss in average gross sales taxes represents an average loss of $\$8,422$ per unit, with an upper/lower bound of $\$17,339/\0 . For total gross sales tax receipts, the loss is 26% and of the order of $\$5,200,000$ with an upper/lower bound of $\$11,400,000/+\$1,000,000$ reflecting the wider confidence interval in the estimation.

Further analysis of these data is important in two dimensions: (i) Relaxing the lagged dependent variables assumption that the effect is identical across all units in the State; (ii) Constructing a data set across all time spectrum, from 1992 to 2005, and examining intervention analysis in the whole time series sample for different towns in the State.

AD II. Addendum to Employment Effects Section

This addendum clarifies the results from Section 3.2 of Baden and Bianconi (2006). This addendum also re-interprets the coefficient significance levels presented in Baden and Bianconi (2006) as one-sided tests (whether the coefficients are negative or not). A confidence interval (range) for lost jobs and wages is provided for the entire New England fishery.

AD.II.1 Amendment 13 and Employment

Traditionally, significance levels for regression coefficients are interpreted according to a t-test test of a null hypothesis of no difference from zero against an alternative hypothesis of different than (greater than or less than) zero – this is the t-statistic generated automatically by the regression analysis software. While this is sufficient in many cases, for the specific test of whether jobs have been lost or not, it is more appropriate to use a 1-tailed t-test, which tests the null hypothesis of no difference in percentage employment to the alternative hypothesis of a decrease in percentage employment. For $n > 30$, this t-statistic is 1.64 for statistical significance at the 95% confidence level, and 1.28 for statistical significance at the 90% confidence level.

Under this test of non-negative coefficient values, several of the results that were marginally significant for 2-tailed tests (of inequality) become significant for the 1-tailed test (of non-negativity). This test is actually more appropriate than a 2-tailed test because the conceptual model behind this analysis posits that the implementation of Amendment 13 could have negative effects on fisheries employment. There is no postulation that Amendment 13 should increase fisheries employment.

For the difference in difference models a 1 tail test is also used. This test assesses whether the difference between employment percentage in manufacturing and employment percentage in fishing is greater than zero. A greater than zero value indicates that the difference between the two percentages increased, which means that employment percentage in fishing fell relative to manufacturing.

- *Results: All Towns and Counties*

Upon re-interpreting the t-statistic values for a one-tailed test of negative values, the regression results for percentage employment in fishing for all fishing areas (the same fixed effects panel model with lagged variables) results in a statistically significant decrease in fisheries employment across all towns at the 90% confidence level.

Regression 3.2.1a: All Places % Fishing

Regression with robust standard errors

Number of obs = 455
 F(14, 441) = 1207.28
 Prob > F = 0.0000
 R-squared = 0.9767
 Root MSE = .29415

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pctf_lag1	.4204863	.0833628	5.04	0.000	.2566486	.5843241
qtr2	.0670156	.0443749	1.51	0.132	-.0201971	.1542282
qtr3	.0801494	.040558	1.98	0.049	.0004384	.1598604
qtr4	.1690242	.0443273	3.81	0.000	.0819053	.2561432
t	.001725	.0017235	1.00	0.317	-.0016622	.0051122
a13	-.0817846	.0546813	-1.50	0.135	-.1892529	.0256838
boston	-.0963192	.0532932	-1.81	0.071	-.2010594	.0084211
BC	-.0538152	.051707	-1.04	0.299	-.155438	.0478075
FH	1.605183	.2527948	6.35	0.000	1.108351	2.102015
GL	1.07981	.1737805	6.21	0.000	.7382692	1.421351
NB	1.61928	.2428367	6.67	0.000	1.142019	2.096541
PC	-.1154694	.0652346	-1.77	0.077	-.2436789	.01274
RP	1.008272	.1703155	5.92	0.000	.6735414	1.343003
WP	.8788293	.1482549	5.93	0.000	.5874553	1.170203

This regression includes a lagged dependent variable to reduce the effects of serial auto-correlation, quarterly dummies (qtr2, etc), a linear time trend variable (t), and town/county fixed effects (Boston, BC, etc.). The effects of Amendment 13 (a13, indicated in gray), are -.08% which represents a decrease of 41% in fisheries employment across all towns and counties. For a one-tailed test, this result is significant at the 93% confidence level.

Table AD 2.1 provides the estimated wages lost per worker across all towns for this regression which takes the average effects. A 90% confidence interval for this result is between a wage loss of \$15,315,000 (175 jobs lost) and a wage gain of \$700,000 (8 jobs gained). To make this calculation, the weekly wage data from the Massachusetts Department of Workforce Development are averaged per year for all towns and adjusted by the consumer price index. All dollar values are in April 2006 constant dollars.

Table AD 2.1: Fishing Employment Changes in All Places

Town	Change in % Employment in Fishing	% Change in Fishing Employment Levels	Wage Change Lower Bound (# of Jobs)	Wage Change (# of Jobs)	Wage Change Upper Bound (# of Jobs)
All Places	- 0.08%	- 5.6 %	-\$15,315,000 (-175)	-\$7,351,000 (-84)	+\$700,000 (+8)

The difference in difference estimator also shows an effect associated with the imposition of Amendment 13, as shown in Regression 3.2.1b. For the 1-tailed test of a positive difference, the coefficient is significant at the 92% confidence level, which indicates that

overall, among Massachusetts fishing towns for which data were available, fishing employment shrunk relative to manufacturing employment.

Regression 3.2.1b: All Places %Manufacturing - % Fishing

Regression with robust standard errors

Number of obs = 454
 F(15, 439) = 6702.37
 Prob > F = 0.0000
 R-squared = 0.9935
 Root MSE = 1.1487

diff_	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
diff_1	.5746526	.1400694	4.10	0.000	.2993628	.8499425
diff_2	-.1427101	.0511507	-2.79	0.006	-.2432408	-.0421794
qtr2	-.6443001	.1732501	-3.72	0.000	-.9848028	-.3037975
qtr3	-.7616543	.2600682	-2.93	0.004	-1.272788	-.250521
qtr4	-.5726569	.2297334	-2.49	0.013	-1.024171	-.121143
t	-.0293859	.0143821	-2.04	0.042	-.0576522	-.0011197
a13	.408303	.2877125	1.42	0.157	-.1571621	.9737681
boston	2.957284	1.013263	2.92	0.004	.9658337	4.948734
BC	3.071578	1.039828	2.95	0.003	1.027918	5.115238
FH	8.748645	2.816808	3.11	0.002	3.212541	14.28475
GL	16.99719	5.55469	3.06	0.002	6.080098	27.91428
NB	13.50254	4.368039	3.09	0.002	4.917672	22.08741
PC	6.463784	2.162224	2.99	0.003	2.214187	10.71338
RP	2.953175	1.036853	2.85	0.005	.9153619	4.990988
WP	3.368106	1.137766	2.96	0.003	1.131961	5.60425

The a13 variable (highlighted in gray) indicates that the relative share of employment in fisheries relative to manufacturing increased during this period.

The results of considering all towns together show some indication of a decline in the fisheries industry associated with the imposition of Amendment 13.

The next section provides estimates for the percentage employment and difference in difference models for each city or county separately.⁹

- *Results: Boston and Cape Ann*

The results for % Fishing in Boston are presented below in Regression 3.2.2a:

⁹ Estimation of the separate city and county regressions as a Seemingly Unrelated Regression system of equations provided no significant efficiency gains.

Regression 3.2.2a: Boston % Fishing

Regression with robust standard errors

Number of obs = 56
 F(6, 49) = 25.67
 Prob > F = 0.0000
 R-squared = 0.6630
 Root MSE = .00197

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pctfB_1	.76078	.0906867	8.39	0.000	.5785384	.9430217
qtr2	-.0006047	.0007789	-0.78	0.441	-.0021699	.0009604
qtr3	-.000578	.0009751	-0.59	0.556	-.0025374	.0013815
qtr4	-.0002039	.0006871	-0.30	0.768	-.0015847	.0011769
t	3.43e-06	.0000208	0.16	0.870	-.0000384	.0000453
a13	-.0009358	.0008094	-1.16	0.253	-.0025624	.0006907
_cons	.0021562	.0012233	1.76	0.084	-.0003021	.0046145

Fishing employment fell in Boston, but not in a statistically significant sense.

The results for the difference in difference model for Boston are unchanged. For Boston, the percentage of fishing employment fell relative to the percentage of manufacturing employment.

Regression 3.2.2b: Boston %Manufacturing - %Fishing

Regression with robust standard errors

Number of obs = 56
 F(6, 49) = 1064.97
 Prob > F = 0.0000
 R-squared = 0.9930
 Root MSE = .02604

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
diffB_1	.7945303	.0686334	11.58	0.000	.6566062	.9324543
qtr2	-.0156774	.0103656	-1.51	0.137	-.0365078	.0051531
qtr3	.0022391	.010335	0.22	0.829	-.0185299	.023008
qtr4	-.0164696	.0099512	-1.66	0.104	-.0364672	.0035281
t	-.0044608	.0016349	-2.73	0.009	-.0077462	-.0011754
a13	.0449555	.0197362	2.28	0.027	.0052941	.0846169
_cons	.7286995	.2506177	2.91	0.005	.2250644	1.232335

The results for Gloucester (Regression 3.2.3a) are unchanged:

Regression 3.2.3a: Gloucester % Fishing

Regression with robust standard errors

Number of obs = 57
 F(5, 51) = 10.09
 Prob > F = 0.0000
 R-squared = 0.4870
 Root MSE = .20369

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pctfGL						
qtr2	-.2157478	.0954511	-2.26	0.028	-.4073738	-.0241218
qtr3	-.0346	.0597442	-0.58	0.565	-.1545414	.0853414
qtr4	.1704469	.0661806	2.58	0.013	.0375838	.30331
t	.0116001	.0030042	3.86	0.000	.005569	.0176312
a13	-.3336981	.1068932	-3.12	0.003	-.5482952	-.1191011
_cons	1.847392	.1008418	18.32	0.000	1.644944	2.049841

These results show that the percentage employment in fishing in Gloucester fell 0.33%, which is a 16% decrease in fishing employment levels (Average percentage fishing employment was 2.05%; 0.33 is 16% of 2.05). For the period studied, this amounts to a layoff of 32 employees for 39 weeks in 2004 and 39 weeks in 2005 during the study period. Average wages per week were \$879 in 2004 and \$953 in 2005.¹⁰ This amounts to \$ 71,448 lost wages per worker, and an aggregate loss of approximately \$2,286,000. Table AD2.2 summarizes these calculations for Gloucester and provides the 90% confidence interval for the calculated wage changes.

Table AD2.2: Fishing Employment Changes in Gloucester

Town	Change in % Employment in Fishing	% Change in Fishing Employment Levels	Wage Change Lower Bound (# of Jobs)	Wage Change (# of Jobs)	Wage Change Upper Bound (# of Jobs)
Gloucester	- 0.33%	- 16.0 %	-\$3,501,000 (-49)	-\$2,286,000 (-32)	-\$1,072,000 (-15)

The results for the difference in difference estimator for Gloucester are unchanged.

¹⁰ All dollar figures are expressed in April 2006 constant dollars. Average wage data are provided with employment data by the Massachusetts Department of Workforce Development.

Regression 3.2.3b: Gloucester %Manufacturing - %Fishing

Regression with robust standard errors

Number of obs =	56
F(6, 49) =	73.87
Prob > F =	0.0000
R-squared =	0.9034
Root MSE =	.72778

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
diff_GL						
diffGL_1	.821762	.0774532	10.61	0.000	.6661139	.9774101
qtr2	-1.593567	.3348431	-4.76	0.000	-2.26646	-.9206751
qtr3	-.9373232	.4051015	-2.31	0.025	-1.751405	-.1232412
qtr4	-.3170561	.3257915	-0.97	0.335	-.9717586	.3376463
t	-.0035634	.0125412	-0.28	0.778	-.028766	.0216391
a13	.3639424	.4125939	0.88	0.382	-.4651961	1.193081
_cons	5.532008	2.476649	2.23	0.030	.5549962	10.5090

Regression 3.2.4a indicates a decrease of 0.26% employment in fisheries in Rockport associated with Amendment 13. This decrease represents a 13% decrease in total fisheries employment. For a test of non-negativity of the coefficient, these results are significant at the 93% confidence level. Table AD2.3 displays the confidence interval and point estimate for lost wages in Rockport.

Regression 3.2.4a: Rockport % Fishing

Regression with robust standard errors

Number of obs =	57
F(5, 51) =	25.86
Prob > F =	0.0000
R-squared =	0.6461
Root MSE =	.36267

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pctFRP						
qtr2	-.809443	.1607451	-5.04	0.000	-1.132152	-.4867337
qtr3	-.6506105	.161118	-4.04	0.000	-.9740684	-.3271527
qtr4	.2753294	.1686449	1.63	0.109	-.0632393	.6138982
t	.0034188	.0045049	0.76	0.451	-.005625	.0124627
a13	-.2559688	.1678485	-1.52	0.133	-.5929386	.081001
_cons	2.197503	.1669278	13.16	0.000	1.862381	2.532624

Table AD2.3: Fishing Employment Changes in Gloucester

Town	Change in % Employment in Fishing	% Change in Fishing Employment Levels	Wage Change Lower Bound (# of Jobs)	Wage Change (# of Jobs)	Wage Change Upper Bound (# of Jobs)
Rockport	- 0.26%	- 13.8 %	-\$278,000 (6)	-\$139,000 (-3)	-\$46,000 (+1)

The results and interpretation for the difference in difference estimator for Rockport are unchanged:

Regression 3.2.4b: Rockport %Manufacturing - %Fishing

Regression with robust standard errors

Number of obs = 56
 F(6, 49) = 35.55
 Prob > F = 0.0000
 R-squared = 0.8182
 Root MSE = .45216

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
diff_RP						
diffRP_1	.7759133	.1009706	7.68	0.000	.5730052	.9788213
qtr2	-.6786383	.240514	-2.82	0.007	-1.161969	-.1953073
qtr3	-.5757476	.1992679	-2.89	0.006	-.9761915	-.1753037
qtr4	-.7367625	.1974918	-3.73	0.000	-1.133637	-.3398878
t	.0069145	.0058223	1.19	0.241	-.0047858	.0186148
a13	.1624593	.2367103	0.69	0.496	-.3132278	.6381464
_cons	.9420086	.3348403	2.81	0.007	.2691219	1.614895

- *Results: Cape Cod*

The regression results and interpretation for Barnstable County are unchanged.

Regression 3.2.5a: Barnstable County % Fishing

Regression with robust standard errors

Number of obs = 56
 F(6, 49) = 15.42
 Prob > F = 0.0000
 R-squared = 0.6840
 Root MSE = .00929

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pctfBC						
pctfBC_1	.3250908	.1065449	3.05	0.004	.1109808	.5392008
qtr2	.0166916	.0037413	4.46	0.000	.0091732	.0242099
qtr3	.0235195	.0035358	6.65	0.000	.0164141	.0306249
qtr4	.0211679	.0038871	5.45	0.000	.0133564	.0289794
t	-.0001816	.0001358	-1.34	0.187	-.0004545	.0000913
a13	.0036443	.0045151	0.81	0.423	-.0054291	.0127177
_cons	.0408571	.0099711	4.10	0.000	.0208194	.0608948

Regression 3.2.5b indicates that the difference between the share of fisheries and manufacturing employment has increased, but not at a statistically significant level.

Regression 3.2.5b: Barnstable County %Manufacturing - %Fishing

Regression with robust standard errors

Number of obs =	56
F(6, 49) =	419.64
Prob > F =	0.0000
R-squared =	0.9780
Root MSE =	.0816

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
diffBC_1	.837439	.0658617	12.72	0.000	.7050849	.9697932
qtr2	-.1888387	.0297494	-6.35	0.000	-.2486224	-.1290549
qtr3	-.0795056	.0495342	-1.61	0.115	-.1790482	.020037
qtr4	.0630207	.0352893	1.79	0.080	-.0078958	.1339373
t	-.0051215	.0024485	-2.09	0.042	-.010042	-.0002009
a13	.0477714	.045946	1.04	0.304	-.0445606	.1401034
_cons	.6945896	.2924715	2.37	0.022	.1068461	1.282333

Taken together, the results do not show a substantial influence of Amendment 13 upon fisheries employment for Barnstable County.

- *Results: Plymouth County*

Exclusive fishing data are unavailable for the municipalities of Plymouth and Marshfield, but such data are available for Plymouth County. The implementation of Amendment 13 coincides with both a substantial decrease in fishing employment and a decrease in the share of fishing relative to manufacturing.

The regression results from Regression 3.2.6a show the effect of Amendment 13 regulations on fishing employment in Plymouth County. The timing of Amendment 13 is highly significant.

Regression 3.2.6a: Plymouth County %Fishing

Regression with robust standard errors

Number of obs = 55
 F(7, 47) = 20.79
 Prob > F = 0.0000
 R-squared = 0.7991
 Root MSE = .00266

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pctfPC_1	.4778169	.1198128	3.99	0.000	.2367846	.7188492
pctfPC_2	.1932644	.1286625	1.50	0.140	-.0655712	.4520999
qtr2	.0043763	.0014261	3.07	0.004	.0015074	.0072453
qtr3	.0036768	.001213	3.03	0.004	.0012366	.0061169
qtr4	.0028255	.0013851	2.04	0.047	.000039	.005612
t	.0001539	.0000484	3.18	0.003	.0000566	.0002513
a13	-.0042596	.0015026	-2.83	0.007	-.0072823	-.0012368
_cons	-.0009775	.0019736	-0.50	0.623	-.0049479	.0029929

Table AD2.4 shows the lost employment and wage estimates as well as the 90% confidence interval for these estimates.

Table AD2.4: Fishing Employment Changes in Plymouth County

Town	Change in % Employment in Fishing	% Change in Fishing Employment Levels	Wage Change Lower Bound (# of Jobs)	Wage Change (# of Jobs)	Wage Change Upper Bound (# of Jobs)
Plymouth County	- 0.004%	- 28.6 %	-\$589,000 (-10)	-\$353,000 (-6)	-\$118,000 (2)

The difference in difference estimator is unchanged.

Regression 3.2.6b: Plymouth County %Manufacturing - %Fishing

Regression with robust standard errors

Number of obs = 56
 F(6, 49) = 353.00
 Prob > F = 0.0000
 R-squared = 0.9637
 Root MSE = .12408

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
diffPC_1	.805747	.0701139	11.49	0.000	.6648478	.9466461
qtr2	-.2245758	.0621541	-3.61	0.001	-.3494792	-.0996725
qtr3	-.1560215	.0907387	-1.72	0.092	-.3383678	.0263248
qtr4	-.0811653	.0702724	-1.16	0.254	-.222383	.0600524
t	-.0055712	.0039944	-1.39	0.169	-.0135983	.0024559
a13	.0774788	.1001482	0.77	0.443	-.1237765	.2787342
_cons	2.031294	.7916691	2.57	0.013	.4403754	3.622213

- *Results: New Bedford Area*

The results for Fairhaven are unchanged.

Regression 3.2.7a: Fairhaven %Fishing

Regression with robust standard errors

Number of obs = 56
 F(6, 49) = 12.14
 Prob > F = 0.0000
 R-squared = 0.5765
 Root MSE = .34889

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pctfFH						
pctfFH_1	-.1364672	.1247292	-1.09	0.279	-.38712	.1141855
qtr2	1.012236	.1431545	7.07	0.000	.7245564	1.299916
qtr3	.8959443	.1781806	5.03	0.000	.537877	1.254012
qtr4	.4732682	.1540333	3.07	0.003	.1637267	.7828097
t	.0041742	.0048543	0.86	0.394	-.0055808	.0139292
a13	.0702211	.1909397	0.37	0.715	-.3134866	.4539288
_cons	2.540868	.3484191	7.29	0.000	1.840693	3.241042

Regression 3.2.7b: Fairhaven %Manufacturing - %Fishing

Regression with robust standard errors

Number of obs = 56
 F(6, 49) = 66.04
 Prob > F = 0.0000
 R-squared = 0.8618
 Root MSE = 1.1879

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
diff_FH						
diffFH_1	.5618228	.1819314	3.09	0.003	.196218	.9274277
qtr2	-.0098569	.43872	-0.02	0.982	-.8914978	.8717839
qtr3	-.1275381	.5578513	-0.23	0.820	-1.248582	.993506
qtr4	.3589172	.530379	0.68	0.502	-.7069194	1.424754
t	-.0113483	.0199423	-0.57	0.572	-.0514238	.0287272
a13	-2.14924	1.260754	-1.70	0.095	-4.682819	.3843393
_cons	6.699502	2.790863	2.40	0.020	1.091054	12.30795

Table AD 2.5 provides the 90% confidence interval around job and wage losses for New Bedford.

Regression 3.2.8a: New Bedford %Fishing

Regression with robust standard errors

Number of obs = 56
 F(6, 49) = 21.97
 Prob > F = 0.0000
 R-squared = 0.6089
 Root MSE = .25912

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pctfNB_1	.1796623	.1405425	1.28	0.207	-.1027686	.4620931
qtr2	.412374	.1294099	3.19	0.003	.1523151	.6724328
qtr3	.3716062	.1166649	3.19	0.003	.1371594	.606053
qtr4	.4570343	.1219105	3.75	0.000	.2120459	.7020227
t	.0152875	.0040154	3.81	0.000	.0072183	.0233567
a13	-.2243452	.1137599	-1.97	0.054	-.4529543	.0042638
_cons	1.747051	.299221	5.84	0.000	1.145744	2.348358

This decrease amounts to a 0.22 decrease in percentage fisheries employment, which is an over-all decline of 7.5%. Table 3.4 displays the calculated effects.

Table AD2.5: Fishing Employment Changes in New Bedford

Town	Change in % Employment in Fishing	% Change in Fishing Employment Levels	Wage Change Lower Bound (# of Jobs)	Wage Change (# of Jobs)	Wage Change Upper Bound (# of Jobs)
New Bedford	- 0.22%	- 7.5 %	-\$12,992,000 (-127)	-\$7,058,000 (-69)	-\$1,228,000 (-12)

The results for the difference in difference estimator for New Bedford are unchanged.

Regression 3.2.8b: New Bedford %Manufacturing - %Fishing

Regression with robust standard errors

Number of obs = 56
 F(6, 49) = 100.48
 Prob > F = 0.0000
 R-squared = 0.8459
 Root MSE = .82598

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
diffNB_1	.8102661	.0690704	11.73	0.000	.6714638	.9490683
qtr2	-.9827944	.4134447	-2.38	0.021	-1.813643	-.1519463
qtr3	-.8449193	.4575356	-1.85	0.071	-1.764372	.0745329
qtr4	-.7629715	.4587783	-1.66	0.103	-1.684921	.1589779
t	-.0052374	.011566	-0.45	0.653	-.0284801	.0180053
a13	.259074	.3860363	0.67	0.505	-.516695	1.034843
_cons	4.744393	1.851263	2.56	0.014	1.02414	8.464646

Regression 3.2.8b does not indicate a statistically significant change in the relative share of fishing employment (to manufacturing) associated with Amendment 13.

The results for Westport are unchanged

Regression 3.2.9a: Westport %Fishing

Regression with robust standard errors

Number of obs = 56
F(6, 49) = 49.57
Prob > F = 0.0000
R-squared = 0.7400
Root MSE = .21093

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
pctfWP						
pctfWP_1	.4338808	.1140125	3.81	0.000	.2047641	.6629974
qtr2	.0565791	.0941256	0.60	0.551	-.1325734	.2457317
qtr3	.129947	.0904153	1.44	0.157	-.0517493	.3116433
qtr4	.1342256	.0862733	1.56	0.126	-.0391471	.3075982
t	-.0078699	.0042961	-1.83	0.073	-.0165033	.0007635
a13	-.1377447	.1141996	-1.21	0.234	-.3672374	.0917481
_cons	1.147227	.3100528	3.70	0.001	.5241523	1.770301

This result is not statistically significant.

The share of fisheries employment in Westport (relative to manufacturing) does significantly fall after Amendment 13 is implemented, however, as indicated in Regression 3.2.9b:

Regression 3.2.9b: Westport %Manufacturing - %Fishing

Regression with robust standard errors

Number of obs = 56
F(6, 49) = 39.49
Prob > F = 0.0000
R-squared = 0.8698
Root MSE = .29816

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
diff_WP						
diffWP_1	.4927644	.1047118	4.71	0.000	.2823381	.7031906
qtr2	-.4807443	.1186322	-4.05	0.000	-.7191446	-.242344
qtr3	-.5296778	.1269372	-4.17	0.000	-.7847677	-.2745878
qtr4	-.5624399	.1432728	-3.93	0.000	-.8503575	-.2745224
t	-.0251751	.0076931	-3.27	0.002	-.0406349	-.0097153
a13	.3645817	.1780013	2.05	0.046	.0068748	.7222887
_cons	2.905633	.5812416	5.00	0.000	1.737585	4.073682

This decrease in the relative share of fisheries employment occurs despite a decrease in manufacturing employment of 36% over the study period.

AD.II.2 Summary of Results and Importance to the Study

Table 3.8 summarizes the results of this chapter. The results for unemployment show higher rates for Gloucester and New Bedford relative to both the US and Massachusetts. The results for Amendment 13 show statistically significant and insignificant results for losses in percentage employment in fishing and the share of employment in fishing.

These results show a preponderance of effects due to Amendment 13. The results for Amendments 5 and 7 also show employment and share losses (though these results are not as methodologically strong as for Amendment 13).

Table 3.8 Results Summary

UNEMPLOYMENT					
Town	US Comparison		MA Comparison		
Gloucester	Higher than US average		Higher than MA average		
New Bedford	Higher than US average		Higher than MA average		
Plymouth	Lower than US average		Close to MA average		
AMENDMENT 13					
Town/County	Employment Loss	Statistically Significant	Wages Lost	Share Loss	Statistically Significant
Boston	Yes	No	--	Yes	Yes
Gloucester	Yes	Yes	\$ 2,286,000	Yes	No
Rockport	Yes	Yes	\$ 139,000	Yes	No
Barnstable County	No	No	--	Yes	No
Plymouth County	Yes	Yes	\$ 353,000	Yes	No
Fairhaven	No	No	--	No	No
New Bedford	Yes	Yes	\$ 7,058,000	Yes	No
Westport	Yes	No	--	Yes	Yes
AMENDMENTS 5 AND 7					
Town	Employment Loss		Share Loss		
Gloucester	Yes		Yes		
New Bedford	Yes		Yes		
Plymouth	Yes		Yes		

The towns of Gloucester and New Bedford have had higher unemployment rates than the rest of Massachusetts and the nation as a whole. Plymouth experienced higher unemployment rates in the 1990s but this trend appears to have reversed itself. These results do not directly show the impact of fishing but indicate that these towns are particularly susceptible to economic distress.

The analysis of fishing employment specifically indicates a loss of fishing jobs in Gloucester, Rockport, and Plymouth County coincident with the implementation of Amendment 13. This lost wage revenue constitutes almost \$10 million, or roughly \$6.5 million per year.¹¹

If this loss value is multiplied by a “multiplier effect” wherein the indirect effects of money entering a community are disseminated and spur further investment and

¹¹ The wage estimates for New Bedford may be biased upwards. New Bedford includes the scallop industry which is a high value industry, and as such may offer higher wages than the groundfish industry. However, no groundfish specific wage data are available. In addition, no estimates of lost wages are made for towns that have apparent (though not statistically significant) job losses – this omission would tend to bias the value of foregone wages downward.

consumption, the estimated impact will be greater. A conservative multiplier of 1.7 would result in an estimated loss of \$16.5 million dollars (\$11 million annually) in community impact across the three towns for who the effects are estimated.¹²

Taken as a whole, the indications are that Amendment 13 affected all 8 of the study areas except Fairhaven (though Barnstable County experienced negligible effects). Boston and Westport lost share in fisheries employment relative to manufacturing. Boston, Rockport, and Westport all experienced some decrease in fisheries employment, though these effects were not significant at the 95% confidence level.

The analysis of aggregated fishing employment data (including agriculture and hunting) for the decennial census of 1990 and 2000 show that employment in these areas decreased in Gloucester, New Bedford, and Plymouth. Difference in difference estimation also indicated that, relative to the share of employment in manufacturing, the share of employment in fishing decreased.

These decreases in fishing employment show that there have been quantifiable losses to the Massachusetts economy. Other losses are harder to quantify due to either a lack of data (data not available earlier in time or in specific towns) or the inability of data to address specific issues. One such issue involves part time fishing. Many former full-time fishers have resorted to fishing part-time. Other former part-time fishers, who may not have appeared as employed in fishing in the official statistics have either had to diversify their activities further or stop fishing entirely. Due to data limitations, the estimated fishing employment effects should be seen as a minimum bound on lost value in employment due to fishing employment foregone.

Should future regulations curtail employment further, some ports (such as Gloucester and Plymouth) could see groundfish fishing cease altogether. As tables 3.2 and 3.3 indicate, relatively modest job losses account for large portions of employment in these fisheries and result in substantial wage losses.

In sum, the results of this chapter indicate that Massachusetts' fishing communities have:

- experienced higher rates of unemployment,
- lost employment in fishing after the passage of Amendment 13, and
- lost employment in fishing during the 1990s.

¹² This multiplier is the lowest multiplier value suggested in Hughes, David (2003) "Policy Uses of Economic Multiplier and Impact Analysis," *Choices*, an online magazine of the American Agricultural Economics Association, vol. 18, no. 2.

AD.III. Economic Models of Fisheries Activities - Addendum

AD.III.I. The Production Function for a Fishing Unit - Addendum

In chapter V of Baden and Bianconi (2006), we presented simulations of a production model for fisheries. Here, we extend those simulations in two directions. First, we consider simulations in reference to Tables 5.1 and 5.2 in Baden and Bianconi (2006), with the alternative assumption that the base set for comparisons remains fixed across different levels of risk taking. Second, we implement a more restrictive policy of 50% reduction in sea space and labor input to show cases where expected profits can become negative, reflecting the possibility for the unit to abandon the productive activity.

Table AD.3.1 present simulations of the original model with the alternative assumption that the base set remain fixed at the no risk level. In reference to the original Table 5.2, we note that changes are smaller in magnitude. This reflects the main point that when a restriction is imposed, it is in a one-size-fits-all fashion, i.e. it applies uniformly to market participants taking alternative levels of risk. Hence, to some participants, the regulation has a much smaller impact since, with different levels of risk, those participants may be more close to the bounds imposed by the restriction.

Tables AD.3.2 and AD.3.3 present simulations of the model with the alternative assumption that the base set remain fixed at the no risk level, but in addition, the restriction on sea level and labor input is set at the higher 50% rate. The shaded boxes illustrate the cases where short run expected profits become negative, because of the imposed regulation. A productive unit could eventually face negative returns in the short run that would make it choose an alternate activity, in effect taking the unit out of business. In this case, again we observe that some changes are negligible reflecting the main point that when a restriction is imposed, it is in a one-size-fits-all fashion. Hence, to some participants, the regulation has a much smaller impact.

Table AD.3.1:
 Model Simulations – Percent Changes from Base Set – Revised - Addendum

	Probability of Good State: 1 (No Risk)	Probability of Good State: 0.85 (Moderate Risk)	Probability of Good State: 0.5 (Maximum Risk)
	All % Changes From Base Set		
30% Reduction of Sea Space: Short Run with Physical Capital Fixed			
Capital/Labor	4.8%	15.1%	-4.8%
Effort/Labor	0.0%	0.0%	0.0%
Expected Profit/Output	0.0%	0.0%	-12.5%
Expected Profit/Labor	8.3%	-8.3%	-10.0%
Expected Output/Labor	0.0%	0.0%	0.0%
30% Reduction of Sea Space: Long Run with Physical Capital Variable			
Capital/Labor	0.0%	0.0%	0.0%
Effort/Labor	0.0%	0.0%	0.0%
Expected Profit/Output	10.0%	22.2%	-12.5%
Expected Profit/Labor	16.7%	16.7%	-10.0%
Expected Output/Labor	0.0%	0.0%	0.0%
30% Reduction of Sea Space and 30% Reduction of Work Activity: Short Run with Physical Capital Fixed			
Capital/Labor	44.9%	167.9%	-29.5%
Effort/Labor	33.3%	46.7%	-11.3%
Expected Profit/Output	10.0%	-66.7%	-25.0%
Expected Profit/Labor	33.3%	-58.3%	-40.0%
Expected Output/Labor	13.6%	36.8%	-10.4%
30% Reduction of Sea Space and 30% Reduction of Work Activity: Long Run with Physical Capital Variable			
Capital/Labor	5.1%	8.3%	-2.6%
Effort/Labor	33.3%	20.0%	-3.8%
Expected Profit/Output	30.0%	55.6%	-25.0%
Expected Profit/Labor	41.7%	58.3%	-20.0%
Expected Output/Labor	4.8%	8.0%	-2.4%

Table AD.3.2: Model Simulations-Addendum

	Probability of Good State: 1 (No Risk)	Probability of Good State: 0.85 (Moderate Risk)	Probability of Good State: 0.5 (Maximum Risk)
Base Set			
Capital/Labor	3.12	3.12	3.12
Effort/Labor	0.03	0.15	0.53
Expected Profit/Output	0.10	0.10	0.08
Expected Profit/Labor	0.12	0.12	0.10
Expected Output/Labor	1.25	1.25	1.25
50% Reduction of Sea Space: Short Run with Physical Capital Fixed			
Capital/Labor	3.40	3.75	3.11
Effort/Labor	0.03	0.15	0.53
Expected Profit/Output	0.10	0.08	0.08
Expected Profit/Labor	0.12	0.10	0.10
Expected Output/Labor	1.25	1.25	1.25
50% Reduction of Sea Space: Long Run with Physical Capital Variable			
Capital/Labor	3.12	3.12	3.12
Effort/Labor	0.03	0.16	0.53
Expected Profit/Output	0.12	0.12	0.08
Expected Profit/Labor	0.15	0.15	0.10
Expected Output/Labor	1.25	1.25	1.25
50% Reduction of Sea Space and 50% Reduction of Work Activity: Short Run with Physical Capital Fixed			
Capital/Labor	6.32	11.70	3.07
Effort/Labor	0.06	0.25	0.53
Expected Profit/Output	0.10	-0.04	0.08
Expected Profit/Labor	0.15	-0.09	0.10
Expected Output/Labor	1.60	1.91	1.25
50% Reduction of Sea Space and 50% Reduction of Work Activity: Long Run with Physical Capital Variable			
Capital/Labor	3.44	3.49	3.13
Effort/Labor	0.06	0.20	0.53
Expected Profit/Output	0.16	0.16	0.08
Expected Profit/Labor	0.21	0.22	0.10
Expected Output/Labor	1.36	1.40	1.25

Table AD.3.3: Model Simulations – Percent Changes from Base Set-addendum

	Probability of Good State: 1 (No Risk)	Probability of Good State: 0.85 (Moderate Risk)	Probability of Good State: 0.5 (Maximum Risk)
	All % Changes From Base Set		
50% Reduction of Sea Space: Short Run with Physical Capital Fixed			
Capital/Labor	9.0%	20.2%	-0.3%
Effort/Labor	0.0%	0.0%	0.0%
Expected Profit/Output	0.0%	-20.0%	0.0%
Expected Profit/Labor	0.0%	-16.7%	0.0%
Expected Output/Labor	0.0%	0.0%	0.0%
50% Reduction of Sea Space: Long Run with Physical Capital Variable			
Capital/Labor	0.0%	0.0%	0.0%
Effort/Labor	0.0%	6.7%	0.0%
Expected Profit/Output	20.0%	20.0%	0.0%
Expected Profit/Labor	25.0%	25.0%	0.0%
Expected Output/Labor	0.0%	0.0%	0.0%
50% Reduction of Sea Space and 50% Reduction of Work Activity: Short Run with Physical Capital Fixed			
Capital/Labor	102.6%	275.0%	-1.6%
Effort/Labor	100.0%	66.7%	0.0%
Expected Profit/Output	0.0%	-140.0%	0.0%
Expected Profit/Labor	25.0%	-175.0%	0.0%
Expected Output/Labor	28.0%	52.8%	0.0%
50% Reduction of Sea Space and 50% Reduction of Work Activity: Long Run with Physical Capital Variable			
Capital/Labor	10.3%	11.9%	0.3%
Effort/Labor	100.0%	33.3%	0.0%
Expected Profit/Output	60.0%	60.0%	0.0%
Expected Profit/Labor	75.0%	83.3%	0.0%
Expected Output/Labor	8.8%	12.0%	0.0%

AD.VI. Concluding Remarks

- The wide and in depth econometric analysis of sales tax data does show some negative impact of Amendments on coastal towns where fishing is active. Main dollar value results are summarized in Table AD.4.1. For Amendment 5, in 1994-1995 US dollars, in the worst case scenario, costal towns had average losses of \$2,168 per month and a total loss of \$2,691,000. In the best case scenario, costal towns had average losses of \$354 per month and a total loss of \$176,468. For Amendment 13, the effect on the Port of Gloucester is in 2005 US dollars, in the worst case scenario, an average loss of \$17,339 per month per business unit and a total loss of \$11,400,000 across the town. In the best case scenario, the town had average of no losses, of \$0 per month per business unit, and a total gain of \$1,000,000 across the town. Hence, the predictive power of future fisheries regulations may have further impacts on gross sales tax receipts that lie within the worst and best case scenarios obtained here.

Table AD.4.1

	<i>Average Gross Sales Tax Receipts</i>	<i>95% Confidence Interval</i>	<i>Total Gross Sales Tax Receipts</i>	<i>95% Confidence Interval</i>
<i>Amendment 5:</i> <u>Impact on all 29 Coastal Towns of MA State - 1994-1995 US Dollars</u>	Receipts Loss per Month per Town \$1,247	Receipts Loss per Month per Town \$2,168 ↔ \$354	Receipts Loss per Month per Town \$1,235,000	Receipts Loss per Month per Town \$2,691,000 ↔ \$176,468
<i>Amendment 13:</i> <u>Impact on the Port of Gloucester - 2005 US Dollars</u>	Receipts Loss per Month per Unit \$8,422	Receipts Loss per Month per Unit \$17,339 ↔ \$0	Receipts Loss per Month \$5,200,000	Receipts Loss (+ Gain) per Month \$11,400,000 ↔ +\$1,000,000

- The employment results in section AD.II show that Massachusetts' fishing communities have:
 - experienced higher rates of unemployment,
 - lost employment in fishing after the passage of Amendment 13, and
 - lost employment in fishing during the 1990s.

- In theory, large restrictions can impose negative expected profits which may take a productive unit out of the business activity as seen in Tables AD.3.2 and AD.3.3.

BIBLIOGRAPHY

Baden, Brett and Marcelo Bianconi (2006) *Economic Impact of Regulations on the Marine Fisheries Industry in the State of Massachusetts*. The Massachusetts Fisheries Recovery Commission, Gloucester, MA, June 30.

AD.V. June 30, 2006-August 15, 2006: Report of Activities

1. Dates of all meetings, major participants, purpose

i. Monday, July 10, 2006 – 2:30-5:30pm, Gloucester, 30 Emerson Ave.: Presentation of Baden and Bianconi (2006), present Marcelo Bianconi, Brett Baden (via phone), Senator Bruce Tarr, Representative Anthony Verga, Paul Diodati, Dr. Brian Rothschild, Vito Calomo, Ann-Margaret Ferrante, and others.

ii. Wednesday, July 10, 2006, Noon: Meeting with Paul Diodati and David Pierce at Boston's Causeway St. office. Marcelo Bianconi and Brett Baden.

iii. Thursday, July 27, 2006, meeting with Dr. Brian Rothschild at UMass Dartmouth SMF, with Vito Calomo, Marcelo Bianconi and Brett Baden.

iii. Tuesday, August 1, 2006, meeting with Senator Bruce Tarr at Massachusetts State House, with Vito Calomo, Marcelo Bianconi and Brett Baden.

Baden and Bianconi, August 27, 2006